

When are Fuzzy Extractors Possible?

Benjamin Fuller* Leonid Reyzin† Adam Smith‡

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Abstract

Fuzzy extractors (Dodis et al., Eurocrypt 2004) convert repeated noisy readings of a high-entropy secret into the same uniformly distributed key. A minimum condition for the security of the key is the hardness of guessing a value that is similar to the secret, because the fuzzy extractor converts such a guess to the key.

We define *fuzzy min-entropy* to quantify this property of a noisy source of secrets. Fuzzy min-entropy measures the success of the adversary when provided with *only* the functionality of the fuzzy extractor, that is, the *ideal* security possible from a noisy distribution. High fuzzy min-entropy is necessary for the existence of a fuzzy extractor.

We ask: *is high fuzzy min-entropy a sufficient condition for key extraction from noisy sources?* If only computational security is required, recent progress on program obfuscation gives evidence that fuzzy min-entropy is indeed sufficient. In contrast, information-theoretic fuzzy extractors are not known for many practically relevant sources of high fuzzy min-entropy.

In this paper, we show that fuzzy min-entropy is also sufficient for information-theoretically secure fuzzy extraction. For every source distribution W for which security is possible we give a secure fuzzy extractor.

Our construction relies on the fuzzy extractor knowing the precise distribution of the source W . A more ambitious goal is to design a single extractor that works for all possible sources. We show that this more ambitious goal is impossible: we give a family of sources with high fuzzy min-entropy for which no single fuzzy extractor is secure. This result emphasizes the importance of accurate models of high entropy sources.

Keywords: Fuzzy extractors, secure sketches, information theory, biometric authentication, error-tolerance, key derivation, error-correcting codes.

1 Introduction

Sources of reproducible secret random bits are necessary for many cryptographic applications. In many situations these bits are not explicitly stored for future use, but are obtained by repeating the same process (such as reading a biometric or a physically unclonable function) that generated them the first time. However, bits obtained this way present a problem: noise [Dau04, ZH93, BS00, EHMS00, MG09, MRW02, PRTG02, GCVDD02, TSv⁺06, SD07, BBR88]. That is, when a secret is read multiple times,

*Email: bfuller@cs.bu.edu. Boston University and MIT Lincoln Laboratory.

†Email: reyzin@cs.bu.edu. Boston University.

‡Email: asmith@cse.psu.edu. Pennsylvania State University. This work performed while visiting Boston University's Hariri Institute for Computing.

readings are close (according to some metric) but not identical. To utilize such sources, it is often necessary to remove noise, in order to derive the same value in subsequent readings.

The same problem occurs in the interactive setting, in which the secret channel used for transmitting the bits between two users is noisy and/or leaky [Wyn75]. Bennett, Brassard, and Robert [BBR88] identify two fundamental tasks. The first, called information reconciliation, removes the noise without leaking significant information. The second, known as privacy amplification, converts the high entropy secret to a uniform random value. In this work, we consider the noninteractive version of these problems, in which these tasks are performed together with a single message.

The noninteractive setting is modeled by a primitive called a fuzzy extractor [DORS08], which consists of two algorithms. The generate algorithm (**Gen**) takes an initial reading w and produces an output key along with a nonsecret helper value p . The reproduce (**Rep**) algorithm takes the subsequent reading w' along with the helper value p to reproduce key. The correctness guarantee is that the key is reproduced precisely when the distance between w and w' is at most t .

The security requirement for fuzzy extractors is that key is uniform even to a (computationally unbounded) adversary who has observed p . This requirement is harder to satisfy as the allowed error tolerance t increases, because it becomes easier for the adversary to guess key by guessing a w' within distance t of w and running $\text{Rep}(w', p)$.

Fuzzy Min-Entropy We introduce a new entropy notion that precisely measures how hard it is for the adversary to guess a value within distance t of the original reading w . Suppose w is sampled from a distribution W . To have the maximum chance that w' is within distance t of w , the adversary would want to maximize the total probability mass of W within the ball $B_t(w')$ of radius t around w' . We therefore define *fuzzy min-entropy* $H_{t,\infty}^{\text{fuzz}}(W) \stackrel{\text{def}}{=} -\log \max_{w'} \Pr[W \in B_t(w')]$. Observe that this quantity can be bounded in terms of min-entropy: $H_\infty(W) \geq H_{t,\infty}^{\text{fuzz}}(W) \geq H_\infty(W) - \log |B_t|$.

Fuzzy min-entropy measures the ideal security of a noisy source of entropy when used for key derivation, similar to distributional notions of program obfuscation (e.g., [BGI⁺01, DS05a, DS05b]; see Section 1.2 for discussion). Superlogarithmic fuzzy min-entropy is *necessary* for nontrivial key extraction (Proposition 2.6).

However, existing constructions do not measure their security in terms of fuzzy min-entropy; instead, their security is shown to be $H_\infty(W)$ minus some loss, for error-tolerance, that is at least $\log |B_t|$. Since $H_\infty(W) - \log |B_t| \leq H_{t,\infty}^{\text{fuzz}}(W)$, it is natural to ask whether this loss is necessary. This question is particularly relevant when the gap between the two sides of the inequality is high.¹ As an example, iris scans appear to have significant $H_{t,\infty}^{\text{fuzz}}(W)$ (because iris scans for different people appear to be well-spread in the metric space [Dau06]) but negative $H_\infty(W) - \log |B_t|$ [BH09, Section 5]. We therefore ask: *is fuzzy min-entropy sufficient for fuzzy extraction?* There is evidence that it may be when the security requirement is computational rather than information-theoretic—see Section 1.2. We provide an answer in two settings.

Sufficiency of $H_{t,\infty}^{\text{fuzz}}(W)$ for a Precisely Known Distribution Ideally, a fuzzy extractor has *precise knowledge* of the probability distribution function of W . We call this the *precise knowledge* setting. In this setting, we show that for every source W with superlogarithmic $H_{t,\infty}^{\text{fuzz}}(W)$, it is possible to construct a fuzzy extractor with a superlogarithmic length key (Corollary 3.7). Our construction crucially utilizes the probability distribution function of W and, in particular, is not polynomial time. This result shows

¹For nearly uniform distributions, $H_{t,\infty}^{\text{fuzz}}(W) \approx H_\infty(W) - \log |B_t|$. In this setting, standard coding based constructions of fuzzy extractors (instantiated with optimum codes) yield keys of size approximately $H_{t,\infty}^{\text{fuzz}}(W)$.

that $H_{t,\infty}^{\text{fuzz}}(W)$ is a necessary and sufficient condition for building a fuzzy extractor for a given distribution W .

A number of previous works in the precise knowledge setting have provided efficient algorithms and tight bounds for specific distributions—generally the uniform distribution or i.i.d. sequences (for example, [JW99, LT03, TG04, HAD06], [WRDI11, IW12]). Our characterization unifies previous work, and justifies using $H_{t,\infty}^{\text{fuzz}}(W)$ as the measure of the quality of a noisy distribution, rather than cruder measures such as $H_\infty(W) - \log |B_t|$.

The Challenge of Precise Distributional Knowledge Assuming precise knowledge of a distribution W may be unrealistic. Indeed, high-entropy distributions can never be fully observed directly and must therefore be modeled. It is imprudent to assume that the designer’s model of a distribution is completely accurate—the adversary, with greater resources, would likely be able to build a better model. Therefore, fuzzy extractor designs cannot usually be tailored to one particular source. Existing designs work for a family of sources (for example, all sources of min-entropy at least m with at most t errors). Thus, the design is fixed with a partially known distribution, and the adversary may know more about the distribution than the designer of the fuzzy extractor. We call this the *distributional uncertainty* setting.

The Cost of Distributional Uncertainty When the distribution is W is uncertain, our results are negative. In this setting, W is known to be an element of a family of distributions \mathcal{W}_Z . We construct a family \mathcal{W}_Z where not even a 2-bit fuzzy extractor can be secure for most distributions in \mathcal{W}_Z . We emphasize that each distribution $W_z \in \mathcal{W}_Z$ has superlogarithmic fuzzy min-entropy—in fact, $H_{t,\infty}^{\text{fuzz}}(W_z) = H_\infty(W_z)$, because all points in W_z are distance at least t apart. Our proof relies on high dimensionality of each W_z and on perfect correctness of the Rep procedure. This result shows that uncertainty in the distribution W is devastating to building a fuzzy extractor. In particular, our positive results (Corollary 3.7) show it is possible to build a fuzzy extractor for each W_z , but no construction can simultaneously secure the whole family. Our result motivates further research into high fidelity descriptions of noisy sources.

	Precise Knowledge	Distributional Uncertainty
Fuzzy Extractor	Yes (Corollary 3.7)	No (Theorem 5.1)
Secure Sketch	Yes (Corollary 3.7)	No (Theorem 4.1)

Table 1: Is fuzzy min-entropy sufficient to extract an information-theoretic superlogarithmic length key?

Stronger Results on Information Reconciliation (Secure Sketches) Traditionally, fuzzy extractors use a secure sketch to perform information reconciliation (mapping w' back to w), followed by randomness extractor [NZ93] to transform w into a uniform key. The security losses incurred in the first of these two steps dominate for typical sources and, indeed, this step is less well understood.² Formally, a secure sketch performs non-interactive information reconciliation via pair of algorithms: **SS** takes w and produces a nonsecret value ss , while **Rec** takes a value w' within distance t of w and uses ss to output the original reading w .

For secure sketches, we show results that are similar to but stronger than our results for fuzzy extractors. Namely, we show in Corollary 3.7 that secure sketches are possible if the distribution W is precisely

²Randomness extractors have matching upper and lower bounds on the security loss: for every extra two bits of output key, they lose one bit of security

known. (In fact, we obtain our fuzzy extractors for the case of a precisely known distribution from this result by applying a randomness extractor.)

On the other hand, there is a family of sources, \mathcal{W}_Z , where each element has $H_{t,\infty}^{\text{fuzz}}(W_z) = H_\infty(W_z) = \omega(\log n)$ for which no secure sketch correcting even a few errors is possible (Theorem 4.1). The impossibility result applies even when Rec is allowed to be incorrect with probability up to $1/4$ (in contrast to our fuzzy extractor impossibility result, which requires perfect reconstruction).

1.1 Our Techniques

Techniques for Positive Results for a Precisely Known Distribution We now explain how to construct a secure sketch for a precisely known distribution W with fuzzy min-entropy (we already explained how to construct a fuzzy extractor from it). We begin with distributions in which all points in the support have the same probability (so-called “flat” distributions). Consider some subsequent reading w' . To achieve correctness, the sketch algorithm must disambiguate which point $w \in W$ within distance t of w' was sketched. Disambiguating multiple points can be accomplished by universal hashing, as long as the size of hash output space is slightly greater than the number of possible points. Thus, our sketch is computed via a universal hash of w . To determine the length of that sketch, consider the heaviest (according to W) ball B^* of radius t . Because the distribution is flat, B^* is also the ball with the most points of nonzero probability. Thus, the length of the sketch needs to be slightly greater than the logarithm of the number of non-zero probability points in B^* . Since $H_{t,\infty}^{\text{fuzz}}(W)$ is determined by the weight of B^* , the number of points cannot be too high and there will be remaining entropy after the sketch is published. This remaining entropy suffices to extract a key.

