

Achieving Differential Privacy with New Imperfect Randomness

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Abstract. We revisit the question of achieving differential privacy with realistic imperfect randomness. In the design of differentially private mechanisms, it's usually assumed that uniformly random source is available. However, in many situations it seems unrealistic, and one must deal with various imperfect random sources. Dodis et al. (CRYPTO'12) proposed that differential privacy can be achieved with Santha-Vazirani (SV) source via adding a stronger property called SV-consistent sampling and left open the question if differential privacy is possible with more realistic (i.e., less structured) sources than SV source. A new source, called Bias-Control Limited (BCL) source, introduced by Dodis (ICALP'01), as a generalization of the SV source and sequential bit-fixing source, is more realistic. Unfortunately, if we naively expand SV-consistent sampling to the BCL source, the expansion is hopeless to achieve differential privacy. One main reason is that SV-consistent sampling requires "consecutive" strings, while some strings can't be generated from "non-trivial" BCL source.

Motivated by this question, we introduce a new appealing property, called compact BCL-consistent sampling, the degeneration of which is different from SV-consistent sampling proposed by Dodis et al. We prove that if the mechanism based on the BCL source satisfies this property, then it's differentially private. Even if the BCL source is degenerated into the SV-source, our proof is much more intuitive and simpler than that of Dodis et al. Further, we construct explicit mechanisms using a new truncation technique as well as arithmetic coding. We also propose its concrete results for differential privacy and accuracy. While the results of [DY14] imply that if there *exist* differentially private mechanisms for imperfect randomness, then some parameters should have some constraints, ours show *explicit* construction of such mechanisms whose parameters match the prior constraints.

1 Introduction

Traditional cryptographic models take for granted the availability of perfect randomness, i.e., sources that output unbiased and independent random bits. However, in many settings this assumption seems unrealistic, and one must

deal with various imperfect sources of randomness. Some well known examples of such imperfect random sources are physical sources, biometric data, secrets with partial leakage, and group elements from Diffie-Hellman key exchange. To abstract this concept, several formal models of realistic imperfect sources have been described. Please see [DY14] for a summary. Roughly speaking, they can be divided into extractable and non-extractable. Extractable sources allow for deterministic extraction of nearly perfect randomness. Moreover, while the question of optimizing the extraction rate and efficiency has been very interesting, from the qualitative perspective such sources are good for any application where perfect randomness is sufficient. Unfortunately, it was quickly realized that many realistic sources are non-extractable [SV86, CG88, Dod01]. The simplest example is Santha-Vazirani (SV) source [SV86], which produces an infinite sequence of bits r_1, r_2, \dots , with the property that $\Pr[r_i = 0 \mid r_1 \dots r_{i-1}] \in [\frac{1-\delta}{2}, \frac{1+\delta}{2}]$, for any setting of the prior bits r_1, \dots, r_{i-1} . Santha and Vazirani [SV86] showed that there exists no deterministic extractor $\text{Enc} : \{0, 1\}^n \rightarrow \{0, 1\}$ capable of extracting even a *single* bit of bias *strictly* less than δ from the δ -SV source, irrespective of how many SV bits r_1, \dots, r_n it is willing to wait for.

Despite this pessimistic result, ruling out the “black-box compiler” from perfect to imperfect (e.g., SV) randomness for *all* applications, people still hope that specific “non-extractable” sources (e.g., SV sources), are sufficient for *concrete* applications. Indeed, there are already a series of positive results for simulating probabilistic polynomial-time algorithms [VV85, SV86, CG88, Zuc96, ACRT99] and *authentication* applications [MW97, DOPS04, DKRS06, ACM⁺14]. Unfortunately, the situation appears to be much less bright when dealing with *privacy* applications, such as encryption, commitment, zero-knowledge, and some others. Please see [DLMV12, DY14] for a survey. While a series of negative results seem to strongly point in the direction that privacy inherently requires extractable randomness, a recent work of Dodis et al. [DLMV12] put a slight dent into this consensus, by showing that SV sources are provably sufficient for achieving a more recent notion of privacy, called *differential privacy* (DP) [DMNS06].

