# Non-Malleable Extractors with Shorter Seeds and Min-Entropy Rate $<\frac{1}{2}$ 

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#### Abstract

Motivated by the problem of how to communicate over a public channel with an active adversary, Dodis and Wichs [DW09] introduced the notion of a non-malleable extractor, as a much stronger version of a strong extractor. A non-malleable extractor nmExt takes two inputs, a weakly-random $W$ and a uniformly random seed $S$, and outputs a string which is nearly uniform, given $S$ as well as $\operatorname{nmExt}(W, \mathcal{A}(S))$, for an arbitrary function $\mathcal{A}$ with $\mathcal{A}(S) \neq S$. Cohen, Raz and Segev in CCC'12 presented an explicit construction of a non-malleable extractor with short seeds. For any integers $n$ and $d$ such that $2.01 \cdot \log n \leq d \leq n$, they proposed an explicit construction of a non-malleable extractor nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ with error exponentially small in $m$. However, their result suffers from some drawbacks: First, the non-malleable extractor is constructed based on Raz's etractor in SOTC'05, while the error estimation in that construction is too rough. Second, though they aimed to shorten the length of the seed, the lower bound of the seed length is not optimal. Moreover, their construction requires the min-entropy rate to be greater than $\frac{1}{2}$. In this paper, we improve the error estimation of Raz's extractor, which plays an extremely important role in the constraints of the non-malleable extractor parameters including the seed length. Then we present an improved explicit construction of non-malleable extractors with shorter seed length by using biased variable sequence for linear tests. More precisely, we construct an explicit (1016, $\frac{1}{2}$ ) - 1 -non-malleable extractor nmExt : $\{0,1\}^{2^{10}} \times\{0,1\}^{d} \rightarrow\{0,1\}$ with seed length 19 , while it should be no less than $\frac{46}{63}+66$ according to Cohen et al. in CCC' 12 . Therefore, it beats the condition " $2.01 \cdot \log n \leq d \leq n$ ", since $d$ is just $1.9 \cdot \log n$ in our construction. We also give a general explicit construction of nonmalleable extractors and analyze the simplification of the constraints on the parameters. Furthermore, we show an explicit construction of non-malleable extractors for the min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ for some constant $\delta>0$, while the min-entropy should be greater than $\frac{1}{2} n$ by Cohen et al. in CCC' 12 . We also propose a general construction of non-malleable extractors with minentropy $k=\left(\frac{1}{2}-\delta\right) n$ from any non-malleable extractor with min-entropy


# $k>\frac{1}{2} n$ by employing a special encoding technique and the property of statical distance. Compared with Li's construction in FOCS'12 by using inner product function, our construction is more general. Finally, we give their application to privacy amplification. 

Keywords: extractors; non-malleable; seed length; min-entropy; privacy amplification

## 1 Introduction

Randomness extractors are functions that convert weakly random sources into nearly uniform bits. Though the motivation of extractors is to simulate randomized algorithms with weak random sources as might arise in nature, randomness extractors have been successfully applied to coding theory, cryptography, complexity, etc. $[9,11,19]$. In this paper, we focus on the extractors that are applied to privacy amplification. In this scenario, two parties Alice and Bob share a weakly random secret key $W \in\{0,1\}^{n}$. The secret source $W$ may be a humanmemorizable password, some biometric data, and physical sources, which are themselves weakly random, or a uniform secret which may have been partially leaked to an adversary Eve. Thus, only the min-entropy of $W$ is guaranteed. Alice and Bob interact over a public communication channel in order to securely agree on a nearly uniform secret key $R \in\{0,1\}^{m}$ in the presence of the adversary, Eve, who can see every message transmitted in the public channel. The min-entropy of $W$ and the length of the public seed are two main measures of efficiency in this setting. If Eve is passive, a (strong) randomness extractor yields the following solution: Alice sends a uniformly random seed $S$ to Bob, then they both compute $R=\operatorname{Ext}(W, X)$ as the nearly uniform secret key [15]. If Eve is active (i.e., it may change the messages in arbitrary ways), some protocols have been proposed to achieve this goal $[4-8,10-12,18,20]$.

As a major progress, Dodis and Wichs [8] introduced non-malleable extractors to study privacy amplification protocols, where the attacker is active and computationally unbounded. If an attacker sees a random seed $S$ and modifies it into an arbitrarily related seed $S^{\prime}$, then the relationship between $R=\operatorname{Ext}(W, S)$ and $R^{\prime}=\operatorname{Ext}\left(W, S^{\prime}\right)$ is bounded to avoid related key attacks. More formally, a non-malleable extractor is a function nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ that takes two inputs, a weakly-random secret source $W$ with min-entropy $k$ and uniformly random seed $S$, and outputs a string which is $\epsilon$-close to uniform in statistical distance, given $S$ as well as $\operatorname{nmExt}(W, \mathcal{A}(S))$, for an arbitrary function $\mathcal{A}$ with $\mathcal{A}(S) \neq S$. They proved that $(k, \epsilon)-$ non-malleable extractors exist as long as $k>2 m+3 \log \frac{1}{\epsilon}+\log d+9$ and $d>\log (n-k+1)+2 \log \frac{1}{\epsilon}+7$.

The first explicit non-malleable extractor was constructed by Dodis, Li, Wooley and Zuckerman [7]. It works for any weakly random input source with the min-entropy $k>\frac{n}{2}$ and uniformly random seed of length $d=n$ (It works even if the seed has entropy only $\Theta(m+\log n))$. However, when outputting more than a logarithmic number of bits, its efficiency relies on a longstanding conjecture on
the distribution of prime numbers. Based on the inner product function, Li [11] later showed the first construction of non-malleable extractors for any weakly random source with the min-entropy $k=\left(\frac{1}{2}-\delta\right) \cdot n$ for some constant $\delta>0$, and a conditional (semi-explicit) construction that can potentially achieve $k=\alpha n$ for any constant $\delta>0$.

In 2012, an alternative explicit construction based on the extractor of Raz [17] was given by Cohen, Raz, and Segev in [5]. Without using any conjecture, their construction works for any weakly random source with the min-entropy $k=\left(\frac{1}{2}+\delta\right) \cdot n$ and uniformly random seed of length $d \geq \frac{23}{\delta} \cdot t m+2 \log n$ (see Theorem 1 for details). However, their result suffers from some drawbacks: First, the non-malleable extractor is constructed based on the explicit seeded extractor of Raz [17], while the error estimation in that construction is too rough. Second, though one main purpose of Cohen et al. in [5] is to shorten the length of the seed, the lower bound of the seed length is not optimal. Moreover, their construction requires the min-entropy rate (i.e., the ratio between the entropy and the length of the weakly random string of the source) to be greater than $\frac{1}{2}$.

## OUR CONTRIBUTIONS AND TECHNIQUES.

- We improve the error estimation of Raz's extractor in SOTC'05 [17], a special case of which was used by Cohen et al. in CCC'12 [5]. For simplicity, denote $\gamma_{1}$ as the error of the extractor in [5], and $\gamma_{2}$ as the counterpart in this paper. Recall that $\gamma_{1}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{k}} \cdot(2 \epsilon)^{\frac{1}{k}}$ in [5] under the assumption that $\epsilon \geq 2^{-\frac{d k}{2}} \cdot k^{k}$ and $0<\delta<\frac{1}{2}$. If $\epsilon \geq \frac{1}{2^{\left(\frac{1}{2}-\delta\right) n+1}}$, then $\gamma_{1}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{k}} \cdot(2 \epsilon)^{\frac{1}{k}} \geq$ 1. In this case, the error estimation is meaningless. One main reason is that in those proofs, the partition method about the sum [5,17] which bounds the error didn't capture the essence of the biased sequence for linear tests. In this paper, we propose another partition method and give a better bound on the sum by employing the combination and permutation formulas. Correspondingly, the error is $\gamma_{2}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{k}} \cdot\left\{2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1+\left[1-2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1\right] \cdot \epsilon\right\}^{\frac{1}{k}}$. Since $\epsilon \geq 2^{-\frac{d k}{2}} \cdot k^{k}$ and $2^{-\frac{d k}{2}} \cdot k^{k}>2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdot 1$ for any even integer $k$, we get $\gamma_{1}>\gamma_{2}$. To simplify this bound, let $k$ be a specific value. For instance, let $k=4$, then the error $\gamma_{2}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{4}} \cdot\left[2^{-2 d} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}}$.
- Based on the improvement of the error estimation, we present an explicit construction of non-malleable extractors, which is an improvement of the construction of Cohen et al. in CCC'12 [5] in the sense that the seed length can be shorter. More concretely, we present an explicit (1016, $\frac{1}{2}$ ) - 1 -non-malleable extractor nmExt : $\{0,1\}^{2^{10}} \times\{0,1\}^{d} \rightarrow\{0,1\}$ with seed length $d=19$, which beats the condition " $2.01 \cdot \log n \leq d \leq n$ " in [5], since $d$ is just $1.9 \cdot \log n$ in our construction while it should be no less than $\frac{46}{63}+66$ according to [5]. We also give a general explicit construction of non-malleable extractors and analyze the simplification of the constraints on the parameters. Similar to [5], the non-malleable extractors are also constructed by using biased variable sequence for linear tests. In the proof, the error of the extractor impacts greatly on the constraints of the
parameters including the length of the seed, the min-entropy of the source, the error of the non-malleable extractor.
- We give an explicit construction of non-malleable extractors for min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ for some constant $\delta>0$ by using biased sequence for linear tests, while the min-entropy $k$ should be greater than $\frac{1}{2} n$ in [5]. We also propose a high level way of constructing non-malleable extractors with min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ from any non-malleable extractor with min-entropy $k>\frac{1}{2} n$ by employing a special encoding technique and the property of statical distance. Compared with Li's construction in FOCS'12 [11] by using inner product function, our construction is more general.