For an arbitrary distribution, we cannot afford to disambiguate points in the ball with the greatest number of points, because there could be too many low-probability points in a single ball despite a high $H_{t,\infty}^{\text{fuzz}}(W)$. We solve this problem by splitting the arbitrary distribution into a number of nearly flat distributions we call “levels.” We then write down, as part of the sketch, the level of the original reading w and apply the above construction considering only points in that level. We call this construction *leveled hashing* (Construction 3.5).

Techniques for Negative Results for Distributional Uncertainty We construct a family of distributions \mathcal{W}_Z and prove impossibility for a uniformly random $W_z \leftarrow \mathcal{W}_Z$. We start by observing the following asymmetry: Gen sees only the sample w (obtained via $W_z \leftarrow \mathcal{W}_Z$ and $w \leftarrow W_z$), while the adversary knows W_z . To exploit the asymmetry, we construct \mathcal{W}_Z so that conditioning on the knowledge of W_z (the outcome z) reduces the distribution to a single affine line, but conditioning on *only* w leaves the rest of the distribution uniform on a large fraction of the entire space.

An adversary can exploit the knowledge of the affine line to reduce the uncertainty about w (in the secure sketch case) or key (in the fuzzy extractor case). In the secure sketch case, ss can be used to find fixed points of $\text{Rec}(\cdot, ss)$ which, by the correctness requirement of the sketch, must be separated by minimum distance t . This means there aren’t too many of them, so few can lie on an average line, permitting the adversary to guess one easily.

In the fuzzy extractor case, the nonsecret value p partitions the metric space into regions that produce a consistent value under Rep (preimages of each key under $\text{Rep}(\cdot, p)$). For each of these regions, the adversary knows that possible w lie t -far from the boundary of the region. However, in the Hamming space, the vast majority of points lie near the boundary (this follows by combining the isoperimetric inequality [Har66] showing that the ball has the smallest boundary and Hoeffding’s inequality [Hoe63] for bounding the volume that is t -away from this boundary). This allows the adversary to rule out so many

possible w that, combined with the adversarial knowledge of the affine line, many regions become empty, leaving key far from uniform.

The result for fuzzy extractors is delicate. It uses the fact that p partitions the space into nonoverlapping regions, which is implied by perfect correctness. Extending this result to imperfect correctness seems challenging and is an interesting open problem. Our result also uses the fact that there are few points far from the boundary of every region, which is implied by the geometry of the high-dimensional Hamming space. This fact seems crucial: in contrast, in low-dimensional Euclidean space, which does not have this property, a single fuzzy extractor can work for any distribution with sufficient $H_{t,\infty}^{\text{fuzz}}$. (Such a construction would use quantization or tiling, similar to, for example, [CK03, LT03, CZC04, LC06, BDH⁺10, VTO⁺10]. Each sample from W would map to the “tile” containing it, from which the output key would be extracted. A randomly chosen quantizer would have the property that few samples lie near the boundary, giving almost-perfect correctness; if perfect correctness is desired, we can give up on security for those rare samples and simply use a special value of p to indicate that one of them was the input.)

1.2 Related Settings

Other settings with close readings: $H_{t,\infty}^{\text{fuzz}}$ is sufficient The security definition of fuzzy extractors and secure sketches can be weakened to protect only against computationally bounded adversaries [FMR13]. In this computational setting, a single fuzzy extractor (or secure sketch) can simultaneously secure all possible distributions by using virtual grey-box obfuscation for all circuits [BCKP14]. The construction places into p the obfuscated program for testing proximity to w and outputting the appropriate value if the test passes.³ This construction is secure when the adversary can rarely learn key with oracle access to the program functionality. This is true for the set of distributions with fuzzy min-entropy (and only those distributions). Thus, extending our negative result to the computational setting would rule out the existence of virtual grey-box obfuscation for all circuits.

Furthermore, the functional definition of fuzzy extractors and secure sketches can be weakened to permit interaction between the party having w and the party having w' . Such a weakening is useful for secure remote authentication [BDK⁺05]. When both interaction and computational assumptions are allowed, secure two-party computation can produce a key that will be secure whenever the distribution W has fuzzy min-entropy. The two-party computation protocol needs to be secure without assuming authenticated channels; it can be assuming the existence of collision-resistant hash functions and enhanced trapdoor permutations [BCL⁺11].

Correlated rather than close readings A different model for the problem of key derivation from noisy sources does not explicitly consider the distance between w and w' , but rather views w and w' as samples of drawn from a correlated pair of random variables. This model is considered in multiple works, including [Wyn75, CK78, AC93, Mau93]; recent characterizations of when key derivation is possible in this model include [RW05] and [TW14]. We compare our positive results to these characterizations in Appendix A. To the best of our knowledge, prior results on correlated random variables are in the precise knowledge setting, we are unaware of works that consider the cost of distributional uncertainty.

Organization The remainder of the paper is organized as follows. In Section 2, we cover preliminaries and fuzzy extractor definitions. In Section 3, we construct a fuzzy extractor in the precise knowledge

³If this construction is used for a secure sketch, W will remain unpredictable conditioned on p , but will not have pseudoentropy (see Section 4.1 for details).

setting. In Sections 4 and 5 we construct families of distributions that no secure sketches and fuzzy extractors can secure, respectively (the distributional uncertainty setting).

2 Preliminaries

Usually, we use capitalized letters for random variables and corresponding lowercase letters for their samples. Unless otherwise noted logarithms are base 2. The *min-entropy* of W is $H_\infty(W) = -\log(\max_w \Pr[W = w])$, and the *average (conditional) min-entropy* of W given P is $\tilde{H}_\infty(W|P) = -\log(\mathbb{E}_{p \in P} \max_w \Pr[W = w|P = p])$ [DORS08, Section 2.4]. Let $H_0(W)$ be the logarithm of the size of the support of W , that is $H_0(W) = \log|\{w | \Pr[W = w] > 0\}|$. We use an average case of remaining support size $\tilde{H}_0(W|P) = \log(\mathbb{E}_{p \in P} |\{w | \Pr[W = w|P = p] > 0\}|)$.

The *statistical distance* between random variables X and Y with the same domain is $\mathbf{SD}(X, Y) = \frac{1}{2} \sum_x |\Pr[X = x] - \Pr[Y = x]|$. For a metric space $(\mathcal{M}, \text{dis})$, the *(closed) ball of radius t around w* is the set of all points within radius t , that is, $B_t(w) = \{w' | \text{dis}(w, w') \leq t\}$. If the size of a ball in a metric space does not depend on w , we denote by $|B_t|$ the size of a ball of radius t . We consider the Hamming metric over vectors in \mathcal{Z}^γ , defined via $\text{dis}(w, w') = |\{i | w_i \neq w'_i\}|$ where \mathcal{Z} is some alphabet. For this metric, $|B_t| = \sum_{i=0}^t \binom{\gamma}{i} (|\mathcal{Z}| - 1)^i$. U_κ denotes the uniformly distributed random variable on $\{0, 1\}^\kappa$. Throughout this work, we consider a sequence of metric spaces \mathcal{M}_n parameterized by n , but we write \mathcal{M} for notational convenience. A *negligible* function $\text{ngl}(n)$ is one that decreases faster than any positive inverse polynomial as $n \rightarrow \infty$.

2.1 Fuzzy Extractors and Secure Sketches

In this section, we define fuzzy extractors and secure sketches. Definitions and lemmas are drawn from the work of Dodis et al. [DORS08, Sections 2.5–4.1] with modifications. First, we allow for error as discussed in [DORS08, Section 8]. Second, in the *distributional uncertainty* setting we consider a general family \mathcal{W}_Z of distributions instead of families containing all distributions of a given min-entropy. Let \mathcal{M} be a metric space with distance function dis .

Definition 2.1. An $(\mathcal{M}, \mathcal{W}_Z, \kappa, t, \epsilon)$ -fuzzy extractor with error δ is a pair of randomized procedures, “generate” (Gen) and “reproduce” (Rep). Gen on input $w \in \mathcal{M}$ outputs an extracted string key $\in \{0, 1\}^\kappa$ and a helper string $p \in \{0, 1\}^*$. Rep takes $w' \in \mathcal{M}$ and $p \in \{0, 1\}^*$ as inputs. (Gen, Rep) have the following properties:

1. Correctness: if $\text{dis}(w, w') \leq t$ and $(\text{key}, p) \leftarrow \text{Gen}(w)$, then

$$\Pr[\text{Rep}(w', p) = \text{key}] \geq 1 - \delta.$$

2. Security: for any distribution $W_z \in \mathcal{W}_Z$, if $(\text{Key}, P) \leftarrow \text{Gen}(W_z)$, then $\mathbf{SD}((\text{Key}, P), (U_\kappa, P)) \leq \epsilon$.

Fuzzy extractors perform two tasks, information-reconciliation and privacy amplification. The standard construction is *sketch-and-extract*: the uniform key is extracted from w (using a randomness extractor [NZ93]) and the error-tolerance is obtained by using a secure sketch [DORS08, Lemma 4.1]. Secure sketches produce a string ss that minimally decreases the entropy of w , while mapping nearby w' to w :

Definition 2.2. An $(\mathcal{M}, \mathcal{W}_Z, \tilde{m}, t)$ -secure sketch with error δ is a pair of randomized procedures, “sketch” (SS) and “recover” (Rec). SS on input $w \in \mathcal{M}$ returns a bit string $ss \in \{0, 1\}^*$. Rec takes an element $w' \in \mathcal{M}$ and $ss \in \{0, 1\}^*$. (SS, Rec) have the following properties:

1. Correctness: $\forall w, w' \in \mathcal{M}$ if $\text{dis}(w, w') \leq t$ then

$$\Pr[\text{Rec}(w', \text{SS}(w)) = w] \geq 1 - \delta.$$

2. Security: for any distribution $W_z \in \mathcal{W}_Z$, $\tilde{H}_\infty(W_z | \text{SS}(W_z)) \geq \tilde{m}$.

In the above definitions, the errors are chosen before ss (resp., p) is known in order for the correctness guarantee to hold: correctness holds for any w' with probability $1 - \delta$ over the coins of the algorithms, but w' cannot be a function of the output of $\text{SS}(w)$.

The Case of a Precisely Known Distribution If in the above definitions we take \mathcal{W}_Z to be a one-element set containing a single distribution W , then the fuzzy extractor/secure sketch is said to be constructed for a *precisely known distribution*. In this case, we need to require correctness only for w that have nonzero probability⁴.

Note that we have no requirement that the algorithms are compact or efficient, and so the distribution can be fully known to them. Finding a natural model of specifying distributions that allows for efficient (yet generic) constructions of sketches and extractors for a precisely known distribution is an interesting problem.