The motivating scenario of differential privacy is a statistical database. The purpose of a privacy-preserving statistical database is to enable the user to learn released statistical facts without compromising the privacy of the individual users whose data is in the database. Differential privacy ensures the removal or addition of a single database item does not (substantially) affect the outcome of any analysis [Dwo08]. More formally, a differentially private mechanism $M(D, \mathbf{r})$ uses its randomness \mathbf{r} to “add enough noise” to the true answer $f(D)$, where D is some sensitive database of users, and f is some useful aggregate information (query) about the users of D . On one hand, to preserve individual users’ privacy, we want M to satisfy ξ -differential privacy, that is, for any neighboring databases D_1 and D_2 (i.e., D_1 and D_2 differ on a single record), and for any possible output z , $e^{-\xi} \leq \frac{\Pr_{\mathbf{r}}[M(D_1, f; \mathbf{r}) = z]}{\Pr_{\mathbf{r}}[M(D_2, f; \mathbf{r}) = z]} \leq e^{\xi}$ for small $\xi > 0$. On the other hand, to keep the utility (or accuracy) of M , we hope the expected value of $|f(D) - M(D, f; \mathbf{r})|$ over random \mathbf{r} to be as small as possible. Usually, we should make a tradeoff between differential privacy and utility.

Additive-noise mechanisms [DMNS06,GRS09,HT10] have the form $M(D, f; \mathbf{r}) = f(D) + X(\mathbf{r})$, where X is an appropriately chosen “noise” distribution added to guarantee ξ -DP. For instance, for counting queries, the right distribution is the Laplace distribution [DMNS06]. However, *we can not generate a “good enough” sample of the Laplace distribution with SV sources*. In fact, any accurate and private additive-noise mechanism for a source R implies the existence of a randomness extractor for R , essentially collapsing the notion of differential privacy to that of traditional privacy, and showing the impossibility of accurate and private additive-noise mechanisms for SV sources [DLMV12]. From another perspective, an additive-noise mechanism must satisfy $T_1 \cap T_2 = \emptyset$, based on which an SV adversary can always succeed in amplifying the ratio $\Pr[\mathbf{r} \in T_1] / \Pr[\mathbf{r} \in T_2]$ (see [DLMV12]), or $|\Pr[\mathbf{r} \in T_1] - \Pr[\mathbf{r} \in T_2]|$ (see [DY14]), where T_i is the set of coins \mathbf{r} with $M(D_i, f; \mathbf{r}) = z$ for $i = 1, 2$.

Dodis et al. [DLMV12] observed a necessary condition, called consistent sampling (i.e., informally, $|T_1 \cap T_2| \approx |T_1| \approx |T_2|$), to build SV-robust mechanisms. They also introduced another condition to match the bit-by-bit property of SV sources. The combination of the above two conditions is called SV-consistent sampling. They build a concrete accurate and private Laplace mechanism by using some truncation and arithmetic coding techniques. Such a mechanism is capable to work with all such distributions, provided that the utility ρ is now relaxed to be polynomial of $1/\varepsilon$, whose degree and coefficients depend on δ , but *not* on the size of the database D . Coupled with the impossibility of traditional privacy with SV sources, this result suggested a qualitative gap between traditional and differential privacy, but left the following open problem.

OPEN QUESTION. *Is differential privacy possible with more realistic (i.e., less structured) sources than SV sources?*

Dodis et al. [Dod01] introduced more realistic source, called Bias-Control Limited (BCL) source, denoted as $\mathcal{BCL}(\delta, b)$, which generates a sequence of bits x_1, x_2, \dots , where for $i = 1, 2, \dots$, the value of x_i can depend on x_1, \dots, x_{i-1} in one of the following two ways: (A) x_i is determined by x_1, \dots, x_{i-1} , but this happens for at most b bits, or (B) $\frac{1-\delta}{2} \leq \Pr[x_i = 1 \mid x_1, \dots, x_{i-1}] \leq \frac{1+\delta}{2}$, where $0 \leq \delta < 1$. (See Definition 2.) In particular, when $b = 0$, it degenerates into SV source of [SV86]; when $\delta = 0$, it yields the bit-fixing source of [LLS89]; when $b = 0$ and $\delta = 0$, it corresponds to the perfect randomness. If $b \neq 0$ and $\delta \neq 0$, we say the BCL source is non-trivial. The BCL source models the problem that each of the bits produced by a streaming source is unlikely to be perfectly random: slight errors (due to noise, measurement errors, and imperfections) of the source are inevitable, and the situation that some of the bits could have non-trivial dependencies on the previous bits (due to internal correlations, poor measurement or improper setup), to the point of being completely determined by them.

Hence, compared with SV source, the BCL source appears much more realistic, especially if the number of interventions b is somewhat moderate. Indeed, since it naturally (and realistically!) relaxes SV source, for which non-trivial differential privacy is possible, it will be interesting to see whether existing results

can be expanded using BCL sources (especially for reasonably high b raised by Dodis [Dod14]). Recently, Dodis and Yao [DY14] have shown an impossibility result for BCL source: when $b \geq \Omega(\log(\xi\rho)/\delta)$, it's impossible to achieve $(\mathcal{BCL}(\delta, b), \xi)$ -differentially private (see Definition 3) and (\mathcal{U}, ρ) -accurate (see Definition 4) mechanism for Hamming weight queries. In other words, if there exists a $(\mathcal{BCL}(\delta, b), \xi)$ -differentially private and (\mathcal{U}, ρ) -accurate mechanism for Hamming weight queries, then $b \leq O(\log(\xi\rho)/\delta)$. This result gives us a bit hope to design differentially private and accurate mechanisms for some b .