ORGANIZATION. The remainder of the paper is organized as follows. In Section 2, we review some notations, concepts, and results. In section 3, we show an existing central lemma about the error estimation of Raz's Extractor and improve it by proposing a new partition method. In section 4 , we propose the explicit construction of non-malleable extractors with shorter seed length compared with that in [5]. Then we show a general explicit construction of non-malleable extractors and analyze the simplification of the constraints on the parameters. In Section 5, we give an explicit construction of non-malleable extractors for min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ for some constant $\delta>0$ based on $\epsilon$ biased variable sequence for linear tests, and propose a general construction with min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ from any non-malleable extractor with min-entropy $k>\frac{1}{2} n$. In Section 6, we show how the non-malleable extractors can be applied to privacy amplification. Section 7 concludes the paper.

## 2 Preliminaries

In this section, we'll recall some notations, definitions, and results that will be used later.

For any positive integer $n$, denote $[n]=\{1,2, \ldots, n\}$. Denote $U_{m}$ as the uniformly random distribution over $\{0,1\}^{m}$. We measure the distance between two distributions by the $\mathcal{L}_{1}$ norm in order to be consistent with [5]. We say the distribution $X$ is $\epsilon$-close to the distribution $Y$ if $\|X-Y\|_{1}=\sum_{s} \mid \operatorname{Pr}[X=s]-\operatorname{Pr}[Y=$ $s] \mid \leq \epsilon{ }^{1}$. The statistical distance of $X$ and $Y$ is defined as $\mathrm{SD}(X, Y)=\frac{1}{2} \| X-$ $Y \|_{1}$. It's well known that for any function $f, \mathrm{SD}(f(X), f(Y)) \leq \mathrm{SD}(X, Y)$. Denote $\mathrm{SD}\left(\left(X_{1}, X_{2}\right),\left(Y_{1}, Y_{2}\right) \mid Z\right)$ as the abbreviation of $\mathrm{SD}\left(\left(X_{1}, X_{2}, Z\right),\left(Y_{1}, Y_{2}, Z\right)\right)$.

The min-entropy of a random variable $W$ is $H_{\infty}(W)=-\log \max _{w} \operatorname{Pr}(W=$ $w)$. For $W \in\{0,1\}^{n}$, we say $W$ is an $(n, k)$-source if $H_{\infty}(W) \geq k$, and we call $W$ has entropy rate $\frac{H_{\infty}(W)}{n}$. We say that a source (i.e., a random variable) is a weak source if its distribution is not uniform. We say $W$ is a flat source if it is a uniform distribution over some subset $S \subseteq\{0,1\}^{n}$. Chor and Goldreich [3]

[^0]observed that the distribution of any $(n, k)$-source is a convex combination of distributions of flat $(n, b)$-sources. Therefore, for general weak sources, it will be enough to consider flat sources instead in most cases.

Definition 1. A function Ext : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ is a $(k, \epsilon)-$ seeded extractor if for every $(n, k)$-source $W$ and an independent uniformly random variable $S$ (called seed) over $\{0,1\}^{d}$, the distribution of $\operatorname{Ext}(W, S)$ is $\epsilon$-close to $U_{m} . \epsilon$ is called the error of the extractor. A seeded extractor is a strong $(k, \epsilon)$-extractor if for $W$ and $S$ as above, the distribution of $(\operatorname{Ext}(W, S), S)$ is $\epsilon$-close to $\left(U_{m}, U_{d}\right)$.
Definition 2. A random variable $Z$ over $\{0,1\}$ is $\epsilon$-biased if $\operatorname{bias}(Z)=\mid \operatorname{Pr}[Z=$ $0]-\operatorname{Pr}[Z=1] \mid \leq \epsilon$ (i.e., $Z$ is $\epsilon$-close to uniform). A sequence of $0-1$ random variables $Z_{1}, \ldots, Z_{N}$ is $\epsilon$-biased for linear tests of size $k$ if for any nonempty $\tau \subseteq[N]$ with $|\tau| \leq k$, the random variable $Z_{\tau}=\oplus_{i \in \tau} Z_{i}$ is $\epsilon$-biased. We also say that the sequence $Z_{1}, Z_{2}, \ldots, Z_{N} \epsilon$-fools linear tests of size $k$.

For every $k^{\prime}, N \geq 2$, variables $Z_{1}, \cdots, Z_{N}$ as above can be explicitly constructed using $2 \cdot\left\lceil\log (1 / \epsilon)+\log k^{\prime}+\log \log N\right\rceil$ random bits [1].

The Extractor of Raz. Raz [17] constructed an extractor based on a sequence of 0-1 random variables that have small bias for linear tests of a certain size. Let $Z_{1}, \cdots, Z_{m \cdot 2^{d}}$ be 0 -1 random variables that are $\epsilon$-biased for linear tests of size $k^{\prime}$ that are constructed using $n$ random bits. The set of indices $\left[m \cdot 2^{d}\right]$ can be considered as the set $\left\{(i, s): i \in[m], s \in\{0,1\}^{d}\right\}$. Define Ext : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ by $\operatorname{Ext}(w, s)=Z_{(1, s)}(w)\left\|Z_{(2, s)}(w) \ldots\right\| Z_{(m, s)}(w)$, where " $\mid$ " is the concatenation operator. Raz proposed that Ext is a seeded extractor with good parameters [17].

Cohen et al. [5] proved that the above extractor is in fact non-malleable. We'll also construct non-malleable extractors based on it. The formal definition of non-malleable extractors is as follows.
Definition 3. (see [8]) A function $\mathcal{A}:\{0,1\}^{d} \rightarrow\{0,1\}^{d}$ is an adversarial function if for every $s \in\{0,1\}^{d}$ it holds that $\mathcal{A}(s) \neq s$. A function nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ is a $(k, \epsilon)$-non-malleable extractor if for every $(n, k)$ source $W$, independent uniformly random variable $S$ over $\{0,1\}^{d}$, and every adversarial function $\mathcal{A}:\{0,1\}^{d} \rightarrow\{0,1\}^{d}$,

$$
\left\|(\operatorname{nmExt}(W, S), \operatorname{nmExt}(W, \mathcal{A}(S)), S)-\left(U_{m}, \operatorname{nmExt}(W, \mathcal{A}(S)), S\right)\right\|_{1} \leq \epsilon
$$

$\epsilon$ is called the error of the extractor.
Definition 4. (see [5]) For any positive integer $t$, define $\mathcal{B}:\{0,1\}^{d} \rightarrow\{0,1\}^{t d}$ as $\mathcal{B}(s)=\mathcal{A}_{1}(s)\left\|\mathcal{A}_{2}(s) \cdots\right\| \mathcal{A}_{t}(s)$, where for all $i \in[t], \mathcal{A}_{i}$ is an adversarial function $\mathcal{A}_{i}:\{0,1\}^{d} \rightarrow\{0,1\}^{d}$ and "||" is the concatenation operator. We say $\mathcal{B}$ is a $t$-adversarial function. A function nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ is a $(k, \epsilon)-t$-non-malleable extractor if for every $(n, k)$-source $W$, independent uniformly random variable $S$ over $\{0,1\}^{d}$, and every $t$-adversarial function $\mathcal{B}$ :

$$
\begin{aligned}
\{0,1\}^{d} \rightarrow & \{0,1\}^{t d}, \\
& \|\left(\operatorname{nmExt}(W, S), \operatorname{nmExt}\left(W, \mathcal{A}_{1}(S)\right), \ldots, \operatorname{nmExt}\left(W, \mathcal{A}_{t}(S)\right), S\right) \\
& -\left(U_{m}, \operatorname{nmExt}\left(W, \mathcal{A}_{1}(S)\right), \ldots, \operatorname{nmExt}\left(W, \mathcal{A}_{t}(S)\right), S\right) \|_{1} \leq \epsilon
\end{aligned}
$$

One-time message authentication code (MAC) is used to guarantee that the received message is sent by a specified legitimate sender in an unauthenticated channel. Formally,
Definition 5. A family of functions $\left\{\mathrm{MAC}_{r}:\{0,1\}^{v} \rightarrow\{0,1\}^{\tau}\right\}_{r \in\{0,1\}^{m}}$ is a $\varepsilon$-secure (one-time) message authentication code (MAC) if for any $\mu$ and any function $f:\{0,1\}^{\tau} \rightarrow\{0,1\}^{v} \times\{0,1\}^{\tau}$, it holds that,

$$
\operatorname{Pr}_{r \leftarrow U_{m}}\left[\mathrm{MAC}_{r}\left(\mu^{\prime}\right)=\sigma^{\prime} \wedge \mu^{\prime} \neq \mu \mid\left(\mu^{\prime}, \sigma^{\prime}\right)=f\left(\mathrm{MAC}_{r}(\mu)\right)\right] \leq \varepsilon .
$$

The main theorem about the explicit construction of non-malleable extractors proposed in [5] is as follows.
Theorem 1. (see [5]) For any integers $n, d, m$ and $t$, and for any $0<\delta<\frac{1}{2}$ such that

$$
\begin{gathered}
d \geq \frac{23}{\delta} \cdot t m+2 \log n \\
n \geq \frac{160}{\delta} \cdot t m \\
\delta \geq 10 \cdot \frac{\log (n d)}{n}
\end{gathered}
$$

there exists an explicit $\left(\left(\frac{1}{2}+\delta\right) \cdot n, 2^{-m}\right)-t$-non-malleable extractor nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$.