From Secure Sketches to Fuzzy Extractors A fuzzy extractor can be produced from a *secure sketch* and an *average case randomness extractor*:

Definition 2.3. Let \mathcal{M}, χ be finite sets. A function $\text{ext} : \mathcal{M} \times \{0, 1\}^d \rightarrow \{0, 1\}^\kappa$ a (\tilde{m}, ϵ) -average case extractor if for all pairs of random variables X, Y over \mathcal{M}, χ such that $\tilde{H}_\infty(X|Y) \geq \tilde{m}$, we have

$$\text{SD}((\text{ext}(X, U_d), U_d, Y), U_\kappa \times U_d \times Y) \leq \epsilon.$$

Lemma 2.4. Assume (SS, Rec) is an $(\mathcal{M}, \mathcal{W}_Z, \tilde{m}, t)$ -secure sketch with error δ , and let $\text{ext} : \mathcal{M} \times \{0, 1\}^d \rightarrow \{0, 1\}^\kappa$ be a (\tilde{m}, ϵ) -average case extractor. Then the following (Gen, Rep) is an $(\mathcal{M}, \mathcal{W}_Z, \kappa, t, \epsilon)$ -fuzzy extractor with error δ :

- $\text{Gen}(w) : \text{sample } x \leftarrow \{0, 1\}^d$, set $p = (\text{SS}(w), x)$, $r = \text{ext}(w; x)$, output (r, p) .
- $\text{Rep}(w', (s, x)) : \text{recover } w = \text{Rec}(w', s)$ and output $r = \text{ext}(w; x)$.

2.2 Fuzzy Min-Entropy: suitability of a noisy distribution for key derivation

The value p allows everyone, including the adversary, to find the output of $\text{Rep}(\cdot, p)$ on any input w' . Ideally, p should not provide any useful information beyond this ability, and the outputs of Rep on inputs that are too distant from w should provide no useful information, either. In this ideal scenario, the adversary is limited to trying to guess a w' that is t -close to w . Letting w' be the center of the maximum-weight ball in W would be optimal for the adversary. We therefore measure the quality of a source by (the negative logarithm of) this weight.

⁴We can extend correctness to all of \mathcal{M} by defining Gen/SS to output the point w as part of p/ss on zero-probability inputs, which will ensure that Rep/Rec can always be correct; this does not affect security.

Definition 2.5. *The t -fuzzy min-entropy of a distribution W in a metric space $(\mathcal{M}, \text{dis})$ is:*

$$H_{t,\infty}^{\text{fuzz}}(W) = -\log \left(\max_{w'} \sum_{w \in W | \text{dis}(w,w') \leq t} \Pr[W = w] \right)$$

Fuzzy min-entropy is a necessary condition for security. It measures the functionality provided to the adversary by providing **Rep** (since p is public). Thus, the fuzzy min-entropy is the ideal security that any fuzzy extractor should be compared against. We defer proofs of statements in this section to Appendix C.

Proposition 2.6. *Let W be a distribution over $(\mathcal{M}, \text{dis})$ with $H_{t,\infty}^{\text{fuzz}}(W) = m$. Let (Gen, Rep) be a $(\mathcal{M}, \{W\}, \kappa, t, \epsilon)$ -fuzzy extractor with error δ . Then*

$$\epsilon \geq 2^{-m} - \delta - 2^{-\kappa}.$$

In particular, for security parameter n , if $m = \Theta(\log n)$, $\delta = \text{ngl}(n)$, $\kappa = \omega(\log n)$, then $\epsilon = 1/\text{poly}(n)$.

The proof of Proposition 2.6 is not specific to the fuzzy extractor setting, it also applies to computational and interactive definitions. Fuzzy min-entropy represents an upper bound on the security from a noisy source. There is evidence it may be possible to (nearly) achieve this upper bound in the computational setting (see Section 1.2). However, in the information-theoretic setting there are many distributions with fuzzy min-entropy with no known fuzzy extractor (or corresponding impossibility result). We now show some properties of fuzzy min-entropy.

Lemma 2.7. $H_{t,\infty}^{\text{fuzz}}(W|P = p) \geq H_{t,\infty}^{\text{fuzz}}(W) + \log \Pr[P = p]$.

Conditional Fuzzy min-entropy In our proofs, we will use a conditional version of fuzzy min-entropy.

Definition 2.8. *The t -conditional fuzzy min-entropy of a distribution $W|P$ in a metric space $(\mathcal{M}, \text{dis})$ is:*

$$\tilde{H}_{t,\infty}^{\text{fuzz}}(W|P) = -\log \left(\mathbb{E}_{p \in P} \max_{w'} \sum_{w \in W | P=p | \text{dis}(w,w') \leq t} \Pr[W = w | P = p] \right).$$

Lemma 2.9. $\tilde{H}_{t,\infty}^{\text{fuzz}}(W|P) \geq H_{t,\infty}^{\text{fuzz}}(W) - H_0(P)$.

3 Sufficiency of $H_{t,\infty}^{\text{fuzz}}(W)$ in the Precise Knowledge Setting

In this section, we build a secure sketch (and thus fuzzy extractors through Lemma 2.4) for each distribution W with $H_{t,\infty}^{\text{fuzz}}(W) = \omega(\log n)$ (using precise knowledge of W). We begin with flat distributions and then turn to arbitrary distributions.

3.1 Flat Distributions

A distribution is flat if all points in its support have the same probability. Let $\text{supp}(W)$ denote the support of W , i.e., the set of points with nonzero probability.

Definition 3.1. *A distribution W is flat if for all $w_0, w_1 \in \text{supp}(W)$, $\Pr[W = w_0] = \Pr[W = w_1]$.*

Denote the largest number of points in a ball of radius t in the support of W as

$$\beta_t = \max_{w' \in \mathcal{M}} |\{w | w \in \text{supp}(W) \wedge \text{dis}(w, w') \leq t\}|.$$

For flat distributions, this quantity defines the fuzzy min-entropy,

$$\begin{aligned} H_{t, \infty}^{\text{fuzz}}(W) &= -\log \left(\max_{w' \in \mathcal{M}} |\{w | w \in \text{supp}(W) \wedge \text{dis}(w, w') \leq t\}| \Pr[W = w] \right) \\ &= -\log \left(\max_{w' \in \mathcal{M}} |\{w | w \in \text{supp}(W) \wedge \text{dis}(w, w') \leq t\}| \cdot 2^{-H_\infty(W)} \right) \\ &= H_\infty(W) - \log \beta_t. \end{aligned} \tag{1}$$

We use universal hashes to construct secure sketches for flat distributions. Skoric et al. constructed secure sketches from universal hashes to correct a polynomial number of error patterns [STGP09].

Definition 3.2 ([CW79]). *Let $F : \mathcal{K} \times \mathcal{M} \rightarrow R$ be a function. We say that F is universal if for all distinct $x_1, x_2 \in \mathcal{M}$:*

$$\Pr_{K \leftarrow \mathcal{K}} [F(K, x_1) = F(K, x_2)] = \frac{1}{|R|}.$$

Construction 3.3. *Let $F : \mathcal{K} \times \mathcal{M} \rightarrow R$ be a universal hash function. Let W be a distribution. Define $\text{SS}_W, \text{Rec}_W$ as:*

SS_W	Rec_W
1. <u>Input</u> : w .	1. <u>Input</u> : $(w', ss = (y, K))$
2. <u>Sample</u> $K \leftarrow \mathcal{K}$.	2. Let $W^* = \{w \in \text{supp}(W) \text{dis}(w, w') \leq t\}$.
3. Set $ss = F(K, w), K$.	3. For $w^* \in W^*$, if $F(K, w^*) = y$, output w^* .
	4. Output \perp .

Lemma 3.4. *Let W be a flat with $H_\infty(W) \geq m$. Then Construction 3.3 is a $(\mathcal{M}, \{W\}, m - \log |R|, t)$ -known distribution secure sketch with error $\delta \leq \frac{\beta_t - 1}{|R|}$.*

Proof. We first argue security. Fix some $W \in \mathcal{W}$. Since \mathcal{K} and W are independent $\tilde{H}_\infty(W|\mathcal{K}) = H_\infty(W) = m$. Then by [DORS08, Lemma 2.2b], $\tilde{H}_\infty(W|\mathcal{K}, F(\mathcal{K}, W)) \geq H_\infty(W) - \log |F(\mathcal{W}, W)| \geq m - \log |R|$. We now argue correctness. Fix some w, w' . Let W^* denote the set of elements in W within distance t of w' . The size of W^* is at most β_t . Since w, w' are independent of SS this set is independent of the choice of \mathcal{K} . The algorithm Rec will never output \perp as the correct w will match the hash. The probability that another element w^* collides is:

$$\begin{aligned} \Pr[\exists w^* \in W^* | w^* \neq w \wedge F(K, w^*) = F(K, w)] \\ &\leq \sum_{w^* \in W^* | w^* \neq w} \Pr[F(K, w^*) = F(K, w)] \\ &= \sum_{w^* \in W^* | w^* \neq w} \frac{1}{|R|} \leq \frac{\beta_t - 1}{|R|} \end{aligned}$$

The inequality proceeds by union bound. The first equality proceeds by the universality of F and the second inequality proceeds by noting the number of wrong neighbors is bounded by $\beta_t - 1$. This completes the proof. \square

3.2 Arbitrary Distributions

The hashing approach used in the previous subsection does not work for arbitrary sources. The reason is that some balls may have many points but low total weight. For example, let W be a distribution consisting of the following balls. Denote by B_t^1 a ball with $2^{\mathsf{H}_\infty(W)}$ points with probability $\Pr[W \in B_t^1] = 2^{-\mathsf{H}_\infty(W)}$. Let $B_t^2, \dots, B_t^{2^{-\mathsf{H}_\infty(W)}}$ be balls with one point each with probability $\Pr[W \in B_t^i] = 2^{-\mathsf{H}_\infty(W)}$. Then the hashing algorithm needs to write down $\mathsf{H}_\infty(W)$ bits to achieve correctness on B_t^1 . However, with probability $1 - 2^{-\mathsf{H}_\infty(W)}$ the initial reading is outside of B_t^1 , and the hash completely reveals the point.

Dealing with non-flat distributions requires a new strategy. Many solutions for manipulating high entropy distributions leverage a solution for flat distributions and use the fact that high entropy distributions are convex combinations of flat distributions. However, a distribution with high fuzzy min-entropy may be formed from component distributions with little or no fuzzy min-entropy. It is unclear how to leverage the convex combination property in this setting.

The main obstacle in the arbitrary setting is distinguishing between a setting where a ball has a few high probability points and a large number of low probability points. To overcome this problem, we write the probability of $w \in W$ in the sketch output. To ensure this information does not completely reveal w we write down only the approximate probability of the outcome w . We then use a universal hash whose output length is proportional to the maximum number of points of the same approximate probability in any ball.

Construction 3.5. *Let \mathcal{M} be a metric space and let $n = \log |\mathcal{M}|$. Let W be a distribution with $\mathsf{H}_\infty(W) = m$. Let $\ell \in \mathbb{Z}^+$ be a parameter. Let $L_i = (2^{-(i+1)}, 2^{-i}]$ for $i = m, \dots, m + \ell - 1$. Let $F_i : \mathcal{K}_i \times \mathcal{M} \rightarrow R_i$ be a parameterized family of universal hash functions. Define $\mathsf{SS}_W, \mathsf{Rec}_W$ as:*

SS_W	Rec_W
<ol style="list-style-type: none"> 1. <u>Input</u>: w. 2. If $\Pr[W = w] \leq 2^{-(m+\ell)}$. Set $ss = 1, w$. 3. Else <ol style="list-style-type: none"> (a) Find i such that $\Pr[W = w] \in L_i$. (b) Sample $K \leftarrow \mathcal{K}_i$. (c) Set $ss = 0, i, F_i(K, w), K$. 	<ol style="list-style-type: none"> 1. <u>Input</u>: (w', ss) 2. If $ss_0 = 1$, output $ss_{1, \dots, y }$. 3. Else <ol style="list-style-type: none"> (a) Parse $(i, y, K) = ss_{1, \dots, y }$. (b) $W^* = \{w \mid \text{dis}(w, w') \leq t \wedge \Pr[W = w] \in L_i\}$. (c) For $w^* \in W^*$, if $F_i(K, w^*) = z$, output w^*. (d) Output \perp.