OUR RESULTS AND TECHNIQUES.

We try to naturally expand SV-consistent sampling to BCL-consistent sampling, but can't get positive results. It's not surprising, as the "interval" property (see Definition 9) is crucial to achieve SV-differential privacy, while the mechanism based on $\mathcal{BCL}(\delta, b)$ with $b \neq 0$ can't be an interval one.

Essentially, to achieve differential privacy, we need to restrict $\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_1 \setminus T_2] / \Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_2]$. Similar to [DLMV12], consistent sampling is still a necessary condition for building BCL-robust, differentially private mechanisms. From the generation procedure of $\mathcal{BCL}(\delta, b, n)$, we can upper bound the numerator and lower bound the denominator by introducing the common prefix \mathbf{u} of T_1 and T_2 . Instead of limiting $|\text{SUFFIX}(\mathbf{u}, n)|/|T_1 \cup T_2| = 2^{n-|\mathbf{u}|}/|T_1 \cup T_2|$ as in [DLMV12], we limit $n - |\mathbf{u}|$. The concept of compact BCL-consistent sampling (Definition 10) emerges from this motivation.

However, we are confronted with some difficulties to construct explicit differentially private mechanisms. According to the method of yielding finite precision mechanisms in [DLMV12], we can't upper bound $n - |\mathbf{u}|$ as a constant! To solve this problem, we find a new truncation trick and hence design a new mechanism (see Section 4.1). Our contributions are as follows.

- We introduce a new concept, called compact BCL-consistent sampling (see Definition 10), to study differentially private mechanisms. It should be noted that if $b = 0$, the degenerated BCL-consistent sampling is not the same as the SV-consistent sampling (see Definition 9) proposed by [DLMV12].
- We prove that if the BCL source satisfies this property, then the corresponding mechanism is differentially private (see Theorem 1). Even if the BCL source is degenerated into SV source, compared with [DLMV12], our proof is much more intuitive and simpler (see Theorem 1 with $b = 0$ and Theorem 4.4 of [DLMV12]).
- We use a new truncation technique and arithmetic coding in the design of a finite-precision mechanism to satisfy compact BCL-consistent sampling (see Section 4.1).
- We also give rigorous proofs about differential privacy and accuracy of this kind of mechanism (Theorems 2 and 3).
- While the result of [DY14] implies if there exists a $(\mathcal{BCL}(\delta, b), \xi)$ -differentially private and (\mathcal{U}, ρ) -accurate mechanism for the Hamming weight queries, then the parameters should satisfy $\rho > \frac{2^{b \cdot \log(1+\delta)-9}}{\xi}$, we build such explicit mechanisms and the parameters match the above condition (Theorem 4).

2 Preliminaries

Let $\{0, 1\}^* \stackrel{\text{def}}{=} \bigcup_{m \in \mathbb{Z}^+} \{0, 1\}^m$. We consider a distribution over $\{0, 1\}^*$ as continuously outputting (possibly correlated) bits. We call a family \mathcal{R} of distributions over $\{0, 1\}^*$ a source. Denote \mathcal{U} as the uniform source, which is the set containing only the distribution U on $\{0, 1\}^*$ that samples each bit independently and uniformly at random. For a set S , we write U_S to denote the uniform distribution over S . For a distribution or a random variable R , let $\mathbf{r} \leftarrow R$ denote the operation of sampling a random \mathbf{r} according to R . For a positive integer n , let $[n] \stackrel{\text{def}}{=} \{1, 2, \dots, n\}$. Denote $\lfloor \cdot \rfloor$ as the nearest integer function.

Definition 1. (*[SV86]*) Let x_1, x_2, \dots be a sequence of Boolean random variables and $0 \leq \delta < 1$. A probability distribution $X = x_1 x_2 \dots$ over $\{0, 1\}^*$ is a δ -Santha-Vazirani (SV) distribution, denoted by $SV(\delta)$, if for all $i \in \mathbb{Z}^+$ and for every string s of length $i-1$, we have $\frac{1-\delta}{2} \leq \Pr[x_i = 1 \mid x_1 x_2 \dots x_{i-1} = s] \leq \frac{1+\delta}{2}$.