Proof. See [5]. Omitted.
For simplicity, we only consider $t=1$ in the rest of the paper.

## 3 The Error Estimation and its Improvement of Raz's Extractor

In this section, we first recall the central lemma used in [5], which is a special case about the error estimation of Raz's Extractor [17]. Then we point out the flaw in the proof and improve its error estimation. Afterwards, we compare our result with the original one and roughly show the role of the improvement.

### 3.1 A Special Case of Raz's Extractor

The central lemma used in [5] is as follows, the proof of which is essentially the same as that in [17]. It can be considered as a special case of Raz's Extractor [17].

Lemma 1. (see Lemma 4.1 in [5]) Let $D=2^{d}$. Let $Z_{1}, \cdots, Z_{D}$ be 0-1 random variables that are $\epsilon$-biased for linear tests of size $k^{\prime}$ that are constructed using $n$ random bits. Define Ext ${ }^{(1)}:\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}$ by Ext ${ }^{(1)}(w, s)=Z_{s}(w)$, that is, Ext ${ }^{(1)}(w, s)$ is the random variable $Z_{s}$, when using $w$ as the value of the $n$ bits needed to produce $Z_{1}, \cdots, Z_{D}$. Then, for any $0<\delta<\frac{1}{2}$ and even integer $k=k^{\prime}$ such that $k \cdot\left(\frac{1}{\epsilon}\right)^{\frac{1}{k}} \leq D^{\frac{1}{2}}$, the function Ext ${ }^{(1)}$ is a $\left(\left(\frac{1}{2}+\delta\right) \cdot n, \gamma\right)$-seededextractor, with $\gamma=\left(\epsilon \cdot 2^{\left(\frac{1}{2}-\delta\right) n+1}\right)^{\frac{1}{k}}$.

Recall that the proof sketch is as follows.
Proof. Let $W$ be a $\left(n,\left(\frac{1}{2}+\delta\right) \cdot n\right)$-source. Let $S$ be a random variable that is uniformly distributed over $\{0,1\}^{d}$ and is independent of $W$. We will show that the distribution of $\operatorname{Ext}^{(1)}(W, S)$ is $\gamma$-close to uniform. As in [3], it is enough to consider the case where $W$ is uniformly distributed over a set $W^{\prime} \subseteq\{0,1\}^{n}$ of size $2^{(1 / 2+\delta) n}$. For every $w \in\{0,1\}^{n}$ and $s \in\{0,1\}^{d}$ denote

$$
e(w, s)=(-1)^{Z_{s}(w)}
$$

Claim 1. For any $r \in[k]$ and any different $s_{1}, \cdots, s_{r} \in\{0,1\}^{d}$,

$$
\sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{r} e\left(w, s_{j}\right) \leq \epsilon \cdot 2^{n} .
$$

Proof.

$$
\begin{aligned}
\sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{r} e\left(w, s_{j}\right) & =\sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{r}(-1)^{Z_{s_{j}}(w)} \\
& =\sum_{w \in\{0,1\}^{n}}(-1)^{Z_{s_{1}}(w) \oplus \cdots \oplus Z_{s_{r}}(w)},
\end{aligned}
$$

and since $Z_{s_{1}}(w) \oplus \cdots \oplus Z_{s_{r}}(w)$ is $\epsilon$-biased, the last sum is at most $\epsilon \cdot 2^{n}$.

The $\mathcal{L}_{1}$ distance of $\operatorname{Ext}^{(1)}(W, S)$ and $U$ is

$$
\begin{aligned}
& \left\|\mathrm{Ext}^{(1)}(W, S)-U\right\|_{1} \\
= & \left|\operatorname{Pr}\left[\mathrm{Ext}^{(1)}(W, S)=0\right]-\operatorname{Pr}[U=0]\right|+\left|\operatorname{Pr}\left[\mathrm{Ext}^{(1)}(W, S)=1\right]-\operatorname{Pr}[U=1]\right| \\
= & \left|\operatorname{Pr}\left[\mathrm{Ext}^{(1)}(W, S)=0\right]-\operatorname{Pr}\left[\mathrm{Ext}^{(1)}(W, S)=1\right]\right| \\
= & \left|\frac{1}{2^{\left(\frac{1}{2}+\delta\right) n}} \cdot \frac{1}{2^{d}}\left(\sum_{w \in W^{\prime}} \sum_{s \in\{0,1\}^{d}} e(w, s)\right)\right| .
\end{aligned}
$$

Denote $\gamma(W, S)=\frac{1}{2^{\left(\frac{1}{2}+\delta\right) n}} \cdot \frac{1}{2^{d}}\left(\sum_{w \in W^{\prime}} \sum_{s \in\{0,1\}^{d}} e(w, s)\right)$.
Define $f:[-1,1] \rightarrow[-1,1]$ by $f(z)=z^{k}$, then $f$ is a convex function for any even positive integer $k$.

Thus, by a convexity argument, we have

$$
\begin{aligned}
& 2^{\left(\frac{1}{2}+\delta\right) n} \cdot\left(2^{d} \cdot \gamma(W, S)\right)^{k} \\
= & 2^{\left(\frac{1}{2}+\delta\right) n} \cdot\left\{\sum_{w \in W^{\prime}}\left[\frac{1}{2^{(1 / 2+\delta) n}} \sum_{s \in\{0,1\}^{d}} e(w, s)\right]\right\}^{k} \\
\leq & 2^{\left(\frac{1}{2}+\delta\right) n} \cdot\left\{\sum_{w \in W^{\prime}} \frac{1}{2^{(1 / 2+\delta) n}}\left[\sum_{s \in\{0,1\}^{d}} e(w, s)\right]^{k}\right\} \\
\leq & \sum_{w \in\{0,1\}^{n}}\left[\sum_{s \in\{0,1\}^{d}} e(w, s)\right]^{k} \\
= & \sum_{w \in\{0,1\}^{n}} \sum_{s_{1}, \cdots, s_{k} \in\{0,1\}^{d}} \prod_{j=1}^{k} e\left(w, s_{j}\right) \\
= & \sum_{s_{1}, \cdots, s_{k} \in\{0,1\}^{d}} \sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{k} e\left(w, s_{j}\right) .
\end{aligned}
$$

The sum over $s_{1}, \cdots, s_{k} \in\{0,1\}^{d}$ is broken into two sums. The first sum is over $s_{1}, \cdots, s_{k} \in\{0,1\}^{d}$ such that in each summand, at least one $s_{j}$ is different than all other elements in $\left\{s_{1}, \cdots, s_{k}\right\}$, and the second sum is over $s_{1}, \cdots, s_{k} \in$ $\{0,1\}^{d}$ such that in each summand every $s_{j}$ is identical to at least one other element in $\left\{s_{1}, \cdots, s_{k}\right\}$. The number of summands in the first sum is trivially bounded by $2^{d \cdot k}$, and by Claim 1 each summand is bounded by $2^{n} \cdot \epsilon$. The number of summands in the second sum is bounded by $2^{d \cdot \frac{k}{2}} \cdot\left(\frac{k}{2}\right)^{k}$, and each summand is trivially bounded by $2^{n}$. Therefore,

$$
\begin{aligned}
2^{\left(\frac{1}{2}+\delta\right) n} \cdot 2^{d \cdot k} \cdot \gamma(W, S)^{k} \leq & 2^{n} \cdot \epsilon \cdot 2^{d \cdot k}+2^{n} \cdot 2^{d \cdot \frac{k}{2}} \cdot\left(\frac{k}{2}\right)^{k} \\
& \leq 2 \cdot 2^{n} \cdot \epsilon \cdot 2^{d \cdot k}
\end{aligned}
$$

where the last inequality follows by the assumption that $k \cdot(1 / \epsilon)^{1 / k} \leq D^{\frac{1}{2}}$. That is,

$$
|\gamma(W, S)| \leq\left(\epsilon \cdot 2^{\left(\frac{1}{2}-\delta\right) n+1}\right)^{\frac{1}{k}}
$$

The above partition method about the sum over $s_{1}, \cdots, s_{k} \in\{0,1\}^{d}$ is not optimal, since it doesn't capture the essence of random variable sequence that is biased for linear tests. Moreover, the bounds of the number of summands in the two sums are too large. The same problem exists in [17].

In fact, when every $s_{j}$ is identical to at least one other element in $\left\{s_{1}, \cdots, s_{k}\right\}$ under the assumption that at least one $s_{j}$ appears odd times in $\left\{s_{1}, \cdots, s_{k}\right\}$, the summand $\sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{k} e\left(w, s_{j}\right)$ is still upper bounded by $2^{n} \cdot \epsilon$, since
$\sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{k} e\left(w, s_{j}\right)=\sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{k}(-1)^{Z_{s_{j}}(w)}=\sum_{w \in\{0,1\}^{n}}(-1)^{Z_{s_{1}}(w) \oplus \cdots \oplus Z_{s_{k}}(w)}$ and $Z_{1}, \cdots, Z_{D}$ are 0-1 random variables that are $\epsilon$-biased for linear tests of size $k^{\prime}$. However, in this case the upper bound of the summand $\sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{k} e\left(w, s_{j}\right)$ was considered to be $2^{n}$ in $[5,17]$.