To set parameters, we restrict our notation of the maximum likelihood ball to points of a given probability. Define $\beta_{t,i}$ as the maximum number of points in a ball in level i . That is,

$$\beta_{t,i} = \max_{w' \in \mathcal{M}} |\{w \mid \text{dis}(w, w') \leq t \wedge \Pr[W = w] \in L_i\}|.$$

Theorem 3.6. *Let W be a distribution over \mathcal{M} where $n = \log |\mathcal{M}|$. Let $\delta > 0$ be an function of n . Let $F_i : \mathcal{K}_i \times \mathcal{M} \rightarrow R_i$ be a parameterized family of universal hash functions where $|R_i| = \beta_{t,i}/\delta$. When $\ell = n$ Construction 3.5 is a $(\mathcal{M}, \{W\}, \tilde{m}, t)$ -known distribution secure sketch with error δ for $\tilde{m} = \mathsf{H}_{t,\infty}^{\text{fuzz}}(W) - \log n - \log 1/\delta - 4$.*

We provide a proof in Appendix D. The main idea is that providing the level information makes the distribution look nearly flat (the probability of points differs by at most a factor of two). We can apply techniques from Lemma 3.4 to show security for each nearly flat distribution. Then, we show that providing the level information does not hurt security too much.

Corollary 3.7. *Let \mathcal{M} be a sequence of metric space parameterized by n where $n = \log |\mathcal{M}|$. For any distribution W over \mathcal{M} with $H_{t,\infty}^{\text{fuzz}}(W) = \omega(\log n)$, there exists a $(\mathcal{M}, \{W\}, \tilde{m}, t)$ -known distribution secure sketch with $\tilde{m} = \omega(\log n)$ and $\delta = \text{ngl}(n)$. (Extendible to a fuzzy extractor using Lemma 2.4.)*

4 Impossibility of Secure Sketches for a Family with $H_{t,\infty}^{\text{fuzz}}$

In the previous section, we showed the sufficiency of $H_{t,\infty}^{\text{fuzz}}(W)$ when the distribution was precisely known. It may be infeasible to completely characterize a high entropy distribution W . Traditionally, algorithms deal with this *distributional uncertainty* by providing security for a family of distributions \mathcal{W}_Z .

In this section, we show distributional uncertainty of W comes at a real cost. The security game of a fuzzy extractor can be thought of as a three stage process: 1) the challenger specifies (SS, Rec) , 2) the adversary sees (SS, Rec) and specifies $W_z \in \mathcal{W}_Z$ 3) the adversary wins if $\tilde{H}_\infty(W_z | \text{SS}(W_z)) < \tilde{m}$.

We prove impossibility in a game that is harder for the adversary to win: 1) the challenger specifies (SS, Rec) 2) the adversary randomly samples $W_z \leftarrow \mathcal{W}_Z$ 3) the adversary wins if $\tilde{H}_\infty(W_z | \text{SS}(W_z)) < \tilde{m}$.⁵

Let V be the process of uniformly sampling $W_z \leftarrow \mathcal{W}_Z$ and then sampling $w \leftarrow W_z$. Let the random variable Z represent the process of sampling $W_z \leftarrow \mathcal{W}_Z$. The view of the challenger is V , while the view of the adversary is $\text{SS}(V)$ and Z . We now show a family of distributions \mathcal{W}_Z that does not admit a secure sketch.

Theorem 4.1. *Let n be a security parameter. There exists a family of distributions \mathcal{W}_Z such that for each element $W_z \in \mathcal{W}_Z$, $H_{t,\infty}^{\text{fuzz}}(W_z) = \omega(\log n)$, and yet for any $(\mathcal{M}, \mathcal{W}_Z, \tilde{m}, t)$ -secure sketch (SS, Rec) with error $\delta < 1/4$ and distance $t \geq 4$, $\tilde{m} < 2$.*

Furthermore, this is true on average. Let V be process of sampling $W_z \leftarrow \mathcal{W}_Z$ and sampling $w \leftarrow W_z$. Then

$$\tilde{H}_\infty(V | \text{SS}(V), Z) < 2.$$

Proof. We prove the stronger average case statement. We first describe a family \mathcal{W}_Z . Let \mathbb{F} be a field where of size $|\mathbb{F}| = \omega(\text{poly}(n))$. Let \mathcal{W}_Z be the set of all distributions of the form

$$w = \begin{pmatrix} \vec{1} \\ a_2 \\ \vdots \\ a_\gamma \end{pmatrix} w_1 + \begin{pmatrix} 0 \\ b_2 \\ \vdots \\ b_\gamma \end{pmatrix}$$

The family is defined by the parameters $z = a_2, \dots, a_\gamma, b_2, \dots, b_\gamma$ (seen by the adversary) where $a_i \neq 0$. The outcome $w \leftarrow W_z$ is sampled by sampling w_1 and computing w . Each distribution W_z is an affine line in space \mathbb{F}^γ . The algorithm SS, Rec never see Z only V . Fix some $4 \leq t < \gamma$. We show the following (in Appendix E):

⁵Our results rule out security for an average member of \mathcal{W}_Z . It may be possible to improve parameters by ruling out only a worst case W_z . In Appendix B, we show that providing security for the set of distributions \mathcal{W}_Z is equivalent to providing security for all distributions Z over that family.

- Proposition E.2: for all $W_z \in \mathcal{W}_Z$, $H_{t,\infty}^{\text{fuzz}}(W_z) = \omega(\log n)$.
- Proposition E.3: the distribution V is uniform.
- Lemma E.4: for any secure sketch on V , the support size of $V|\text{SS}(V)$ decreases significantly. Here we show the minimum distance between points of $V|\text{SS}(V)$ is at least t .
- Lemma E.5: $\tilde{H}_0(V|\text{SS}(V), Z) < 2$ and thus $\tilde{H}_\infty(V|\text{SS}(V)) < 2$.

□

Note: There is a tradeoff between the size of \mathbb{F} and the error tolerance required for the counter example. By increasing t it is possible to show a counter example for a smaller \mathbb{F} .

4.1 Implications for Computational Secure Sketches

Fuller et al. show that computational secure sketches that provide pseudoentropy imply information-theoretic secure sketches with almost the same parameters [FMR13, Corollary 3.8]. The definition of Fuller et al. uses a weak version of pseudoentropy [HILL99] due to Gentry and Wichs [GW11].

Definition 4.2. $W|S$ has relaxed HILL entropy, denoted $H_{\epsilon_{\text{sec}}, s_{\text{sec}}}^{\text{HILL-r1x}}(W|S) \geq \tilde{m}$, if there exists a joint distribution (X, Y) , such that $\tilde{H}_\infty(X|Y) \geq \tilde{m}$ and

$$\delta^{\mathcal{D}_{s_{\text{sec}}}}((W, S), (X, Y)) \leq \epsilon_{\text{sec}}.$$

By the contrapositive of [FMR13, Corollary 3.8], no sketch can retain HILL entropy for the same family of distributions:

Corollary 4.3. Let n be a security parameter and let $\mathcal{M} = |\mathbb{F}|^\gamma$. There exists a family of distributions \mathcal{W}_Z over \mathcal{M} such that for each element $W_z \in \mathcal{W}_Z$, $(\mathcal{M}, \mathcal{W}_Z, \tilde{m}, t)$ -secure sketch (SS, Rec) with error δ , then

$$H_{\epsilon_{\text{sec}}, s_{\text{sec}}}^{\text{HILL-r1x}}(W|\text{SS}(W), Z) < 4.$$

if $t \geq 4$, $s_{\text{sec}} \geq t(|\text{Rec}| + \gamma \log |\mathbb{F}|)$, and $\epsilon_{\text{sec}} + t\delta < 1/16$.

Secure sketches that provide computational unpredictability are implied by virtual-grey box obfuscation of proximity functions [BCKP14]. Our impossibility result says nothing about this weaker form of a secure sketch. Extraction from unpredictability entropy can be done using an extractor with a reconstruction property [BSW03, HLR07]; however, in the computational setting, the obfuscated function can simply hide a randomly generated key, and therefore extraction is not necessary to obtain a fuzzy extractor.

5 Impossibility of Fuzzy Extractors for a Family with $H_{t,\infty}^{\text{fuzz}}$

In the previous section, we showed a family of distributions that does not admit a secure sketch. We provide an analogous result for fuzzy extractors.

Theorem 5.1. Let n be a security parameter. There exists a family of distributions \mathcal{W}_Z over $\{0, 1\}^n$ satisfying the following conditions. For each element $W_z \in \mathcal{W}_Z$, $H_{t,\infty}^{\text{fuzz}}(W_z) = \omega(\log n)$. Let $\kappa \geq 2$ and $t = \omega(n^{1/2} \log n)$. Any $(\mathcal{M}, \mathcal{W}_Z, \kappa, t, \epsilon)$ -fuzzy extractor with error $\delta = 0$ has $\epsilon > 1/8 - \text{ngl}(n)$.

Furthermore, this is true on average. Let V be process of uniformly sampling $W_z \leftarrow \mathcal{W}_Z$ and sampling $w \leftarrow W_z$. Let $(\text{Key}, P) \leftarrow \text{Gen}(V)$. Then,

$$\text{SD}((\text{Key}, P, Z), (U_\kappa, P, Z)) > 1/8 - \text{ngl}(n).$$

Proof Outline. We prove the stronger average case statement. Let $\nu = \omega(\log n)$ and $\nu = o(n^{1/2}/\log n)$. Let $t = 4\nu n^{1/2} + 1$ and note that $n/\nu > t$.