We define the δ -Santha-Vazirani source $\mathcal{SV}(\delta)$ to be the set of all δ -SV distributions. For $SV(\delta) \in \mathcal{SV}(\delta)$, we define $SV(\delta, n)$ as $SV(\delta)$ restricted to the first n coins $x_1 x_2 \dots x_n$. We let $\mathcal{SV}(\delta, n)$ be the set of all distributions $SV(\delta, n)$.

Definition 2. (*[Dod01]*) Let x_1, x_2, \dots be a sequence of Boolean random variables and $0 \leq \delta < 1$. A probability distribution $X = x_1 x_2 \dots$ over $\{0, 1\}^*$ is a (δ, b) -Bias-Control Limited (BCL) distribution, denoted by $BCL(\delta, b)$, if for all $i \in \mathbb{Z}^+$ and for every string s of length $i-1$, the value of x_i can depend on x_1, x_2, \dots, x_{i-1} in one of the following two ways:

(A) x_i is determined by x_1, \dots, x_{i-1} , but this happens for at most b bits. This process of determining a bit is called intervention.

(B) $\frac{1-\delta}{2} \leq \Pr[x_i = 1 \mid x_1 x_2 \dots x_{i-1} = s] \leq \frac{1+\delta}{2}$.

We define the (δ, b) -Bias-Control Limited source $\mathcal{BCL}(\delta, b)$ to be the set of all (δ, b) -BCL distributions. For a distribution $BCL(\delta, b) \in \mathcal{BCL}(\delta, b)$, we define $BCL(\delta, b, n)$ as the distribution $BCL(\delta, b)$ restricted to the first n coins $x_1 x_2 \dots x_n$. We let $\mathcal{BCL}(\delta, b, n)$ be the set of all distributions $BCL(\delta, b, n)$.

This source models the facts that physical sources can never produce completely perfect bits and some of the bits generated by a physical source could be determined from the previous bits.

Remark 1. In particular, if $b = 0$, $\mathcal{BCL}(\delta, b, n)$ degenerates into $\mathcal{SV}(\delta, n)$ [SV86]; if $\delta = 0$, it yields the sequential-bit-fixing source of Lichtenstein, Linial, and Saks [LLS89]. The definitions and results in the remainder of this paper can be degenerated into the counterparts for SV and sequential bit-fixing sources.

Consider a statistical database as an array of rows from some countable set. Two databases are neighboring if they differ in exactly one row. Let \mathcal{D} be the space of all databases. For simplicity, we only consider the query function $f : \mathcal{D} \rightarrow \mathbb{Z}$. Recall some concepts mentioned in [DLMV12] as follows.

Definition 3. Let $\xi \geq 0$, \mathcal{R} be a source, and $\mathcal{F} = \{f : \mathcal{D} \rightarrow \mathbb{Z}\}$ be a family of functions. A mechanism M is (\mathcal{R}, ξ) -differentially private for \mathcal{F} if for all neighboring databases $D_1, D_2 \in \mathcal{D}$, all $f \in \mathcal{F}$, all possible outputs $z \in \mathbb{Z}$, and all distributions $R \in \mathcal{R}$: $\Pr_{\mathbf{r} \leftarrow R}[M(D_1, f; \mathbf{r}) = z] / \Pr_{\mathbf{r} \leftarrow R}[M(D_2, f; \mathbf{r}) = z] \leq 1 + \xi$.

In what follows we employ the upper bound of the ratio of probabilities introduced in [DLMV12] other than the traditional upper bound “ e^ξ ” to make later calculations a little simpler. It’s reasonable since when $\xi \in [0, 1]$, which is the main useful range, we have $e^\xi \approx 1 + \xi$, and when $\xi \geq 0$, we always have $1 + \xi \leq e^\xi$.

Remark 2. As observed by Dodis et al. [DLMV12], here we assume that the randomness \mathbf{r} as input of the mechanism M is in $\{0, 1\}^*$, i.e., M has at its disposal a possibly infinite number of random bits, but for two neighboring databases $D_1, D_2 \in \mathcal{D}$, query $f \in \mathcal{F}$, and fixed outcome z , M needs only a finite number of coins $n \stackrel{\text{def}}{=} \tilde{\tau}(D_1, D_2, f, z)$, where $\tilde{\tau}$ is a function, to determine whether $M(D_1, f) = z$ and $M(D_2, f) = z$. Furthermore, we assume that if $M(D_1, f; \mathbf{r}) = z$ and $M(D_2, f; \tilde{\mathbf{r}}) = z$ where $\mathbf{r}, \tilde{\mathbf{r}} \in \{0, 1\}^n$, then providing M with extra coins doesn’t change the output. Namely, for any \mathbf{r}' with \mathbf{r} as its prefix and $\tilde{\mathbf{r}}'$ with $\tilde{\mathbf{r}}$ as its prefix, we still have $M(D_1, f; \mathbf{r}') = z$ and $M(D_2, f; \tilde{\mathbf{r}}') = z$.