### 3.2 Improvement of the Error Estimation

We improve the error estimation of Raz's extractor as follows. Unlike [5, 17], we present another partition method of the sum in the following proof. The combination and permutation formulas are exploited to show a tight bound of the sum. Correspondingly, the error can be reduced.
Theorem 2. Let $D=2^{d}$. Let $Z_{1}, \cdots, Z_{D}$ be $0-1$ random variables that are $\epsilon$ biased for linear tests of size $k^{\prime}$ that are constructed using $n$ random bits. Define Ext ${ }^{(1)}:\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}$ by Ext ${ }^{(1)}(w, s)=Z_{s}(w)$, that is, Ext ${ }^{(1)}(w, s)$ is the random variable $Z_{s}$, when using $w$ as the value of the $n$ bits needed to produce $Z_{1}, \cdots, Z_{D}$. Then, for any even integer $k=k^{\prime}$, the function Ext ${ }^{(1)}$ is a $(\alpha, \gamma)$-seeded-extractor, with

$$
\gamma=2^{\frac{n-\alpha}{k}} \cdot\left[2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{k}} .
$$

Proof. We improve the proof by proposing another method for partitioning the sum $\sum_{s_{1}, \cdots, s_{k} \in\{0,1\}^{d}} \sum_{w \in\{0,1\}^{n}} \prod_{j=1}^{k} e\left(w, s_{j}\right)$ into two sums. The first sum is over $s_{1}, \cdots, s_{k} \in\{0,1\}^{d}$ such that in each summand, at least one $s_{j}$ appears odd times in $\left\{s_{1}, \cdots, s_{k}\right\}$, and the second sum is over $s_{1}, \cdots, s_{k} \in\{0,1\}^{d}$ such that in each summand every $s_{j}$ appears even times in $\left\{s_{1}, \cdots, s_{k}\right\}$. Then the number of summands in the second sum is
and each summand is $2^{n}$. Therefore, the number of summands in the first sum is

$$
2^{d k}-2^{\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1,
$$

and by Claim 1 each summand is bounded by $2^{n} \cdot \epsilon$. Hence,

$$
\begin{aligned}
& 2^{\alpha} \cdot 2^{d \cdot k} \cdot \gamma(W, S)^{k} \\
\leq & 2^{n} \cdot\left[2^{\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdot 1\right]+2^{n} \cdot \epsilon \cdot\left[2^{d k}-2^{\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdot 1\right]
\end{aligned}
$$

Therefore,

$$
\begin{aligned}
\gamma(W, S)^{k} & \leq \frac{2^{n} \cdot 2^{d k}}{2^{\alpha} \cdot 2^{d \cdot k}} \cdot\left[2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdot 1 \cdot(1-\epsilon)+\epsilon\right] \\
& =2^{n-\alpha} \cdot\left[2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots \cdot 1 \cdot(1-\epsilon)+\epsilon\right]
\end{aligned}
$$

Consequently, we have

$$
|\gamma(W, S)| \leq 2^{\frac{n-\alpha}{k}} \cdot\left[2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{k}} .
$$

### 3.3 Comparison

For simplicity, in the rest of the paper, denote $\gamma_{1}$ as the error of the extractor in Lemma 1, and $\gamma_{2}$ as the counterpart in Theorem 2.
Proposition 1. $2^{\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1 \leq 2^{d \cdot \frac{k}{2}} \cdot\left(\frac{k}{2}\right)^{k}$ for any positive even integer $k$, and "=" holds iff $k=2$.
Proof. When $k=2$, it's trivial that $2^{\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1=2^{d \cdot \frac{k}{2}} \cdot\left(\frac{k}{2}\right)^{k}$. In the following, we only consider any positive even integer $k$ with $k>2$.

Since $\frac{k!}{\left(\frac{k}{2}\right)!}<\frac{k^{k}}{2^{\frac{k}{2}}}$, we have $\frac{k!}{\left(\frac{k}{2}\right)!\cdot 2^{\frac{k}{2}}}<\frac{k^{k}}{2^{k}}$. Hence,

$$
2^{\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots \cdot 1=2^{\frac{d k}{2}} \cdot \frac{k!}{\left(\frac{k}{2}\right)!\cdot 2^{\frac{k}{2}}}<2^{\frac{d k}{2}} \cdot \frac{k^{k}}{2^{k}}
$$

- When $k$ is large enough, $\left(\frac{k}{2}\right)^{k}$ is much greater than $(k-1) \cdot(k-3) \cdots \cdots 1$. For instance, when $k=6$, we have $\left(\frac{k}{2}\right)^{k}=729$ and $(k-1) \cdot(k-3) \cdots \cdots 1=15$. Therefore, "The number of summands in the second sum is $2^{\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots$ 1 , and each summand is $2^{n}$." in the proof of Theorem 2 is a great improvement on "The number of summands in the second sum is bounded by $2^{d \cdot \frac{k}{2}} \cdot\left(\frac{k}{2}\right)^{k}$, and each summand is trivially bounded by $2^{n}$." in the proof of Lemma 1 .
- The error estimation of the extractor in Theorem 1 is better than that in Lemma 1. $\gamma_{1}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{k}} \cdot(2 \epsilon)^{\frac{1}{k}}$ in [5] under the assumption that $\epsilon \geq 2^{-\frac{d k}{2}} \cdot k^{k}$ and $0<\delta<\frac{1}{2}$, and $\gamma_{2}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{k}} \cdot\left[2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{k}}=$ $2^{\frac{\left(\frac{1}{2}-\delta\right) n}{k}} \cdot\left\{2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1+\left[1-2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1\right] \cdot \epsilon\right\}^{\frac{1}{k}}$. Since $\epsilon \geq 2^{-\frac{d k}{2}} \cdot k^{k}$ and $2^{-\frac{d k}{2}} \cdot k^{k}>2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1$ for any even integer $k$, we get $\gamma_{1}>\gamma_{2}$. Moreover, according to the Stirling's Formula, $2^{-\frac{d k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1=2^{-\frac{d k}{2}} \cdot \frac{k!}{\left(\frac{k}{2}\right)!\cdot 2^{\frac{k}{2}}} \approx 2^{-\frac{d k}{2}} \cdot \frac{k^{k}}{\left(\frac{k}{2}\right)^{\frac{k}{2}} \cdot 2^{\frac{k}{2}}}=2^{-\frac{d k}{2}} \cdot \sqrt{k}^{k}$.
- If $\epsilon \geq \frac{1}{2^{\left(\frac{1}{2}-\delta\right) n+1}}$, then $\gamma_{1}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{k}} \cdot(2 \epsilon)^{\frac{1}{k}} \geq 1$. In this case, the error estimation is meaningless.
- To simplify $\gamma_{2}$, let $k$ be a specific value. For instance, let $k=4$, then the error $\gamma_{2}=2^{\frac{\left(\frac{1}{2}-\delta\right) n}{4}} \cdot\left[2^{-2 d} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}}$.


### 3.4 Important Role

It should be noticed that $\gamma$ constrains on the parameters in Theorem 1. The main idea of the proof about Theorem 1 is as follows. Assuming for contradiction that Ext is not a non-malleable extractor, then after some steps, an inequality $\gamma_{1}>A$ is deduced, where $A$ denotes a certain formula. On the other hand, from the assumption of Theorem $1, \gamma_{1}<A$ should hold. Thus Ext is a non-malleable extractor. Essentially, the constraints on the parameters in Theorem 1 are chosen according to the inequality $\gamma_{1}<A$. From Proposition 1, we have $\gamma_{1}>\gamma_{2}$ for any positive even integer $k \geq 4$. Therefore, we may relax the constraints on the parameters in Theorem 1 according to $\gamma_{2}<A$. Correspondingly, the seed length may be further shortened.

## 4 Explicit Constructions of Non-malleable Extractors with Shorter Seed Length

In this section, we propose the explicit construction of a non-malleable extractor, which has shorter seed length than that in Theorem 1. Then we give a general explicit construction of non-malleable extractors, and analyze the simplification of the constraints on the parameters.

### 4.1 Constructions

We first review two lemmas that will be used later.
Lemma 2. (see [5]) Let $X$ be a random variable over $\{0,1\}^{m}$. Let $Y, Z$ be two random variables. Then,

$$
\left\|(X, Y, Z)-\left(U_{m}, Y, Z\right)\right\|_{1}=\mathbb{E}_{z \sim Z}\left\|\left.(X, Y)\right|_{Z=z}-\left.\left(U_{m}, Y\right)\right|_{Z=z}\right\|_{1} .
$$

Lemma 3. (see [5]) Let $X$ be a random variable over $\{0,1\}^{m}$. Let $Y$ be a random variable over $\{0,1\}^{n}$. Then

$$
\left\|(X, Y)-\left(U_{m}, Y\right)\right\|_{1} \leq \sum_{\emptyset \neq \sigma \subseteq[m], \tau \subseteq[n]} \operatorname{bias}\left(X_{\sigma} \oplus Y_{\tau}\right),
$$

where $X_{i}$ is the $i$ th bit of $X, Y_{j}$ is the $j$ th bit of $Y, X_{\sigma}=\oplus_{i \in \sigma} X_{i}$, and $Y_{\tau}=$ $\oplus_{j \in \tau} Y_{j}$.
Theorem 3. Let $n=2^{10}, d=19, m=1, \alpha=1016, \xi=2^{\theta}, \theta=-1$, and $t=1$. Then there exists an explicit $(\alpha, \xi)-t$-non-malleable extractor nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$.