Our counterexample uses a slightly different family of distributions \mathcal{W}_Z than the counterexample for secure sketches. We will work over a binary alphabet (we used a large alphabet in our counterexample for secure sketches). A property of the binary Hamming space is that a large fraction of any set of bounded size is the near ‘‘boundary’’ of that set. This will be crucial in our proof. We will embed the larger alphabet in the secure sketch counterexample into the binary Hamming metric. Let $x_1, \dots, x_\nu \in \{0, 1\}^\nu$. Let \mathbb{F} denote the field of size 2^ν . Let $a_2, \dots, a_{n/\nu} \in \mathbb{F}$ such that $a_i \neq 0$ and let $b_2, \dots, b_{n/\nu} \in \mathbb{F}$. Interpret x_1, \dots, x_ν as a element $x \in \mathbb{F}$ and let

$$w = \begin{pmatrix} \bar{1} \\ a_2 \\ \vdots \\ a_{n/\nu} \end{pmatrix} x + \begin{pmatrix} 0 \\ b_2 \\ \vdots \\ b_{n/\nu} \end{pmatrix}.$$

The multiplication is in \mathbb{F} . The family \mathcal{W}_Z is indexed by $z = a_2, \dots, a_{n/\nu}, b_2, \dots, b_{n/\nu}$ where $a_i \neq 0$. Define V as the process of uniformly choosing $W_z \leftarrow \mathcal{W}_Z$ and then sampling from $w \leftarrow W_z$. The adversary sees $\text{SS}(V)$ and Z . We then show the following (proofs in Appendix F):

- Proposition F.1: for all $W_z \in \mathcal{W}_Z$, $\text{H}_{t, \infty}^{\text{fuzz}}(W_z) = \omega(\log n)$.
- Proposition F.2: the distribution V is uniform.
- Lemma F.3: In expectation across Z , a large subset of keys are not possible. In more detail,
 - Half the keys have at most $2^{n-\kappa}$ pre images in the metric space (this is at most half the metric space). Denote this set as $\text{Key}_{\text{small}}$.
 - Consider some key $\in \text{Key}_{\text{small}}$. Consider the set of $V_{\text{key}} = \{w \mid \text{Rep}(w, p) = \text{key}\}$. All points in $V \mid \text{SS}(V)$ are distance t from a boundary of V_{key} (the functionality of Rep guarantees that for the true w all nearby points map to the same key). We show that most of V_{key} is near a boundary. A result of Frankel and Füredi says that the boundary of a region is minimized by a ball containing the same number of points [FF81]. Hoeffding’s inequality says that most of a ball lies near its boundary [Hoe63]. The number of w that could produce key is small.
 - There are many possible values for z_1, z_2 for the side information Z (and these possible values are equally likely). Furthermore, the distributions $V \mid Z = z_1$ and $V \mid Z = z_2$ have disjoint support outside of a single point w .
 - For most values of possible Z , the intersection between the viable pre images of $V \mid Z$ and V_{key} contains at most one point (the received point v). Checking if $V \mid Z \cap V_{\text{key}}$ is nonempty is an effective distinguisher.

□

Note: As stated in Section 1.2, using strong computational assumptions it is possible to avoid this result. For the specific family \mathcal{W}_Z , Canetti et al. [CFP⁺14, Construction 5.3] construct computational fuzzy extractors for this family of distributions when \mathbb{F} is large enough under weaker assumptions. (Their construction is stated with imperfect correctness. A construction with perfect correctness is obtained by using a code that corrects t bidirectional errors instead of a code that corrects t unidirectional errors.)

Comparison with Theorem 4.1 The parameters in this result are weaker than those in Theorem 4.1. This result requires: 1) higher error tolerance $t = \omega(n^{1/2} \log n)$ 2) the fuzzy extractor must have perfect correctness. The secure sketch counter example needs $t = 4$ and allows the Rec to be wrong almost 1/4 of the time.

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A Key Derivation from Correlated Random Variables

In this work, we consider two draws from some physical source whose distance was bounded according to some metric space. Instead of considering w, w' that have bounded distance, we can treat W, W' as a pair of correlated random variables. Renner and Wolf [RW05] study this setting, firsting consider information-reconciliation and privacy amplification separately. They show that $H_\infty(W)$ is a necessary and sufficient condition for privacy amplification.⁶ Second, they show that the length of p must grow with the worst case number of possible outcomes for W conditioned on W' .⁷ That is, the length of the public value

$$|p| \geq \max_{w' \in W} \log |\{w | \Pr[W = w | W' = w'] > 0\}|.$$

⁶The results of Renner and Wolf use smooth notions of entropy. A random variable has smooth entropy if it is statistically close to a distribution with true entropy. We describe their results in the terminology of non smooth entropy.

⁷This result also uses a smooth notion of entropy. We describe the non smooth version for simplicity.

Furthermore, they show there exists a protocol with this length p using optimal encoding functions. Intuitively, the public information must describe which possible outcome for W actually occurred. This result describes the maximal length of p and does not argue how p effects security. It may be possible to construct a p that reduces the entropy of w by less than $\log |p|$. In Section 3, we construct schemes with variable length p , providing information-reconciliation for distributions where the bounds of Renner and Wolf provide no security guarantees.

Lastly, the work of Renner and Wolf shows characterize when key derivation is possible from correlated random variables [RW05, Theorem 3]. Fuzzy min-entropy can be generalized to this setting. Fuzzy extractors consider the worst case w' . When considering correlated readings, it is natural to treat W' as a random variable:⁸

Definition A.1. *Let (W, W') be a pair of correlated random variables. The correlated fuzzy min-entropy of W, W' is:*

$$H^{\text{corr}}(W, W') = -\log \left(\max_{w' \in \text{supp}(W')} \sum_{w \in W | \Pr[W=w|W'=w'] > 0} \Pr[W = w] \right).$$

In Definition 2.5, the sum is implicitly over $W = w|W' = w'$ since we assume any w' within distance t is possible. In Section 3 we showed sufficiency of $H_{t, \infty}^{\text{fuzz}}(W)$ for key derivation from noisy sources (Definition 2.5).

Connection to the characterization of [RW05] Renner and Wolf characterize when it is possible to derive keys from correlated random variables [RW05, Theorem 3]. They consider all possible (randomized) transforms T, T' of W, W' into a new pair of variables V, V' . They show that

$$|\text{key}| \leq \sup_{(V, V') \leftarrow (T(W), T'(W))} \left(H_{\infty}(V|T') - \log \max_{v' \in V'} |\{v | \Pr[V = v|T' \wedge V' = v'] > 0\}| \right).$$

Furthermore, they show that there is a transformation that achieves a key of nearly this length. The result is nonconstructive as there is no guidance on how to find the transforms T, T' . Since there is no known bound on the length of T, T' it is not clear how to search the transform space even with unlimited time. Construction 3.5 can be used to derive keys from correlated random variables. The main change is to define

$$W^* = \{w | \Pr[W = w|W' = w'] > 0 \wedge \Pr[W = w] \in L_i\}.$$

Our result shows if one is satisfied with obtaining a strong key when possible (our protocol has losses of $2 \log 1/\epsilon + \log n + \log 1/\delta$), then a protocol is possible (and explicitly constructible) in the original space.

B A Definitional Equivalence

As described in Section 4, our negative results rule out security for an average member of \mathcal{W}_Z . It may be possible to significantly improve parameters by only ruling out security for a single member W_z .

⁸Fuzzy extractors are defined to require high probability of correctness for all pairs w, w' . In the correlated setting, it may make sense to provide an average-case guarantee, where the probability of correctness is also over the draw of w, w' . Renner and Wolf use a smoothed notion of entropy that removes the δ fraction of the probability mass of $W = w|W' = w'$ with the most points to improve parameters under such a definition. In this work, we consider worst case correctness and use unsmoothed entropy.

Recall the security game of a fuzzy extractor: 1) the challenger specifies (SS, Rec) , 2) the adversary specifies a source $W_z \in \mathcal{W}_Z$ 3) The challenger wins if $\tilde{H}_\infty(W_z|\text{SS}(W_z)) \geq \tilde{m}$. Instead of just thinking of the uniform distribution over \mathcal{W}_Z , consider an arbitrary distribution V over elements of \mathcal{W}_Z . The minimax theorem says we can reverse which of these actions is announced first [vN28] if \mathcal{A} announces V instead of a single element W_z . That is, the following two player games have the same equilibrium:

<p>Experiment $\text{Exp}_1^{\mathcal{W}_Z}(\mathcal{A}, \mathcal{C}, \tilde{m})$: $(\text{SS}, \text{Rec}) \leftarrow \mathcal{C}(\mathcal{W}_Z)$ $W_z \leftarrow \mathcal{A}(\mathcal{W}, \text{SS}, \text{Rec})$ If $W_z \notin \mathcal{W}_Z$, \mathcal{C} wins. If $\tilde{H}_\infty(W_z \text{SS}(W_z)) \geq \tilde{m}$, \mathcal{C} wins. Else \mathcal{A} wins.</p>	<p>Experiment $\text{Exp}_2^{\mathcal{W}_Z}(\mathcal{A}, \mathcal{C}, \tilde{m})$: $V \leftarrow \mathcal{A}(\mathcal{W}_Z)$ $(\text{SS}, \text{Rec}) \leftarrow \mathcal{C}(V, \mathcal{W}_Z)$ $W_z \leftarrow V$ If $\tilde{H}_\infty(W_z \text{SS}(W_z)) \geq \tilde{m}$, \mathcal{C} wins. Else \mathcal{A} wins.</p>
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This means that showing security for a family of distributions \mathcal{W}_Z is equivalent to showing security for all distributions V when the distribution is known to the algorithms (SS, Rec) . In our negative results, the adversary uses the uniform distribution over \mathcal{W}_Z . However, it may be possible to improve parameters by using a different V .

C Fuzzy min-entropy

In this section, we provide proofs of statements in Section 2.2.

Proof of Proposition 2.6. Let W be a distribution where $H_{t,\infty}^{\text{fuzz}}(W) = m$. This means that there exists a point $w' \in \mathcal{M}$ such that $\Pr_{w \in W}[\text{dis}(w, w') \leq t] = 2^{-m}$. Consider the following distinguisher D :

- Input key, p .
- If $\text{Rep}(w', p) = \text{key}$, output 1.
- Else output 0.

Clearly, $\Pr[D(\text{Key}, P) = 1] \geq 2^{-m} - \delta$, while $\Pr[D(U_\kappa, P) = 1] = 1/2^{-\kappa}$. Thus,

$$\text{SD}((\text{Key}, P), (U_\kappa, P)) \geq \delta^D((\text{Key}, P), (U_\kappa, P)) \geq 2^{-m} - \delta - 2^{-\kappa}.$$

This completes the proof of Proposition 2.6. □

Proof of Lemma 2.7.

$$\begin{aligned} H_{t,\infty}^{\text{fuzz}}(W|P=p) &= -\log \left(\max_{w'} \sum_{w \in (W|P=p)|\text{dis}(w,w') \leq t} \Pr[W=w|P=p] \right) \\ &= -\log \left(\max_{w'} \sum_{w \in (W|P=p)|\text{dis}(w,w') \leq t} \frac{\Pr[W=w \wedge P=p]}{\Pr[P=p]} \right) \\ &\geq -\log \left(\max_{w'} \sum_{w \in W|\text{dis}(w,w') \leq t} \frac{\Pr[W=w]}{\Pr[P=p]} \right) \\ &= H_{t,\infty}^{\text{fuzz}}(W) + \log \Pr[P=p]. \end{aligned}$$

This completes the proof of Lemma 2.7. □

Proof of Lemma 2.9.

$$\begin{aligned}
\tilde{H}_{t,\infty}^{\text{fuzz}}(W|P) &= -\log \left(\mathbb{E}_{p \leftarrow P} \max_{w'} \sum_{w \in W|P=p|\text{dis}(w,w') \leq t} \Pr[W = w|P = p] \right) \\
&= -\log \left(\sum_p \max_{w'} \sum_{w \in W|P=p|\text{dis}(w,w') \leq t} \Pr[W = w|P = p] \Pr[P = p] \right) \\
&= -\log \left(\sum_p \max_{w'} \sum_{w \in W|P=p|\text{dis}(w,w') \leq t} \Pr[W = w \wedge P = p] \right) \\
&\geq -\log \left(\sum_p \max_{w'} \sum_{w \in W|P=p|\text{dis}(w,w') \leq t} \Pr[W = w] \right) \\
&\geq -\log \left(\sum_p \max_{w'} \sum_{w \in W|\text{dis}(w,w') \leq t} \Pr[W = w] \right) \\
&\geq -\log \left(2^{H_0(P)} \left(\max_{w'} \sum_{w \in W|\text{dis}(w,w') \leq t} \Pr[W = w] \right) \right) \\
&\geq H_{t,\infty}^{\text{fuzz}}(W) - H_0(P).
\end{aligned}$$

This completes the proof of Lemma 2.9. □

D Proof of Theorem 3.6

Proof. Throughout the proof we assume that $\ell = n$ is the number of levels. The proof can be carried out for an arbitrary ℓ but it leads to a complicated theorem statement.