Definition 4. Let $\rho > 0$, \mathcal{R} be a source, and $\mathcal{F} = \{f : \mathcal{D} \rightarrow \mathbb{Z}\}$ be a family of functions. A mechanism M has (\mathcal{R}, ρ) -utility (or accuracy) if for all databases $D \in \mathcal{D}$, all queries $f \in \mathcal{F}$, and all distributions $R \in \mathcal{R}$: $\mathbb{E}_{\mathbf{r} \leftarrow R}[|M(D, f; \mathbf{r}) - f(D)|] \leq \rho$.

One core problem in the area of differential privacy is to design accurate and private mechanisms.

Definition 5. We say a function family \mathcal{F} admits accurate and private mechanisms w.r.t. \mathcal{R} if there exists function $g(\cdot)$ s.t. for all $\xi > 0$ there exists mechanism $M_{(\xi)}$ that is (\mathcal{R}, ξ) -differentially private and has $(\mathcal{R}, g(\xi))$ -utility. $\mathcal{M} = \{M_{(\xi)}\}$ is called a class of accurate and private mechanisms for \mathcal{F} w.r.t. \mathcal{R} .

Though there are already some infinite additive mechanisms based on gaussian, binomial, and Laplace distributions, we must specify how to approximate them under finite precision in practice. When perfect randomness is available, we can simply approximate a continuous sample within some “good enough” finite precision, which is omitted in most differential privacy papers. Dodis et al. build finite-precision mechanisms under imperfect randomness $\mathcal{SV}(\delta)$ [DLMV12].

Definition 6. For query $f : \mathcal{D} \rightarrow \mathbb{Z}$, the sensitivity of f is defined as $\Delta f \stackrel{\text{def}}{=} \max_{D_1, D_2} \|f(D_1) - f(D_2)\|$ for all neighboring databases $D_1, D_2 \in \mathcal{D}$. For $d \in \mathbb{Z}^+$, denote $\mathcal{F}_d = \{f : \mathcal{D} \rightarrow \mathbb{Z} \mid \Delta f \leq d\}$.

For clarity, in this paper we only consider the case $d = 1$. It’s straightforward to extend all our results to any sensitivity bound d .

Definition 7. The Laplace (or double exponential) distribution with mean μ and standard deviation $\frac{\sqrt{2}}{\varepsilon}$, denoted as $\text{Lap}_{\mu, \frac{1}{\varepsilon}}$, has probability density function $\text{PDF}_{\mu, \frac{1}{\varepsilon}}^{\text{Lap}}(x) = \frac{\varepsilon}{2} \cdot e^{-\varepsilon|x-\mu|}$. The cumulative distribution function is given by $\text{CDF}_{\mu, \frac{1}{\varepsilon}}^{\text{Lap}}(x) = \frac{1}{2} + \frac{1}{2} \cdot \text{sgn}(x - \mu) \cdot (1 - e^{-\varepsilon|x-\mu|})$.

If a random variable X has this distribution, denote $X \sim \text{Lap}_{\mu, \frac{1}{\varepsilon}}$ ¹.

3 Compact BCL-Consistent Sampling and SV-Consistent Sampling

Dodis et al. [DLMV12] introduced SV-consistent sampling. However, the proof of “SV-consistent sampling implies differential privacy” (see Theorem 4.4 in [DLMV12] for details) is complex. Moreover, its natural expansion to BCL source is difficult and unknown to achieve differential privacy, as the proof of Theorem 4.4 in [DLMV12] depends on the fact that the values in T_2 (resp. T_1) constitutes an interval (see Definition 9), while it may not be the case for BCL sources.

In this section, we introduce the concept of compact (ζ, c) -BCL-consistent sampling. If $b = 0$, we get the concept of compact (ζ', c) -SV-consistent sampling. Then we observe that these concepts are sufficient to design finite-precision differentially private and accurate mechanisms based on BCL and SV sources.