Proof Idea. We borrow the reductio ad absurdum approach in the proof of Theorem 1. The first part of the proof is almost the same as that in [5]. The second part jumps out of the idea in [5]. We exploit the error estimation of the extractor in Theorem 2 instead of Lemma 1. We use a trick such that the
even integer $k$ is just 4 instead of the largest even integer that is not larger than $\frac{\left\lceil\frac{\delta n}{8}\right\rceil}{(t+1) m}$. Therefore the extractor error can be simplified and complex assumptions to satisfy $k \cdot\left(\frac{1}{\epsilon}\right)^{\frac{1}{k}} \leq\left(m \cdot 2^{d}\right)^{\frac{1}{2}}$ is omitted.

Proof. The explicit construction we present is the extractor constructed in [17]. Alon et al. [1] observed that for every $k^{\prime}, N \geq 2$, the sequence of $0-1$ random variables $Z_{1}, \cdots, Z_{N}$ that is $\epsilon$-biased for linear tests of size $k^{\prime}$ can be explicitly constructed using $2 \cdot\left\lceil\log (1 / \epsilon)+\log k^{\prime}+\log \log N\right\rceil$ random bits. Therefore, let $D=2^{d}$ and $\epsilon=2^{-\frac{n}{2}+r}$ with $r=1+\log k^{\prime}+\log \log D$, then we can construct a sequence of $0-1$ random variables $Z_{1}, \cdots, Z_{D}$ that is $\epsilon$-biased for linear tests of size $k^{\prime}$ by using $n$ random bits. Let $k^{\prime}=8$. Define Ext : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ by $\operatorname{Ext}(w, s)=Z_{s}(w)$.

Let $S$ be a random variable uniformly distributed over $\{0,1\}^{d}$.
Assume for contradiction that Ext is not a (1016, $\frac{1}{2}$ )-non-malleable-extractor. Then there exists a source $W$ of length $n$ with min-entropy $\alpha$, and an adversarialfunction $\mathcal{A}:\{0,1\}^{d} \rightarrow\{0,1\}^{d}$ such that

$$
\left\|(\operatorname{Ext}(W, S), \operatorname{Ext}(W, \mathcal{A}(S)), S)-\left(U_{m}, \operatorname{Ext}(W, \mathcal{A}(S)), S\right)\right\|_{1}>\xi=2^{\theta}
$$

As in [3], we assume that $W$ is uniformly distributed over a set $W^{\prime} \subseteq\{0,1\}^{n}$ of size $\alpha$.

For every $s \in\{0,1\}^{d}$, let $X_{s}$ be the random variable $\operatorname{Ext}(W, s)$. By Lemma 2 and 3 , we have

$$
\sum_{\emptyset \neq \sigma \subseteq[1], \tau \subseteq[1]} \mathbb{E}_{s \sim S}\left[\operatorname{bias}\left(\left(X_{s}\right)_{\sigma} \oplus\left(X_{\mathcal{A}(s)}\right)_{\tau}\right)\right]>2^{\theta}
$$

Let $\sigma^{*}, \tau^{*} \subseteq[1]$ be the indices of (one of) the largest summands in the above sum. For every $s \in\{0,1\}^{d}$, let

$$
Y_{s}=\left(X_{s}\right)_{\sigma^{*}} \oplus\left(X_{\mathcal{A}(s)}\right)_{\tau^{*}}
$$

There is a set $S^{\prime \prime} \subseteq\{0,1\}^{d}$ satisfying that $\left|S^{\prime \prime}\right|>\frac{\xi \cdot 2^{d-2}}{2^{m t}\left(2^{m}-1\right)(t+1)^{2}}=\xi \cdot 2^{d-5}$. The $S^{\prime \prime}$ here is the same as that in the proof of Theorem 1 by replacing $t$ there with 1 and the error $2^{-m}$ there with $\xi$. Please see [5] for details.

Define a random variable $Y_{S^{\prime \prime}}$ over $\{0,1\}$ as follows: To sample a bit from $Y_{S^{\prime \prime}}$, uniformly sample a string $s$ from $S^{\prime \prime}$, and then independently sample a string $w$ uniformly from $W^{\prime}$. The sampled value is $Y_{s}(w)$. We have that $\operatorname{bias}\left(Y_{S^{\prime \prime}}\right)>$ $\frac{\xi}{2^{m t+1}\left(2^{m}-1\right)(t+1)}=\frac{\xi}{2^{3}}$. For every $s \in S^{\prime \prime}$, let

$$
Y_{s}^{\prime}=Z_{(1, s)} \oplus\left(\oplus_{j \in \tau^{*}} Z_{(j, \mathcal{A}(s))}\right),
$$

where $Z_{(1, s)}=Z_{s}$.
Let $t=1$ and $m=1$ in Claim 7.2 of [5], we get the following claim.
Claim 2. The set of random variables $\left\{Y_{s}^{\prime}\right\}_{s \in S^{\prime \prime}} \epsilon-$ fools linear tests of size $\frac{k^{\prime}}{2}$.

We apply Theorem 2 on the random variables $\left\{Y_{s}^{\prime}\right\}_{s \in S^{\prime \prime}}$. For simplicity of presentation we assume $\left|S^{\prime \prime}\right|=2^{d^{\prime}}$. By Theorem 2, the distribution of Ext ${ }^{(1)}\left(W, S^{\prime \prime}\right)$ is $\gamma$-biased for $\gamma=2^{\frac{n-\alpha}{k}} \cdot\left[2^{-\frac{d^{\prime} k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{k}}$. Let $k=\frac{k^{\prime}}{2}=4$, then $\gamma=2^{\frac{n-\alpha}{4}} \cdot\left[2^{-2 d^{\prime}} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}}$. We note that Ext ${ }^{(1)}\left(W, S^{\prime \prime}\right)$ has the same distribution as $Y_{S^{\prime \prime}}$. In particular, both random variables have the same bias. Therefore, we get

$$
2^{\frac{n-\alpha}{4}} \cdot\left[2^{-2 d^{\prime}} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}} \geq \operatorname{bias}\left(Y_{S^{\prime \prime}}\right)>\frac{\xi}{2^{3}}
$$

Moreover, since $2^{d^{\prime}}=\left|S^{\prime \prime}\right|>\xi \cdot 2^{d-5}$, we have

$$
2^{\frac{n-\alpha}{4}} \cdot\left[\xi^{-2} \cdot 2^{-2 d+10} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}}>2^{\frac{n-\alpha}{4}} \cdot\left[2^{-2 d^{\prime}} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}}>\frac{\xi}{2^{3}}
$$

That is,

$$
\begin{equation*}
2^{-2 d}>\frac{\xi^{4} \cdot 2^{-12-n+\alpha}-\epsilon}{3(1-\epsilon) \cdot 2^{10} \cdot \xi^{-2}} \tag{a}
\end{equation*}
$$

where $\epsilon=2^{-\frac{n}{2}+r}$ and $r=4+\log d$.
On the other hand, from the assumption that $n=2^{10}, \alpha=1016, d=19$, $m=1, \xi=2^{\theta}, \theta=-1$, we have

$$
2^{-2 d}<\frac{\xi^{4} \cdot 2^{-12-n+\alpha}-\epsilon}{3(1-\epsilon) \cdot 2^{10} \cdot \xi^{-2}}
$$

which is in contradiction to the inequality (a).

Comparison. In Theorem 1, the seed length $d$ and the source length $n$ should satisfy $d \geq \frac{23}{\delta} m+2 \log n$ with $0<\delta<\frac{1}{2}$. However, in the above construction, we have $d=1.9 \log n$. In other words, let $n=2^{10}, m=1, \delta=\frac{504}{1024}$, and $t=1$ in Theorem 1 , then it can be easily verified that $n \geq \frac{160}{\delta} \cdot t m$. To construct an explicit $\left(\left(\frac{1}{2}+\delta\right) \cdot n, 2^{-m}\right)-t$-non-malleable extractor nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ ( that is, an explicit (1016, $\frac{1}{2}$ )-1-non-malleable extractor nmExt), by Theorem 1 , the seed length $d$ should satisfy $d \geq \frac{23}{\delta} \cdot t m+2 \log n=\frac{46}{63}+66$. Moreover, when $d \leq 2^{41}$, the precondition $\delta \geq 10 \cdot \frac{\log (n d)}{n}$ in Theorem 1 is satisfied. Meanwhile, by Theorem 3, the seed length $d$ can just be 19. In this sense, our construction is much better than that of [5].
Theorem 4. Assuming
$0<2^{\log 3-2 \theta+4 m+8}-2^{\log 3-\frac{n}{2}+4+\log d-2 \theta+4 m+8} \leq 2^{2 d+4 \theta-8 m-8-n+\alpha}-2^{2 d-\frac{n}{2}+4+\log d}$.
Then there exists an explicit $\left(\alpha, \xi=2^{\theta}\right)$ - 1-non-malleable extractor nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$.

Proof. The proof is similar to Theorem 3. Please see the appendix for details.

Analysis of the Assumption. In order to construct an explicit non-malleable extractor, it's enough to guarantee that the parameters satisfies
$0<2^{\log 3} \cdot\left(1-2^{-\frac{n}{2}+4+\log d}\right) \cdot 2^{-2 \theta+4 m+8} \leq 2^{2 d+4 \theta-8 m-8-n+\alpha}-2^{2 d-\frac{n}{2}+4+\log d}$. (b)
For simplicity, denote

$$
\begin{aligned}
& A^{\prime}=\log 3-2 \theta+4 m+8, B^{\prime}=\log 3-\frac{n}{2}+4+\log d-2 \theta+4 m+8, \\
& C^{\prime}=2 d+4 \theta-8 m-8-n+\alpha, D^{\prime}=2 d-\frac{n}{2}+4+\log d,
\end{aligned}
$$

then

$$
\text { (b) holds. } \Leftrightarrow 0<2^{A^{\prime}}-2^{B^{\prime}} \leq 2^{C^{\prime}}-2^{D^{\prime}} \text {. }
$$

In the following, we discuss what happens under the assumption (b) in three cases.