Correctness: Fix some w, w' . If $\Pr[W = w] \leq 2^{-(m+\ell)} = 2^{-(m+n)}$, then w is simply transmitted to Rec and correctness is clear. When $\Pr[W = w] > 2^{-(m+n)}$ let L_i^* be the level of $\Pr[W = w]$.

Let W^* denote the set of elements of W in L_i within distance t of w' . The size of W^* is at most $\beta_{t,i}$. The choice of w, w' is independent of SS , so this set is independent of \mathcal{K}_i (it does effect the value of i but not the particular outcome from \mathcal{K}_i). The probability that another element w^* matches the hash is:

$$\begin{aligned}
&\Pr[\exists w^* \in W^* | w^* \neq w \wedge F_i(K, w^*) = F_i(K, w)] \\
&\leq \sum_{w^* \in W^* | w^* \neq w} \Pr[F_i(K, w^*) = F_i(K, w)] \\
&= \sum_{w^* \in W^* | w^* \neq w} \frac{1}{|R_i|} \leq \frac{\beta_{t,i} - 1}{|R_i|} \leq \frac{\beta_{t,i}}{|R_i|} = \delta
\end{aligned}$$

The first inequality is by union bound. The first equality follows from the universality of F_i . The second inequality follows since the number of neighbors is bounded by $\beta_{t,i}$.

Security: We now argue security of the construction. First note that the total weight of points whose probability is less than $2^{-(n+m)}$ is at most 2^{-m} (there are at most 2^n points in the distribution). Let 1_{low} be the indicator random variable for $\Pr[W = w] \leq 2^{-(n+m)}$. Then

$$\begin{aligned}\tilde{H}_\infty(W|\text{SS}(W)) &= -\log\left(\Pr[1_{\text{low}} = 1] * 1 + \Pr[1_{\text{low}} = 0]2^{-\tilde{H}_\infty(W|\text{SS}(W)\wedge 1_{\text{low}}=0)}\right) \\ &\geq -\log\left(2^{-m} + (1 - 2^{-m})2^{-\tilde{H}_\infty(W|\text{SS}(W)\wedge 1_{\text{low}}=0)}\right)\end{aligned}\quad (2)$$

For the remainder of the proof, we seek a bound on $\tilde{H}_\infty(W|\text{SS}(W) \wedge 1_{\text{low}} = 0)$. Let L_I be the random variable of the level information (in what follows, L_I takes values between m and $m + n$, where possible we omit the range of i for notational clarity).

$$\begin{aligned}\tilde{H}_\infty(W|\text{SS}(W) \wedge 1_{\text{low}} = 0) &= -\log\left(\mathbb{E}_i 2^{-\tilde{H}_\infty(W|\text{SS}(W)\wedge 1_{\text{low}}=0\wedge L_I=i)}\right) \\ &= -\log\left(\mathbb{E}_i 2^{-\tilde{H}_\infty(W|\mathcal{K}_i\wedge F_i(K_i,W)\wedge 1_{\text{low}}=0\wedge L_I=i)}\right) \\ &= -\log\left(\mathbb{E}_i 2^{-\tilde{H}_\infty(W|F_i(K_i,W)\wedge 1_{\text{low}}=0\wedge L_I=i)}\right) \\ &\geq -\log\left(\mathbb{E}_i 2^{-(\tilde{H}_\infty(W|1_{\text{low}}=0\wedge L_I=i)-\log(\beta_{t,i})+\log\delta)}\right) \\ &\geq -\log\left(\frac{1}{\delta}\mathbb{E}_i 2^{-(\tilde{H}_\infty(W|1_{\text{low}}=0\wedge L_I=i)-\log(\beta_{t,i}))}\right) \\ &\geq -\log\left(\mathbb{E}_i 2^{-(\tilde{H}_\infty(W|1_{\text{low}}=0\wedge L_I=i)-\log(\beta_{t,i}))}\right) - \log\frac{1}{\delta}\end{aligned}\quad (3)$$

The third line follows because the only dependence between \mathcal{K}_i and W is in i . The fourth line follows by [DORS08, Lemma 2.2]. We now show that the fuzzy min-entropy conditioned on the level information is proportional to $\beta_{t,i}$. For convenience denote by $J \stackrel{\text{def}}{=} (L_I \wedge 1_{\text{low}} = 0)$. That is, $\Pr[J = i] \stackrel{\text{def}}{=} \Pr[1_{\text{low}} = 0 \wedge L_I = i]$.

Claim D.1. $H_\infty(W|J = i) - \log\beta_{t,i} \geq H_{t,\infty}^{\text{fuzz}}(W|J = i) - 1$.

Proof.

$$\begin{aligned}
& \mathbf{H}_{t,\infty}^{\text{fuzz}}(W|J=i) = \\
& = -\log \left(\max_{w' \in \mathcal{M}} \left(\sum_{\substack{w \in W | \Pr[W=w] \in L_i \\ \wedge \text{dis}(w,w') \leq t}} \Pr[W=w|J=i] \right) \right) \\
& \leq -\log \left(\max_{w' \in \mathcal{M}} \left(\sum_{\substack{w \in W | \Pr[W=w] \in L_i \\ \wedge \text{dis}(w,w') \leq t}} \left(\min_{w^* \in W=w|J=i} \Pr[W=w^*|J=i] \right) \right) \right) \\
& \leq -\log \left(\max_{w' \in \mathcal{M}} \left(\sum_{\substack{w \in W | \Pr[W=w] \in L_i \\ \wedge \text{dis}(w,w') \leq t}} \left(\max_{w^* \in W=w|J=i} \Pr[W=w^*|J=i] \right) \right) \right) + 1 \\
& = -\log \left(\max_{w' \in \mathcal{M}} \left(\sum_{\substack{w \in W | \Pr[W=w] \in L_i \\ \wedge \text{dis}(w,w') \leq t}} 2^{-\mathbf{H}_\infty(W|J=i)} \right) \right) + 1 \\
& = \mathbf{H}_\infty(W|J=i) - \log \left(\max_{w' \in \mathcal{M}} \left(\sum_{\substack{w \in W | \Pr[W=w] \in L_i \\ \wedge \text{dis}(w,w') \leq t}} 1 \right) \right) + 1 \\
& = \mathbf{H}_\infty(W|J=i) \\
& - \log \left(\max_{w' \in \mathcal{M}} |\{w | \text{dis}(w,w') \leq t \wedge \Pr[W=w] \in L_i\}| \right) + 1 \\
& = \mathbf{H}_\infty(W|J=i) + 1 - \log \beta_{t,i}.
\end{aligned}$$

Where the fourth line follows by Lemma D.2, which follows the proof of Theorem 3.6. This completes the proof of Claim D.1. \square

We now return to the proof of Theorem 3.6. Together Equation 3 and Claim D.1 yield:

$$\begin{aligned}
\tilde{\mathbf{H}}_\infty(W|\mathbf{SS}(W) \wedge 1_{\text{low}}=0) & \geq -\log \left(\mathbb{E}_i 2^{-\mathbf{H}_{t,\infty}^{\text{fuzz}}(W|I=i \wedge 1_{\text{low}}=0)-1} \right) - \log \frac{1}{\delta} \\
& = -\log \left(\mathbb{E}_i 2^{-\mathbf{H}_{t,\infty}^{\text{fuzz}}(W|I=i \wedge 1_{\text{low}}=0)} \right) - 1 - \log \frac{1}{\delta} \\
& = \tilde{\mathbf{H}}_{t,\infty}^{\text{fuzz}}(W|I \wedge 1_{\text{low}}=0) - 1 - \log \frac{1}{\delta} \\
& \geq \mathbf{H}_{t,\infty}^{\text{fuzz}}(W|1_{\text{low}}=0) - 1 - \log n - \log \frac{1}{\delta} \\
& \geq \mathbf{H}_{t,\infty}^{\text{fuzz}}(W) - 1 - \log n - \log \frac{1}{\delta} + \log(1 - 2^{-m})
\end{aligned}$$

Where the second-to-last line follows by Lemma 2.9. The last line follows by Lemma 2.7 and by noting that $\Pr[1_{\text{low}} = 0] \geq (1 - 2^{-m})$. Returning to equation 2 one has:

$$\begin{aligned}
\tilde{H}_\infty(W|\text{SS}(W)) &\geq -\log\left(2^{-m} + (1 - 2^{-m})2^{-\tilde{H}_\infty(W|\text{SS}(W)\wedge 1_{\text{low}}=0)}\right) \\
&\geq -\log\left(2^{-m} + 2^{-(H_{t,\infty}^{\text{fuzz}}(W)-1-\log n-\log 1/\delta+2\log(1-2^{-m}))}\right) \\
&\geq -\log\min\{2^{-m}, 2^{-(H_{t,\infty}^{\text{fuzz}}(W)-1-\log n-\log 1/\delta+2\log(1-2^{-m}))}\} - 1 \\
&\geq H_{t,\infty}^{\text{fuzz}}(W) - 2 - \log n - \log 1/\delta + 2\log(1 - 2^{-m}) \\
&\geq H_{t,\infty}^{\text{fuzz}}(W) - 4 - \log n - \log 1/\delta
\end{aligned}$$

Where the fourth line follows from the third because $H_{t,\infty}^{\text{fuzz}}(W) \leq H_\infty(W) = m$. The last line follows from the fourth because if $m \geq 1$, then $\log(1 - 2^{-m}) \geq -1$ and if $m < 1$ the entire bound is vacuous as $H_{t,\infty}^{\text{fuzz}}(W) < 1$. \square

Lemma D.2. *Let W be a distribution and let $S \subset W$ such that for all $w_1, w_2 \in S$, $\Pr[W = w_1] \geq \Pr[W = w_2]/2$. Let J be as above, then for all $w_1, w_2 \in (S \wedge W|J = j)$, $\Pr[W = w_1|J = j] \geq \Pr[W = w_2|J = j]/2$.*

Proof.

$$\begin{aligned}
\Pr[W = w_1|J = j] &= \frac{\Pr[W = w_1 \wedge J = j]}{\Pr[J = j]} \\
&= \frac{\Pr[W = w_1]}{\Pr[J = j]} \\
&\geq \frac{\Pr[W = w_2]}{2\Pr[J = j]} \\
&= \frac{\Pr[W = w_2 \wedge J = j]}{2\Pr[J = j]} \\
&= \frac{\Pr[W = w_2|J = j]}{2}.
\end{aligned}$$

Where the first and last equality follow by Bayes' rule. The second and fourth equality follow by noting that $\Pr[W = w \wedge J = j] = \Pr[W = w \wedge 1_{\text{low}} = 0 \wedge I = i] = \Pr[W = w]$ for all $w \in L_i$. The inequality proceeds by assumption. This completes the proof of Lemma D.2. \square

E Proof of Theorem 4.1

Let $c' \leftarrow \text{Neigh}_t(c)$ sample a uniform point within distance t of c . The proof of Theorem 4.1 uses the definition of a Shannon code:

Definition E.1. *Let C be a set over space \mathcal{M} . We say that C is an (t, δ) -Shannon code if there exists a procedure Rec such that for all $c \in C$, $\Pr[c' \leftarrow \text{Neigh}_t(c) \wedge \text{Rec}(c') \neq c] \leq \delta$.*

We now prove item in the outline of Theorem 4.1.