Consider a mechanism M with randomness space $\{0, 1\}^*$. Denote $\tilde{T}(D_i, f, z) \stackrel{\text{def}}{=} \{\mathbf{r} \in \{0, 1\}^* \mid z = M(D_i, f; \mathbf{r})\}$, where $i \in \{1, 2\}$, as the set of all coins such that M outputs z when running on two neighboring databases D_1 and D_2 , query f , and randomness \mathbf{r} . It should be noted that in our model only $n \stackrel{\text{def}}{=} \tilde{\tau}(D_1, D_2, f, z)$ coins need to be sampled to determine if $M(D_1, f) = z$ and $M(D_2, f) = z$. Therefore, let $T(D_i, f, z) \stackrel{\text{def}}{=} \{\mathbf{r} \in \{0, 1\}^n \mid z = M(D_i, f; \mathbf{r})\}$ for $i \in \{1, 2\}$, wlog, we assume that $\tilde{T}(D_i, f, z) = T(D_i, f, z)$ for $i \in \{1, 2\}$.

For $m \in \mathbb{Z}^+$ and $\mathbf{x} = x_1, \dots, x_m \in \{0, 1\}^m$, let $\text{SUFFIX}(\mathbf{x}) \stackrel{\text{def}}{=} \{\mathbf{y} = y_1, y_2, \dots \in \{0, 1\}^* \mid x_i = y_i \text{ for all } i \in [m]\}$ as the set of all bit strings having \mathbf{x} as a prefix. For $n \in \mathbb{Z}^+$ where $n \geq m$, let $\text{SUFFIX}(\mathbf{x}, n) \stackrel{\text{def}}{=} \text{SUFFIX}(\mathbf{x}) \cap \{0, 1\}^n$.

Now consider two neighboring databases D_1 and D_2 , $f \in \mathcal{F}$, and a possible outcome z . Denote $n \stackrel{\text{def}}{=} \tilde{\tau}(D_1, D_2, f, z)$. Let $T_1 \stackrel{\text{def}}{=} T(D_1, f, z)$, $T_2 \stackrel{\text{def}}{=} T(D_2, f, z)$, and $\mathbf{u} \stackrel{\text{def}}{=} \text{argmax}\{|\mathbf{u}'| \mid \mathbf{u}' \in \{0, 1\}^{\leq n} \text{ and } T_1 \cup T_2 \subseteq \text{SUFFIX}(\mathbf{u}', n)\}$. Then the ratio is

$$\frac{\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b, n)}[\mathbf{r} \in T_1 \setminus T_2]}{\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b, n)}[\mathbf{r} \in T_2]} = \frac{\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b, n)}[\mathbf{r} \in T_1 \setminus T_2 \mid \mathbf{r} \in \text{SUFFIX}(\mathbf{u})]}{\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b, n)}[\mathbf{r} \in T_2 \mid \mathbf{r} \in \text{SUFFIX}(\mathbf{u})]}.$$

Since SV source and BCL source both generate strings bit by bit, the calculation of the ratio can be simplified.

Recall that the concepts of consistent sampling, interval mechanism, and SV-consistent sampling [DLMV12] are as follows.

¹ In this paper, we only consider the case that $\frac{1}{\varepsilon} \in \mathbb{Z}$.

Definition 8. A mechanism M has ζ -consistent sampling if for all potential outputs $z \in \mathbb{Z}$, all queries $f \in \mathcal{F}$, all neighboring databases $D_1, D_2 \in \mathcal{D}$: $\frac{|T_1 \setminus T_2|}{|T_2|} \leq \zeta$, where $T_1 \stackrel{def}{=} T(D_1, f, z)$, $T_2 \stackrel{def}{=} T(D_2, f, z) \neq \emptyset$.

Definition 9. Let $\tilde{c} > 1$ and $\zeta > 0$. We say a mechanism M is an interval mechanism if for all $f \in \mathcal{F}$, all $D \in \mathcal{D}$, and all possible outcomes $z \in \mathbb{Z}$, $T \stackrel{def}{=} T(D, f, z)$, set $\{\sum_{i=1}^n r_i \cdot 2^{n-i} \mid r_1 \dots r_n \in T\}$ contains consecutive integers. An interval mechanism has (ζ, \tilde{c}) -SV-consistent sampling if it has ζ -consistent sampling and for all $f \in \mathcal{F}$, all neighboring databases $D_1, D_2 \in \mathcal{D}$, all possible outcomes $z \in \mathbb{Z}$, which define T_1, T_2 , and \mathbf{u} as above, $\frac{|SUFFIX(\mathbf{u}, n)|}{|T_1 \cup T_2|} \leq \tilde{c}$ holds.

Note that when $b \neq 0$, $\mathcal{BCL}(\delta, b, n)$ can't generate all n -bit strings. The corresponding mechanism can't be an interval mechanism. Though Dodis et al. [DLMV12] proposed that if M has (ζ, \tilde{c}) -SV-consistent sampling, then M is $(\mathcal{SV}(\delta), \xi)$ -differentially private. In that proof, the "interval" property is a basic condition, we can't follow that thought. We resort to a new property instead.