Case 1. Assuming $A^{\prime} \geq C^{\prime}$ and $B^{\prime} \geq D^{\prime}$. Since " $B^{\prime} \geq D^{\prime}$ " implies " $A^{\prime} \geq$ $C^{\prime \prime \prime}$, we only need to consider $B^{\prime} \geq D^{\prime}$ (i.e., $\log 3-2 \theta+4 m+8 \geq 2 d$ ). Let $1-\epsilon=1-2^{-\frac{n}{2}+4+\log d}=2^{\rho^{\prime}}$.

From $\log 3+8+4 m \geq 2 d+2 \theta, \alpha \leq n, m \geq 1$, and $\theta<0$, we get

$$
\begin{aligned}
-16 & >-8 m-8+4 \theta-n+\alpha \\
& =(\log 3+8+4 m)+4 \theta-12 m-16-\log 3-n+\alpha \\
& \geq 2 d+2 \theta+4 \theta-12 m-16-\log 3-n+\alpha
\end{aligned}
$$

Let $\rho^{\prime} \geq-16$. Then we have $\rho^{\prime}>2 d+2 \theta+4 \theta-12 m-16-\log 3-n+\alpha$.
Therefore, $\log 3+\rho^{\prime}-2 \theta+4 m+8>2 d+4 \theta-8 m-8-n+\alpha$, which is in contradiction to the inequality (b).

Consequently, when $\epsilon \in\left(0,1-2^{-16}\right], A^{\prime} \geq C^{\prime}$, and $B^{\prime} \geq D^{\prime}$, (b) does not hold. From Theorem 2, only if $\epsilon$ is small enough, the corresponding seeded extractor is useful. Therefore, we assume that $\epsilon \in\left(0,1-2^{-16}\right]$.

Case 2. Assuming $A^{\prime} \geq C^{\prime}$ and $B^{\prime}<D^{\prime}$, then it's in contradiction to the inequality (b).

Case 3. Assuming $A^{\prime}<C^{\prime}$, then it's trivial that $B^{\prime}<D^{\prime}$. Thus, we only need to consider $A^{\prime}<C^{\prime}$. Since $A^{\prime}>B^{\prime}$, we have $C^{\prime}>D^{\prime}$, that is, $4 \theta-8 m-$ $12-\frac{n}{2}+\alpha>\log d$.

Therefore, we obtain the following theorem.
Theorem 5. Assuming $\epsilon=2^{-\frac{n}{2}+4+\log d} \in\left(0,1-2^{-16}\right]$ and
$2^{\log 3} \cdot\left(1-2^{-\frac{n}{2}+4+\log d}\right) \cdot 2^{-2 \theta+4 m+8} \leq 2^{2 d+4 \theta-8 m-8-n+\alpha}-2^{2 d-\frac{n}{2}+4+\log d}$.
Then there exists an explicit $\left(\alpha, \xi=2^{\theta}\right)$ - 1-non-malleable extractor nmExt : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$.

In particular, the parameters of the non-malleable extractor can be chosen according to the inequality system

$$
\left\{\begin{array}{l}
\log 3-6 \theta+16+12 m+n-\alpha<2 d  \tag{1}\\
4 \theta-8 m-12-\frac{n}{2}+\alpha>\log d \\
2^{-\frac{n}{2}+4+\log d} \leq 1-2^{-16}
\end{array}\right.
$$

then check whether they satisfy the inequality
$2^{\log 3-2 \theta+4 m+8}-2^{\log 3-\frac{n}{2}+4+\log d-2 \theta+4 m+8} \leq 2^{2 d+4 \theta-8 m-8-n+\alpha}-2^{2 d-\frac{n}{2}+4+\log d}$.
Remark. $\alpha$ can't be less than $\frac{n}{2}$, since $4 \theta-8 m-12-\frac{n}{2}+\alpha>\log d$. We can also replace $\alpha$ with $\left(\frac{1}{2}+\delta\right) n$ where $0<\delta<\frac{1}{2}$.

## 5 Non-Malleable Extractors for Entropy Rate Less than $\frac{1}{2}$

In the following, we give an explicit construction of non-malleable extractors for min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ for some constant $\delta>0$ by employing a special encoding technique in $[2,11,16]$, while the min-entropy $k$ should be greater than $\frac{1}{2} n$ in [5]. In this technique, for a weak source $X$ with min-entropy rate $\frac{1}{2}-$ $\delta$ for some constant $\delta>0, X$ is encoded into $\operatorname{Enc}(X)$ such that $3 \operatorname{Enc}(X)$ is statistically close to having min-entropy rate $\frac{1}{2}+\delta$. Unlike Li's construction in FOCS'12 by using inner product function [11], our construction is based on $\epsilon$ biased variable sequence for linear tests. Furthermore, we propose a high level way of constructing non-malleable extractors with min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ from any non-malleable extractor with min-entropy $k>\frac{1}{2} n$.

Now we recall the way to encode a source $[2,11,16]$ as follows. Consider the finite field $\mathbb{F}_{2^{p}}$ where $p$ is a prime. Let $g$ be a generator in $\mathbb{F}_{2^{p}}^{*}$. Define Enc : $\mathbb{F}_{2^{p}}^{*} \rightarrow \mathbb{F}_{2^{p}}^{*} \times \mathbb{F}_{2^{p}}^{*}$ by $\operatorname{Enc}(x)=\left(x, g^{x}\right)$. Each element in the support of $3 \operatorname{Enc}(X)$ has the form $\left(x_{1}+x_{2}+x_{3}, g^{x_{1}}+g^{x_{2}}+g^{x_{3}}\right)$, where $x_{1}, x_{2}, x_{3} \in \mathbb{F}_{2^{p}}^{*}$. From [2, 11, 16], we get
Lemma 4. Let $\alpha_{0}>0$ be the constant in Theorem 3.21 of [11] ${ }^{2}$ and $\delta=\frac{\alpha_{0}}{8+2 \alpha_{0}}$. Assume that $X$ is any weak source with min-entropy $k=\left(\frac{1}{2}-\delta\right) p$, then $3 \operatorname{Enc}(X)$ is $2^{1-\frac{\alpha_{0} k}{2}}$-close to a source with min-entropy $\left(\frac{1}{2}+\delta\right)(2 p)$.

Proof. Recall that it has been proved that $3 \mathrm{Enc}(X)$ is $2^{1-\frac{\alpha_{0} k}{2}}$-close to having min-entropy $\left(1+\frac{\alpha_{0}}{2}\right) \cdot 2 \cdot\left(\frac{1}{2}-\delta\right) p$ in the proof of Claim 6.4 in [11] under the assumption that $p=n$. Let $\left(\frac{1}{2}+\delta\right)(2 p)=\left(1+\frac{\alpha_{0}}{2}\right) \cdot 2 \cdot\left(\frac{1}{2}-\delta\right) p$, then $\delta=\frac{\alpha_{0}}{8+2 \alpha_{0}}$.

Therefore, the explicit construction in Theorem 5 can be modified to break the $\frac{1}{2}$ entropy rate barrier as follows.

[^1]Theorem 6. Let $\alpha_{0}>0$ be the constant in Theorem 3.21 of [11] and $\delta=\frac{\alpha_{0}}{8+2 \alpha_{0}}$. Let $\left\{Z_{(i, s)} \mid i \in[m], s \in\{0,1\}^{d}\right\}$ be a sequence of $0-1$ random variables that is $\epsilon$-biased for linear tests of size $k^{\prime}$ that are constructed using $2 p$ random bits.

Assuming Ext : $\{0,1\}^{2 p} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ defined by

$$
\operatorname{Ext}(w, s)=Z_{(1, s)}(w)\left\|Z_{(2, s)}(w) \ldots\right\| Z_{(m, s)}(w)
$$

is a $\left(\left(\frac{1}{2}+\delta\right)(2 p), \xi\right)$-1-non-malleable extractor, then Ext $:\{0,1\}^{p} \times\{0,1\}^{d} \rightarrow$ $\{0,1\}^{m}$ defined by $E_{x t} t^{\prime}(x, s)=\operatorname{Ext}(3 E n c(x), s)$ is a $\left(\left(\frac{1}{2}-\delta\right) p, 2^{2-\frac{\alpha_{0} k}{2}}+\xi\right)$-1-nonmalleable extractor.