Proposition E.2. *For each $W_z \in \mathcal{W}_Z$, $H_{t,\infty}^{\text{fuzz}}(W_z) = \omega(\log n)$.*

Proof. Consider some $W_z \in \mathcal{W}_Z$. The value w_1 is uniform in a field of size $\omega(\text{poly}(n))$, so $H_\infty(W_z) \geq \omega(\log n)$. We now show that for any $w, w' \in W_z$, $\text{dis}(w, w') = \gamma > t$. This shows that $H_{t, \infty}^{\text{fuzz}}(W_z) = H_\infty(W_z)$. Fix some $w, w' \in W_z$. Clearly, $w_1 \neq w'_1$, for any i , $w_i = a_i w_1 + b_i$ and $w'_i = a_i w'_1 + b_i$. Since $a_i \neq 0$, $a_i w_1 \neq a_i w'_1$ and thus $a_i w_1 + b_i \neq a_i w'_1 + b_i$. That is, $\text{dis}(w, w') = \gamma$. \square

Proposition E.3. V is the uniform distribution over \mathbb{F}^γ .

Proof. Consider some $w \in V$. Then w was drawn from an intermediate distribution W_z with coefficients $a_2, b_2, \dots, a_\gamma, b_\gamma$. The value w_1 is uniformly random and w_i are uniformly random since b_2, \dots, b_γ are uniformly distributed. \square

Lemma E.4. Fix some SS, Rec algorithm with error $\delta < 1/4$, then $\tilde{H}_0(V|\text{SS}(V)) \leq (\gamma - t + 1) \log |\mathbb{F}| + 1$.⁹

Proof. We assume that Rec is deterministic in our analysis. Any randomness necessary for the Rec algorithm can be provided by SS . This is the same as considering Rec that outputs any coin it flips. Since w, w' are independent of ss this does not effect correctness. Security is defined based on the output of SS so outputting the coins of Rec does not effect security. By the definition of correctness for (SS, Rec) ,

$$\forall w, w', \text{dis}(w, w') \leq t, \Pr_{ss \leftarrow \text{SS}(w)} [\text{Rec}(w', ss) \neq w] < \delta.$$

Fix some w . By Markov's inequality, there exists a set A_{ss} such that $\Pr[ss \in A_{ss}] \geq 1/2$ and $\forall ss \in A_{ss}$,

$$\frac{|\{w' | \text{dis}(w', w) \leq t \wedge \text{Rec}(w', ss) \neq w\}|}{|\{w' | \text{dis}(w', w) \leq t\}|} \leq 2\delta.$$

Consider some $ss^* \in A_{ss}$. We now show that $H_0(V|\text{SS}(V) = ss^*) \leq (\gamma - t + 1) \log |\mathbb{F}|$. For the sketched value w , at most a 2δ fraction of nearby w' do not map to w . Thus, this is also true for every value in $V|\text{SS}(V) = ss^*$. This makes the support of $V|\text{SS}(V) = ss^*$ a $(t, 2\delta)$ -Shannon code (see Definition E.1). This implies that for all $w_1, w_2 \in V|\text{SS}(V) = ss^*$, $\text{dis}(w_1, w_2) \geq t$ (since $2\delta < 1/2$). That is $V|\text{SS}(V) = ss^*$ is a set with minimum distance at least t .

By the Singleton bound, this implies that $H_0(V|\text{SS}(V) = ss^*) \leq (\gamma - t + 1) \log |\mathbb{F}|$. Averaging over $\text{SS}(V) = ss^*$ one has that $\tilde{H}_0(V|\text{SS}(V)) \leq (\gamma - t + 1) \log |\mathbb{F}| + 1$. \square

Lemma E.5. $\tilde{H}_0(V|\text{SS}(V), Z) < 2$ and thus $\tilde{H}_\infty(V|\text{SS}(V), Z) < 2$.

Proof. Recall that Z consists of 2γ coefficients and there are $(|\mathbb{F}| - 1)^{\gamma-1} |\mathbb{F}|^{\gamma-1}$ equally likely values for Z . As described above, the view of SS is a uniform distribution V . The length of this point is $|\mathbb{F}|^\gamma$. Conditioned on this information there are still many possible values for Z . That is,

$$\forall v, H_0(Z|V = v) = \log \left(\frac{(|\mathbb{F}| - 1)^{\gamma-1} |\mathbb{F}|^{\gamma-1}}{|\mathbb{F}|^\gamma} \right) = \log ((|\mathbb{F}| - 1)^{\gamma-1} / |\mathbb{F}|).$$

Consider two possible z_1, z_2 that are possible values of Z (having seen v). The distributions $V|Z = z_1$ and $V|Z = z_2$ intersect at one point (namely v).

We now show for any sketch algorithm there are few possible values of $V|Z$ in the support of $V|\text{SS}(V)$. The distributions $V|Z = z_1$ and $V|Z = z_2$ for possible z_1, z_2 (having seen v) overlap only at the point

⁹This result is an extension of lower bounds from [DORS08, Appendix C]. Dealing with imperfect correctness makes the bound more complicated. In particular, we can only argue about the average remaining support size.

v . This means for any $v^* \in V|\text{SS}(V)$ (other than the true v) there is at most one z such that $v^* \in V|\text{SS}(V), Z = z$.

The optimum strategy is to include these values uniformly from different Z values. We show this across different sketch values. Consider some fixed sketch value s and let $h_s = H_0(V|\text{SS}(V) = s)$. Recall that

$$\tilde{H}_0(V|\text{SS}(V)) = \log \mathbb{E}_{s \in \text{SS}(V)} 2^{H_0(V|\text{SS}(V)=s)} = \log \mathbb{E}_{s \in \text{SS}(V)} 2^{h_s}$$

Conditioned on seeing the point V there are $(|\mathbb{F}| - 1)^{\gamma-1}/|\mathbb{F}|$ possible values for Z with disjoint support outside of the sketched point. Consider these possible values for Z as containers to be filled with the $2^{h_{ss}}$ items (possible values of $V|\text{SS}(V) = ss$). Each container receives automatically receives one free point (all the distributions share v). The average number of items in each container is maximized when the containers are filled equally. That is, the average number of items in each container is bounded by the number of items divided by the number of container. That is,

$$\begin{aligned} \tilde{H}_0(V|Z, \text{SS}(V) = ss) &\leq \log \left(\frac{\# \text{ items} + \# \text{ containers}}{\# \text{ containers}} \right) \\ &= \log \left(\frac{2^{h_{ss}} |\mathbb{F}|}{(|\mathbb{F}| - 1)^{\gamma-1}} + 1 \right) \end{aligned}$$

Then averaging over the possible values of s , we have the following as long as $t \geq 4$:

$$\begin{aligned} \tilde{H}_0(V|Z, \text{SS}(V)) &= \log \mathbb{E}_{s \in \text{SS}(V)} 2^{\tilde{H}_0(V|\text{SS}(V)=ss, (Z|\text{SS}(V)=ss))} \\ &= \log \mathbb{E}_{s \in \text{SS}(V)} \left(\frac{2^{h_s} |\mathbb{F}|}{(|\mathbb{F}| - 1)^{\gamma-1}} + 1 \right) \\ &\leq \max \left\{ \log \left(\frac{|\mathbb{F}|}{(|\mathbb{F}| - 1)^{\gamma-1}} \mathbb{E}_{s \in \text{SS}(V)} 2^{h_s} \right) + 1, 1 \right\}. \end{aligned}$$

Where the inequality follows because $\log(x + 1) \leq \max\{\log x + 1, 1\}$ for $x \geq 0$. The left operand to max is bounded by 2 (bounding the max by 2). This argument uses a technical lemma that appears directly after (Lemma E.6).

$$\begin{aligned} &\log \left(\frac{|\mathbb{F}|}{(|\mathbb{F}| - 1)^{\gamma-1}} \mathbb{E}_{s \in \text{SS}(V)} 2^{h_s} \right) + 1 \\ &= \log |\mathbb{F}| - (\gamma - 1) \log(|\mathbb{F}| - 1) + \log \left(\mathbb{E}_{s \in \text{SS}(V)} 2^{h_s} \right) + 1 \\ &= \log |\mathbb{F}| - (\gamma - 1) \log(|\mathbb{F}| - 1) + \tilde{H}_0(V|\text{SS}(V)) + 1 \\ &\leq \log |\mathbb{F}| - (\gamma - 1) \log(|\mathbb{F}| - 1) + (\gamma - t + 1) \log |\mathbb{F}| + 2 \\ &\leq (\gamma - t + 2) \log |\mathbb{F}| - (\gamma - 1) \log(|\mathbb{F}| - 1) + 2 \\ &< (\gamma - t + 2) \log |\mathbb{F}| - (\gamma - 2) \log |\mathbb{F}| + 2 \quad (\text{by Lemma E.6}) \\ &\leq (4 - t) \log |\mathbb{F}| + 2 < 2. \end{aligned}$$

□

Lemma E.6. For any real numbers $\alpha \leq \eta$ with $\eta \geq e + 1$ (in particular, $\eta \geq 4$ suffices), the following holds: $\alpha \log(\eta - 1) > (\alpha - 1) \log \eta$.