Definition 10. Let c be a constant and $\zeta > 0$. A mechanism is a compact (ζ, c) -BCL-consistent sampling mechanism if it has ζ -consistent sampling and for all queries $f \in \mathcal{F}$, all neighboring databases $D_1, D_2 \in \mathcal{D}$, and all possible outcomes $z \in \mathbb{Z}$, which define T_1, T_2 and \mathbf{u} as above, we have $n - |\mathbf{u}| \leq c$.

Theorem 1. If Mechanism M is a compact (ζ, c) -BCL-consistent sampling mechanism for (δ, b) -BCL-sources, then M is $(\mathcal{BCL}(\delta, b), \xi)$ -differentially private, where $\xi \leq (\frac{1+\delta}{1-\delta})^c \cdot [\frac{1}{2}(1+\delta)]^{-b} \cdot \zeta$. In particular, for $\delta \in [0, 1)$, and $c = O(1)$, we have $\lim_{\zeta \rightarrow 0} (\frac{1+\delta}{1-\delta})^c \cdot [\frac{1}{2}(1+\delta)]^{-b} \cdot \zeta = 0$.

Proof. Assume that $\frac{|T_1 \setminus T_2|}{|T_2|} \leq \zeta$ and $n - |\mathbf{u}| \leq c$. For any $\mathbf{r}, \mathbf{r}' \in \{0, 1\}^n$, denote $\mathbf{r} = r_1 \dots r_n$ and $\mathbf{r}' = r'_1 \dots r'_n$ where $r_i, r'_i \in \{0, 1\}$ for $i \in [n]$. Then

$$\begin{aligned} \frac{\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_1 \setminus T_2]}{\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_2]} &= \frac{\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_1 \setminus T_2 \mid \mathbf{r} \in SUFFIX(\mathbf{u})]}{\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_2 \mid \mathbf{r} \in SUFFIX(\mathbf{u})]} \\ &= \frac{\sum_{\mathbf{r}' \in T_1 \setminus T_2} \Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} = \mathbf{r}' \mid \mathbf{r}' \in SUFFIX(\mathbf{u})]}{\sum_{\mathbf{r}' \in T_2} \Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} = \mathbf{r}' \mid \mathbf{r}' \in SUFFIX(\mathbf{u})]} \end{aligned}$$

For any fixed $\mathbf{r}' \in \{0, 1\}^n$, we have $\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} = \mathbf{r}' \mid \mathbf{r}' \in SUFFIX(\mathbf{u})] = \Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[r_{|\mathbf{u}|+1} = r'_{|\mathbf{u}|+1} \mid r_1 \dots r_{|\mathbf{u}|} = \mathbf{u}] \times \dots \times \Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[r_n = r'_n \mid r_1 \dots r_{|\mathbf{u}|} r_{|\mathbf{u}|+1} \dots r_{n-1} = \mathbf{u} r'_{|\mathbf{u}|+1} \dots r'_{n-1}]$. Therefore, $\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_2] \geq [\frac{1}{2}(1-\delta)]^{n-|\mathbf{u}|} \cdot |T_2|$ and $\Pr_{\mathbf{r} \leftarrow \mathcal{BCL}(\delta, b, n)}[\mathbf{r} \in T_1 \setminus T_2] \leq [\frac{1}{2}(1+\delta)]^{n-|\mathbf{u}|} \cdot |T_1 \setminus T_2|$. Correspondingly,

$$\begin{aligned}
 & \frac{\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b, n)}[\mathbf{r} \in T_1 \setminus T_2]}{\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b, n)}[\mathbf{r} \in T_2]} \leq \frac{[\frac{1}{2}(1+\delta)]^{n-|\mathbf{u}|-b}}{[\frac{1}{2}(1-\delta)]^{n-|\mathbf{u}|}} \cdot \frac{|T_1 \setminus T_2|}{|T_2|} \\
 & \leq \left(\frac{1+\delta}{1-\delta}\right)^{n-|\mathbf{u}|} \cdot [\frac{1}{2}(1+\delta)]^{-b} \cdot \zeta \leq \left(\frac{1+\delta}{1-\delta}\right)^c \cdot [\frac{1}{2}(1+\delta)]^{-b} \cdot \zeta
 \end{aligned}$$

□

Remark 3. When $b = 0$, Theorem 1 holds for SV sources, while Theorem 4.4 of [DLMV12] can be naturally expanded for BCL sources, mainly because of the “consecutive strings” requirement in the latter. Further, the proof here is much simpler and more intuitive than that of [DLMV12].