Proof. From Lemma 4, there exists a source $W$ with min-entropy $\left(\frac{1}{2}+\delta\right)(2 p)$ such that $\|3 \operatorname{Enc}(X)-W\|_{1} \leq 2^{1-\frac{\alpha_{0} k}{2}}$. Hence,
$S D\left(\left(\operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), S), \operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), \mathcal{A}(S))\right),\left(\operatorname{Ext}^{\prime}(W, S), \operatorname{Ext}^{\prime}(W, \mathcal{A}(S))\right) \mid S\right) \leq 2^{-\frac{\alpha_{0} k}{2}}$.
Moreover, we have

$$
\begin{gathered}
S D\left(\left(\operatorname{Ext}^{\prime}(W, S), \operatorname{Ext}^{\prime}(W, \mathcal{A}(S))\right),\left(U_{m}, \operatorname{Ext}^{\prime}(W, \mathcal{A}(S))\right) \mid S\right) \leq \frac{1}{2} \xi \\
S D\left(\left(U_{m}, \operatorname{Ext}^{\prime}(W, \mathcal{A}(S))\right),\left(U_{m}, \operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), \mathcal{A}(S))\right) \mid S\right) \leq 2^{-\frac{\alpha_{0} k}{2}}
\end{gathered}
$$

Therefore, we get

$$
\begin{aligned}
& S D\left(\left(\operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), S), \operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), \mathcal{A}(S))\right),\left(U_{m}, \operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), \mathcal{A}(S))\right) \mid S\right) \\
& \leq 2^{1-\frac{\alpha_{0} k}{2}}+\frac{1}{2} \xi
\end{aligned}
$$

Consequently,
$\left\|\left(\operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), S), \operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), \mathcal{A}(S)), S\right)-\left(U_{m}, \operatorname{Ext}^{\prime}(3 \operatorname{Enc}(X), \mathcal{A}(S)), S\right)\right\|_{1}$ $\leq 2^{2-\frac{\alpha_{0} k}{2}}+\xi$.

More generally, we have the following result, which proposes a high level way of constructing non-malleable extractors with min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ for some $0<\delta<\frac{1}{2}$.
Theorem 7. Let $\alpha_{0}>0$ be the constant in Theorem 3.21 of [11] and $\delta=$ $\frac{\alpha_{0}}{8+2 \alpha_{0}}$. Assuming Ext : $\{0,1\}^{2 p} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ is a $\left(\left(\frac{1}{2}+\delta\right)(2 p), \xi\right)$-1-nonmalleable extractor, then Ext $:\{0,1\}^{p} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ defined by Ext $(x, s)=$ $\operatorname{Ext}(3 \operatorname{Enc}(x), s)$ is a $\left(\left(\frac{1}{2}-\delta\right) p, 2^{2-\frac{\alpha_{0} k}{2}}+\xi\right)$-1-non-malleable extractor.

## 6 Application to Privacy Amplification

In this section, we show how the non-malleable extractor is applied to the privacy amplification protocol of Dodis and Wichs [7, 8] (also known as an informationtheoretic key agreement protocol).

Roughly speaking, in this scenario, Alice and Bob share a shared weak secret $W$, the min-entropy of which is only guaranteed. They communicate over a public and unauthenticated channel to securely agree on a nearly uniform secret key $R$, where the attacker Eve is active and computationally unbounded. To achieve this goal, the protocol is designed as follows. Assuming we'll authenticate the

| Alice: W | Eve | Bob: W |
| :---: | :---: | :---: |
| Sample random S. |  |  |
|  | $S \longrightarrow S^{\prime}$ |  |
|  | $\left(S_{0}^{\prime}, T_{0}^{\prime}\right) \longleftarrow\left(S_{0}, T_{0}\right)$ | Sample random $S_{0}$. $\begin{gathered} R^{\prime}=\operatorname{nmExt}\left(W, S^{\prime}\right) \\ T_{0}=\operatorname{MAC}_{R^{\prime}}\left(S_{0}\right) . \end{gathered}$ <br> Reach KeyDerived state. <br> Output $R_{B}=\operatorname{Ext}\left(W, S_{0}\right)$. |
| $R=\mathrm{nmExt}(W, S)$. |  |  |
| If $T_{0}^{\prime} \neq \mathrm{MAC}_{R}\left(S_{0}^{\prime}\right)$, output $R_{A}=\perp$. |  |  |
| Otherwise, reach KeyConfirmed state, and output $R_{A}=\operatorname{Ext}\left(W, S_{0}^{\prime}\right)$. |  |  |

Table 1. The Dodis-Wichs privacy amplification protocol.
seed $S_{0}$. Alice initiates the conversation by transmitting a uniformly random seed $S$ to Bob. During this transmission, $S$ may be modified by Eve into any value $S^{\prime}$. Then Bob samples a uniform seed $S_{0}$, computes the authentication key $R^{\prime}=\operatorname{nmExt}\left(W, S^{\prime}\right)$, and sends $S_{0}$ together with the authentication tag $T_{0}=$ $\mathrm{MAC}_{R^{\prime}}\left(S_{0}\right)$ to Alice. At this point, Bob reaches the KeyDerived state and outputs $R_{B}=\operatorname{Ext}\left(W, S_{0}\right)$. During this transmission, $\left(S_{0}, T_{0}\right)$ may be modified by Eve into any pair $\left(S_{0}^{\prime}, T_{0}^{\prime}\right)$. Alice computes the authentication key $R=\mathrm{nmExt}(W, S)$ and verifies that $T_{0}^{\prime}=\operatorname{MAC}_{R}\left(S_{0}^{\prime}\right)$. If the verification fails then Alice rejects and outputs $R_{A}=\perp$. Otherwise, Alice reaches the KeyConfirmed state and outputs $R_{A}=\mathrm{nmExt}\left(W, S_{0}^{\prime}\right)$.

The security can be analyzed in two cases [5, 7]. Case 1: Eve does not modified the seed $S$ in the first round. Then Alice and Bob share the same authentication key (i.e., $R^{\prime}=R$ ), which is statistically close to a uniform distribution. Therefore, Eve has only a negligible probability of getting a valid authentication tag $T_{0}^{\prime}$ for any seed $S_{0}^{\prime} \neq S_{0}$. Case 2 : Eve does modify the seed $S$ to a different seed $S^{\prime}$. Since $T_{0}$ is a deterministic function of $S_{0}$ and $R^{\prime}$, Eve may guess $R^{\prime}$. According to the definition of non-malleable extractors, the authentication key $R$ computed by Alice is still statistically close to a uniform distribution. Thus, again, the adversary has only a negligible probability of computing a valid authentication $T_{0}^{\prime}$ for any seed $S_{0}^{\prime}$ with respect to the authentication key $R$. Consequently, the above protocol is secure. The formal definition of privacy amplification protocol is given in the appendix.

Theorem 8. (see $[5,8])$ Assume that nmExt: $\{0,1\}^{n} \times\{0,1\}^{d_{1}} \rightarrow\{0,1\}^{m_{1}}$ is a $\left(k, \epsilon_{\text {nmExt }}\right)$-non-malleable extractor, Ext : $\{0,1\}^{n} \times\{0,1\}^{d_{2}} \rightarrow\{0,1\}^{m_{2}}$ is a strong $\left(k-\left(d_{1}+m_{1}\right)-\log \frac{1}{\epsilon^{\prime}}, \epsilon_{E x t}\right)$-extractor, and $\left\{M A C_{r}:\{0,1\}^{d_{2}} \rightarrow\{0,1\}^{\tau}\right\}_{r \in\{0,1\}^{m_{1}}}$ is a $\varepsilon_{M A C}$-secure message authentication code. Then for any integers $n$ and $k<n$, the protocol in Table 1 is a 2-round ( $n, k, m, \delta$ )-privacy amplification protocol, with communication complexity $d_{1}+d_{2}+\tau$, where $\delta=\max \left\{\epsilon^{\prime}+\epsilon_{\text {Ext }}, \epsilon_{n m E x t}+\right.$ $\left.\varepsilon_{M A C}\right\}$.

The explicit non-malleable extractor in this paper can be applied to construct the above privacy amplification protocol with low communication complexity.

## 7 Conclusion

Non-malleable extractor is a powerful theoretical tool to study privacy amplification protocols, where the attacker is active and computationally unbounded.

In this paper, we improved the error estimation of Raz's extractor. Based on the improvement, we presented an improved explicit construction of nonmalleable extractors with shorter seed length. More precisely, we constructed an explicit (1016, $\frac{1}{2}$ )-1-non-malleable extractor nmExt : $\{0,1\}^{2^{10}} \times\{0,1\}^{d} \rightarrow\{0,1\}$ with seed length 19 , while it should be no less than $\frac{46}{63}+66$ according to Cohen et al. in CCC'12 [5]. We also gave a general explicit construction of non-malleable extractors and analyzed the simplification of the constraints on the parameters. How to further simplify the constraints is an open problem.

Furthermore, we gave an explicit construction of non-malleable extractors for min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ for some constant $\delta>0$, while the min-entropy $k$ should be greater than $\frac{1}{2} n$ by Cohen et al. in CCC'12 [5]. By employing a special encoding technique, we also proposed a high level way to construct non-malleable extractors with min-entropy $k=\left(\frac{1}{2}-\delta\right) n$ from any non-malleable extractor with min-entropy $k>\frac{1}{2} n$. Compared with Li's construction in FOCS'12 [11] by using inner product function, our construction is more general. Finally, we showed their application to the privacy amplification protocol (or information-theoretic key agreement protocol).

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## A

Theorem 4. Assuming
$0<2^{\log 3-2 \theta+4 m+8}-2^{\log 3-\frac{n}{2}+4+\log d-2 \theta+4 m+8} \leq 2^{2 d+4 \theta-8 m-8-n+\alpha}-2^{2 d-\frac{n}{2}+4+\log d}$.
Then there exists an explicit $\left(\alpha, \xi=2^{\theta}\right)$-non-malleable extractor $n m E x t:\{0,1\}^{n} \times$ $\{0,1\}^{d} \rightarrow\{0,1\}^{m}$.