Proof. Because $\eta - 1$ is positive, and $1 + x < e^x$ for positive x ,

$$1 + \frac{1}{\eta - 1} < e^{\frac{1}{\eta - 1}}.$$

Therefore,

$$\left(1 + \frac{1}{\eta - 1}\right)^{\alpha - 1} < e^{\frac{\alpha - 1}{\eta - 1}} \leq e < \eta - 1$$

(since $\alpha \leq \eta$). Multiplying both sides by $(\eta - 1)^{\alpha - 1}$, we obtain

$$\eta^{\alpha - 1} < (\eta - 1)^\alpha.$$

Taking the logarithm of both sides yields the statement of the lemma. \square

F Proof of Theorem 5.1

Proposition F.1. *For each $W_z \in \mathcal{W}_Z$, $H_{t, \infty}^{\text{fuzz}}(W_z) = \omega(\log n)$.*

Proof. Consider some fixed $W_z \in \mathcal{W}_Z$. The bits $w_{1, \dots, \nu} = x$ are uniform, so $H_\infty(W_z) = \omega(\log n)$. Recall that $t = o(n/\nu)$. Fix some $w, w' \in W_z$. Denote by x, x' the values that produce w, w' respectively. Clearly, $x \neq x'$. Thus, for any i , $a_i x + b_i \neq a_i x' + b_i$. This implies that $w_{i\nu+1, \dots, (i+1)\nu} \neq w'_{i\nu+1, \dots, (i+1)\nu}$. That is, at least one of the bits in each block differs between w and w' , and so $\text{dis}(w, w') \geq n/\nu > t$. Since no two values in the support of W_z lie in the same ball of radius t , we have $H_{t, \infty}^{\text{fuzz}}(W_z) = H_\infty(W_z) = \omega(\log n)$. \square

Proposition F.2. *V is the uniform distribution over \mathbb{F}^γ .*

Proof. Consider some $w \in V$ over $\{0, 1\}^n$. Then $w \leftarrow W_z$ for coefficients $z = a_2, b_2, \dots, a_\gamma, b_\gamma$. The value $w_{1, \dots, \nu} = x$ is uniformly random and $w_{i\nu+1, \dots, (i+1)\nu}$ are uniformly random since b_2, \dots, b_γ are random. \square

Lemma F.3. *Fix some (Gen, Rep) algorithm with $\kappa \geq 2$. There exists an information theoretic distinguisher between (Key, P, Z) and (U_κ, P, Z) with advantage $\epsilon = 1/8 - \text{ngl}(n)$.*

Proof. As in the proof of Theorem 4.1, we assume that Rep is deterministic. Denote by $(\text{Key}, P) \leftarrow \text{Gen}(V)$. By Markov's inequality, there exists a set A_ϵ such that $\Pr[p \in A_\epsilon] \geq 1/2$ and $\forall p \in A_\epsilon$,

$$(\text{Key}|P = p, P = p) \approx_{2\epsilon} (U_\kappa, P = p).$$

Consider some $p^* \in A_\epsilon$. The distribution $\text{Key}|P = p^*$ is the set of possible keys. The distribution $\text{Key}|P = p^*$ induces a partition on the metric space. That is, for every $w \in \mathcal{M}$, there exists a unique value key such that $\text{Rep}(w, p^*) = \text{key}$. Denote this partition by $Q_{p^*, \text{key}} = \{w | \text{Rep}(w, p^*) = \text{key}\}$.

There exists a set $\text{Key}_{\text{small}}$ where $|\text{Key}_{\text{small}}| \geq 2^{\kappa-1}$ such that for all $\text{key} \in \text{Key}_{\text{small}}$, $|Q_{p^*, \text{key}}| \leq \mathcal{M}/2^\kappa = 2^{n-\kappa}$. If not, then $\cup_{\text{key}} |Q_{p^*, \text{key}}| > |\mathcal{M}|$. For the remainder of the proof we restrict ourselves to elements in $\text{Key}_{\text{small}}$. Only points that are distance t from points outside of $Q_{p^*, r}$ are viable points in the metric space. These are points where all points within distance t map to the same key. We call this set the interior of $Q_{p^*, \text{key}}$:

$$\text{Inter}(Q_{p^*, \text{key}}) = \{w | \text{Rep}(w, p^*) = \text{key} \wedge (\forall w', \text{dis}(w, w') \leq t \wedge \text{Rep}(w', p^*) = \text{key})\},$$

The isoperimetric inequality says $\text{Inter}(Q_{p^*, \text{key}^*})$ must be of bounded size. This statement use the sets which are nearly balls in the Hamming space¹⁰:

Definition F.4. *A set S is a η -near ball if there exists a point x such that $B_{\eta-1}(x) \subseteq S \subseteq B_\eta(x)$.*

We now show for any key^* , $\text{Inter}(Q_{p^*, \text{key}^*})$ is small:

Lemma F.5. $|\text{Inter}(Q_{p^*, \text{key}^*})| \leq 2^{n-4\nu}$.

Proof. By the isoperimetric inequality on the Hamming space (we use a version due to [FF81, Theorem 1], the original result is due to Harper [Har66]), there exists a η -near ball S_{p^*, key^*} centered at 0^n and a near ball D , centered at 1^n , such that $|S_{p^*, \text{key}^*}| = |\text{Inter}(Q_{p^*, \text{key}^*})|$, $|D| = |Q_{p^*, \text{key}^*}^c|$ (where \cdot^c denotes the complement of a set) and $\forall s \in S_{p^*, \text{key}^*}, d \in D, \text{dis}(s, d) \geq t$ (that is, the distance between the sets is t). Since S_{p^*, key^*} is a near ball and the set D^c contains all points of distance less than t from S_{p^*, key^*} . Thus, the set $S_{p^*, \text{key}^*} \cup D^c$ contains a near ball of radius is $\eta + t - 1$. We now bound the size of S_{p^*, key^*} .

Recall that $|S_{p^*, \text{key}^*} \cup D^c| = |Q_{p^*, \text{key}^*}| \leq 2^{n-\kappa} \leq |\mathcal{M}|/2$. Since this set contains less than half the points in the metric space we know its radius at most $n/2$. This means that $|S_{p^*, \text{key}^*}|$ is a near ball of radius at most $n/2 - t + 1$. Let X denote a uniform string on $\{0, 1\}^n$. We use Hoeffding's inequality [Hoe63]:

$$\begin{aligned} |S_{p^*, \text{key}^*}| &\leq \{x | \text{dis}(x, 0) \leq \frac{n}{2} - t + 1\} \\ &= 2^n \Pr_{X \leftarrow \{0, 1\}^n} [wt(X) \leq (\frac{1}{2} - \frac{t-1}{n})n] \\ &\leq 2^n e^{-n((t-1)/n)^2} = 2^n e^{-4\nu} \leq 2^{n-4\nu}. \end{aligned}$$

where the second to last equality follows from the definition of t (as $4\nu\sqrt{n} + 1$) at the beginning of the proof of Theorem 5.1. \square

We have shown that $|\text{Inter}(Q_{p^*, \text{key}^*})| \leq 2^{n-4\nu}$. To complete the proof it suffices to show that for most values of the auxiliary information Z there are many parts Q_{p^*, key^*} that do not receive any points. Recall that Z consists of $2n/\nu$ coefficients and there are $(2^{n/\nu} - 1)^{\nu-1} 2^{n-\nu}$ equally likely values for Z . As described above, the view of Gen, Rep is a uniform distribution V . We know show there are many possible values for $Z|P = p^*$. The only information about Z is contained in the point $V = v$. The length of this point is 2^n . Conditioned on this information there are still many possible values for Z . That is,

$$\begin{aligned} \forall v, H_0(Z|V = v) &= \log \left(\frac{(2^{n/\nu} - 1)^{\nu-1} 2^{n-\nu}}{2^n} \right) \\ &= \log \frac{(2^{n/\nu} - 1)^{\nu-1}}{2^\nu} \\ &> \log \frac{(2^{n/\nu})^{\nu-2}}{2^\nu} \quad (\text{by Lemma E.6}) \\ &= \log \frac{2^{(n-2\nu)}}{2^\nu} = n - 3\nu. \end{aligned}$$

¹⁰In most statements of the isoperimetric inequality, this type of set is simply called a ball. We use the term near ball for emphasis.

Consider two possible z_1, z_2 that are possible values of Z . The distributions $V|Z = z_1$ and $V|Z = z_2$ intersect at one point (namely v).

This means that the **Gen** algorithm may include points for possible Z values into parts Q_{p^*, key^*} (other than v) and these values are disjoint. The optimum strategy is to include these values uniformly from different Z values. Consider the set of all preimages of Key_{small} denoted $Q_{small} = \cup_{\text{key} \in \text{Key}_{small}} \text{Inter}(Q_{\text{key}, p^*})$. Note that $Q_{small} \leq 2^{n-4\nu} |\text{Key}_{small}|$. We now show that the intersection between Q_{key, p^*} is small for most possible values z . As before each container (the values of z) receives one item for free (the point v).

$$\begin{aligned} \mathbb{E}_z |Q_{small} \cap (V|P = p^* \wedge Z = z)| &\leq \left(\frac{\# \text{ items} + \# \text{ containers}}{\# \text{ containers}} \right) \\ &\leq \frac{2^{n-4\nu} |\text{Key}_{small}|}{2^{n-3\nu}} + 1 \\ &= \frac{|\text{Key}_{small}|}{2^\nu} + 1 \end{aligned}$$

In expectation across Z ,

$$\frac{\frac{|\text{Key}_{small}|}{2^\nu} + 1}{|\text{Key}_{small}|} \leq \frac{1}{2^\nu} + \frac{1}{|\text{Key}_{small}|}$$

fraction of Key_{small} receive any support. We now present a distinguisher D_{p^*} for a particular p^* :

1. On input key, z .
2. Compute $V|P = p^* \wedge Z = z$ and $Q_{p^*, \text{key}}$.
3. If $(Q_{p^*, \text{key}} \cap V|P = p^* \wedge Z = z) = \emptyset$ output $b = 0$.
4. Else output $b = 1$.

The distinguisher $D(\text{key}, p, z)$ is formed by calling $D_p(\text{key}, z)$ when $p \in A_\epsilon$ and outputting a random

bit otherwise. The advantage of D is

$$\begin{aligned}
& \Pr[D(\mathbf{Key}, P, Z) = 1] - \Pr[D(U, P, Z) = 1] \\
&= (\Pr[D(\mathbf{Key}, P, Z) = 1|P \in A_\epsilon] - \Pr[D(U, P, Z) = 1|P \in A_\epsilon]) \Pr[P \in A_\epsilon] \\
&\geq \sum_{p^* \in A_\epsilon} \Pr[P = p^*] (1 - \Pr[D_{p^*}(U, Z) = 1]) \\
&\geq \sum_{p^* \in A_\epsilon} \Pr[P = p^*] (1 - \Pr[D_{p^*}(U, Z) = 1|U \in \mathbf{Key}_{small}] \Pr[U \in \mathbf{Key}_{small}] \\
&\quad - \Pr[U \notin \mathbf{Key}_{small}]) \\
&\geq \sum_{p^* \in A_\epsilon} \Pr[P = p^*] \left(1 - \left(\frac{1}{|\mathbf{Key}_{small}|} + \frac{1}{2^\nu}\right) \Pr[U \in \mathbf{Key}_{small}]\right) \\
&\quad - \Pr[U \notin \mathbf{Key}_{small}] \\
&\geq \sum_{p^* \in A_\epsilon} \Pr[P = p^*] \left(1 - \frac{1}{2^\nu} - \frac{1}{2} \Pr[U \in \mathbf{Key}_{small}] - \Pr[U \notin \mathbf{Key}_{small}]\right) \\
&\geq \sum_{p^* \in A_\epsilon} \Pr[P = p^*] \left(1 - \frac{1}{2^\nu} - \frac{1}{2} \Pr[U \in \mathbf{Key}_{small}] - \Pr[U \notin \mathbf{Key}_{small}]\right) \\
&\geq \sum_{p^* \in A_\epsilon} \Pr[P = p^*] \left(1 - \frac{1}{2^\nu} - 1 + \frac{1}{2} \Pr[U \in \mathbf{Key}_{small}]\right) \\
&\geq \sum_{p^* \in A_\epsilon} \Pr[P = p^*] (1/4 - \mathbf{ngl}(n)) \geq \frac{1}{8} - \mathbf{ngl}(n).
\end{aligned}$$

The sixth line follows since $|\mathbf{Key}_{small}| \geq 2^{\kappa-1} \geq 2$. The eighth line follows because $\Pr[U \in \mathbf{Key}_{small}] \geq 1/2$. The last inequality proceeds because $\Pr[P \in A_\epsilon] \geq 1/2$. This completes the proof of Lemma F.3. \square