Proof. The explicit construction we present is the extractor constructed in [17]. Alon et al. [1] observed that for every $k^{\prime}, N \geq 2$, the sequence of $0-1$ random variables $Z_{1}, \cdots, Z_{N}$ that is $\epsilon$-biased for linear tests of size $k^{\prime}$ can be explicitly constructed using $2 \cdot\left\lceil\log (1 / \epsilon)+\log k^{\prime}+\log \log N\right\rceil$ random bits. Therefore, let $D=m \cdot 2^{d}$ and $\epsilon=2^{-\frac{n}{2}+r}$ with $r=1+\log k^{\prime}+\log \log D$, then we can construct a sequence of $0-1$ random variables $Z_{1}, \cdots, Z_{D}$ that is $\epsilon$-biased for linear tests of size $k^{\prime}$ by using $n$ random bits. Let $k^{\prime}=8 m$. We interpret the set of indices $[D]$ as the set $\left\{(i, s): i \in[m], s \in\{0,1\}^{d}\right\}$. Define Ext : $\{0,1\}^{n} \times\{0,1\}^{d} \rightarrow\{0,1\}^{m}$ by

$$
\operatorname{Ext}(w, s)=Z_{(1, s)}(w)\left\|Z_{(2, s)}(w) \ldots\right\| Z_{(m, s)}(w)
$$

where "||" is the concatenation operator.
Let $S$ be a random variable uniformly distributed over $\{0,1\}^{d}$.
Assume for contradiction that Ext is not a $\left(\alpha, 2^{\theta}\right)$-non-malleable-extractor. Then there exists a source $W$ of length $n$ with min-entropy $\alpha$, and an adversarialfunction $\mathcal{A}:\{0,1\}^{d} \rightarrow\{0,1\}^{d}$ such that

$$
\left\|(\operatorname{Ext}(W, S), \operatorname{Ext}(W, \mathcal{A}(S)), S)-\left(U_{m}, \operatorname{Ext}(W, \mathcal{A}(S)), S\right)\right\|_{1}>\xi=2^{\theta}
$$

As in [3], we may assume that $W$ is uniformly distributed over a set $W^{\prime} \subseteq\{0,1\}^{n}$ of size $\alpha$.

For every $s \in\{0,1\}^{d}$, let $X_{s}$ be the random variable $\operatorname{Ext}(W, s)$. By Lemma 2 and 3, we have

$$
\sum_{\emptyset \neq \sigma \subseteq[m], \tau \subseteq[m]} \mathbb{E}_{s \sim S}\left[\operatorname{bias}\left(\left(X_{s}\right)_{\sigma} \oplus\left(X_{\mathcal{A}(s)}\right)_{\tau}\right)\right]>2^{\theta} .
$$

Let $\sigma^{*}, \tau^{*} \subseteq[m]$ be the indices of (one of) the largest summands in the above sum. For every $s \in\{0,1\}^{d}$, let

$$
Y_{s}=\left(X_{s}\right)_{\sigma^{*}} \oplus\left(X_{\mathcal{A}(s)}\right)_{\tau^{*}}
$$

There is a set $S^{\prime \prime} \subseteq\{0,1\}^{d}$ satisfying that $\left|S^{\prime \prime}\right|>\frac{\xi \cdot 2^{d-2}}{2^{m t}\left(2^{m}-1\right)(t+1)^{2}}=\frac{\xi \cdot 2^{d-2}}{2^{m+2}\left(2^{m}-1\right)}$. The $S^{\prime \prime}$ here is the same as that in the proof of Theorem 1 by replacing $t$ there with 1 and the error $2^{-m}$ there with $\xi$. Please see [5] for details.

Define a random variable $Y_{S^{\prime \prime}}$ over $\{0,1\}$ as follows: To sample a bit from $Y_{S^{\prime \prime}}$, uniformly sample a string $s$ from $S^{\prime \prime}$, and then independently sample a string $w$ uniformly from $W^{\prime}$. The sampled value is $Y_{s}(w)$. We have that

$$
\operatorname{bias}\left(Y_{S^{\prime \prime}}\right)>\frac{\xi}{2^{m t+1}\left(2^{m}-1\right)(t+1)}=\frac{\xi}{2^{m+2}\left(2^{m}-1\right)}
$$

For every $s \in S^{\prime \prime}$, let

$$
Y_{s}^{\prime}=\oplus_{i \in \sigma^{*}} Z_{(i, s)} \oplus\left(\oplus_{j \in \tau^{*}} Z_{(j, \mathcal{A}(s))}\right)
$$

Let $t=1$ in Claim 7.2 of [5], we get the following claim.
Claim 2'. The set of random variables $\left\{Y_{s}^{\prime}\right\}_{s \in S^{\prime \prime}} \epsilon$-fools linear tests of size $\frac{k^{\prime}}{(t+1) m}=4$.

We apply Theorem 2 on the random variables $\left\{Y_{s}^{\prime}\right\}_{s \in S^{\prime \prime}}$. For simplicity of presentation, we assume $\left|S^{\prime \prime}\right|=2^{d^{\prime}}$. By Theorem 2, the distribution of Ext ${ }^{(1)}\left(W, S^{\prime \prime}\right)$ is $\gamma$-biased for $\gamma=2^{\frac{n-\alpha}{k}} \cdot\left[2^{-\frac{d^{\prime} k}{2}} \cdot(k-1) \cdot(k-3) \cdots \cdots 1 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{k}}$. Let $k=4$, then $\gamma=2^{\frac{n-\alpha}{4}} \cdot\left[2^{-2 d^{\prime}} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}}$. We note that Ext ${ }^{(1)}\left(W, S^{\prime \prime}\right)$ has the same distribution as $Y_{S^{\prime \prime}}$. In particular, both random variables have the same bias. Therefore, we get

$$
2^{\frac{n-\alpha}{4}} \cdot\left[2^{-2 d^{\prime}} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}} \geq \operatorname{bias}\left(Y_{S^{\prime \prime}}\right)>\frac{\xi}{2^{m+2}\left(2^{m}-1\right)}
$$

Moreover, since $2^{d^{\prime}}=\left|S^{\prime \prime}\right|>\frac{\xi \cdot 2^{d-2}}{2^{m+2}\left(2^{m}-1\right)}$, we have

$$
2^{\frac{n-\alpha}{4}} \cdot\left[\xi^{-2} \cdot 2^{-2 d+2 m+8} \cdot\left(2^{m}-1\right)^{2} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]^{\frac{1}{4}}>\frac{\xi}{2^{m+2} \cdot\left(2^{m}-1\right)}
$$

Hence,

$$
2^{n-\alpha} \cdot\left[2^{-2 \theta} \cdot 2^{-2 d+4 m+8} \cdot 3 \cdot(1-\epsilon)+\epsilon\right]>\frac{2^{4 \theta}}{2^{8 m+8}}
$$

That is,

$$
2^{-2 d}>\frac{2^{4 \theta-8 m-8-n+\alpha}-\epsilon}{3(1-\epsilon) 2^{-2 \theta+4 m+8}}
$$

with $\epsilon=2^{-\frac{n}{2}+4+\log d}$, which is in contradiction to the assumption of the theorem.

Definition 6. (see [5, 8]) In an ( $n, k, m, \delta$ )-privacy amplification protocol (or information-theoretic key agreement protocol), Alice and Bob share a weak secret $W$, and have two candidate keys $r_{A}, r_{B} \in\{0,1\}^{m} \cup \perp$, respectively. For any adversarial strategy employed by Eve, denote two random variables $R_{A}, R_{B}$
as the values of the candidate keys $r_{A}, r_{B}$ at the conclusion of the protocol execution, and random variable $T_{E}$ as the transcript of the (entire) protocol execution as seen by Eve. We require that for any weak secret $W$ with minentropy at least $k$ the protocol satisfies the following three properties:

- Correctness: If Eve is passive, then one party reaches the state, the other party reaches the KeyConfirmed state, and $R_{A}=R_{B}$.
- Privacy: Denote KeyDerived ${ }_{A}$ and KeyDerived $_{B}$ as the indicators of the events in which Alice and Bob reach the KeyDerived state, respectively. Then for any adversarial strategy employed by Eve, if Bob reaches the KeyDerived ${ }_{B}$ state during the protocol execution, then $\mathrm{SD}\left(\left(R_{B}, T_{E}\right),\left(U_{m}, T_{E}\right)\right) \leq \delta$; if Alice reaches the KeyDerived ${ }_{A}$ state during the protocol execution, then $\operatorname{SD}\left(\left(R_{A}, T_{E}\right),\left(U_{m}, T_{E}\right)\right)$ $\leq \delta$.
- Authenticity: Denote KeyConfirmed ${ }_{A}$ and KeyConfirmed ${ }_{B}$ as the indicators of the events in which Alice and Bob reach the KeyConfirmed state, respectively. Then, for any adversarial strategy employed by Eve, it holds that

$$
\operatorname{Pr}\left[\left(\text { KeyConfirmed }_{A} \vee \text { KeyConfirmed }_{B}\right) \wedge R_{A} \neq R_{B}\right] \leq \delta
$$


[^0]:    ${ }^{1}$ In other literature (e.g., $\left.[11,21]\right), X$ is $\epsilon$-close to $Y$ if $\left.\frac{1}{2}\|X-Y\|_{1}=\frac{1}{2} \sum_{s} \right\rvert\, \operatorname{Pr}[X=$ $s]-\operatorname{Pr}[Y=s] \mid \leq \epsilon$.

[^1]:    ${ }^{2} \alpha_{0}$ here means $\alpha$ in Theorem 3.21 of [11].