4 Accurate and Private BCLCS Mechanisms

In this section, we construct finite-precision mechanisms that achieve compact $(\zeta, O(1))$ -BCL-consistent sampling with sensitivity 1. We also propose that the precision of the specific mechanism based on Laplace distribution introduced by Dodis et al. [DLMV12] can be modified via this technique such that it becomes a compact SV-consistent sampling mechanism. Then by Theorems 1 and 2, the mechanism here and the modified mechanism of [DLMV12] are $(\text{BCL}(\delta, b), \xi)$ -differentially private and $(\text{SV}(\delta), \xi')$ -differentially private, where ξ' is a specific ξ by letting $b = 0$. We also show that these mechanisms have good bound on utility when the random sampling is generated from the BCL source.

4.1 Explicit Construction

We construct an infinite-precision mechanism, called $M_\varepsilon^{\text{CBCLCS}}$, then modify it to a finite precision one, denoted as $\overline{M}_\varepsilon^{\text{CBCLCS}}$. Recall that some truncation method was proposed in [DLMV12] in order to get a finite mechanism, which leads to the non-intuitive notion of SV-consistent sampling. However, it can't be transplanted to BCL sources. In this section, we develop another truncation technique. The finite-precision mechanism is designed as follows.

Explicit Construction of the Mechanism:

Step 1 On input any neighboring databases $D_1, D_2 \in \mathcal{D}$, $f \in \mathcal{F}$, the infinite-precision mechanism $M_\varepsilon^{\text{CBCLCS}}$ computes $f(D_1)$ and $f(D_2)$. Without loss of generality, assume that $f(D_1) = y$ and $f(D_2) = y - 1$. $M_\varepsilon^{\text{CBCLCS}}(D_1, f)$ (resp. $M_\varepsilon^{\text{CBCLCS}}(D_2, f)$) outputs $z_1 \leftarrow \frac{1}{\varepsilon} \cdot \lceil \varepsilon \cdot (y + \text{Lap}_{0, \frac{1}{\varepsilon}}) \rceil$ (resp. $z_2 \leftarrow \frac{1}{\varepsilon} \cdot \lceil \varepsilon \cdot (y - 1 + \text{Lap}_{0, \frac{1}{\varepsilon}}) \rceil$). Denote Z_y (resp. Z_{y-1}) as the output distribution of $M_\varepsilon^{\text{CBCLCS}}(D_1, f)$ (resp. $M_\varepsilon^{\text{CBCLCS}}(D_2, f)$) using arithmetic coding (see [DLMV12]).

Step 2 Suppose that y is fixed. Let $s_y(k) \stackrel{\text{def}}{=} \text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}\left(\frac{k+\frac{1}{2}}{\varepsilon}\right)$ and $s_{y-1}(k) \stackrel{\text{def}}{=} \text{CDF}_{y-1, \frac{1}{\varepsilon}}^{\text{Lap}}\left(\frac{k+\frac{1}{2}}{\varepsilon}\right)$ for all $k \in \mathbb{Z}$. Denote $I_y(k) = [s_y(k-1), s_y(k))$ and $I_{y-1}(k) =$

$[s_{y-1}(k-1), s_{y-1}(k)]$. Let $\bar{s}_{y-1}(k-1)$ (resp. $\bar{s}_{y-1}(k)$) be $s_{y-1}(k-1)$ (resp. $s_{y-1}(k)$), rounded to the first $n \stackrel{\text{def}}{=} \tau(\min(f(D_1), f(D_2)), k/\varepsilon) = \tau(y-1, k/\varepsilon)$ bits after the binary point. We round $s_y(k-1)$ (resp. $s_y(k)$) to the first $n \stackrel{\text{def}}{=} \tau(y-1, k/\varepsilon)$ bits after the binary point. Assume the binary decimal representation of the rounded $s_y(k-1)$ (resp. $s_y(k)$) is $0.r_1r_2\dots r_n$ (resp. $0.q_1q_2\dots q_n$), then let $\bar{s}_y(k-1) = 0.r_1r_2\dots r_n + 0.r'_1r'_2\dots r'_n$ (resp. $\bar{s}_y(k) = 0.q_1q_2\dots q_n + 0.q'_1q'_2\dots q'_n$), where $r'_i = 0$ for $i \in [n-1]$, and $r'_n = 1$ (resp. $q'_i = 0$ for $i \in [n-1]$ and $q'_n = 1$). Denote $\bar{I}_{y-1}(k) = [\bar{s}_{y-1}(k-1), \bar{s}_{y-1}(k)]$ and $\bar{I}_y(k) = [\bar{s}_y(k-1), \bar{s}_y(k)]$.

Step 3 Denote \bar{Z}_y (resp. \bar{Z}_{y-1}) as the output distribution of $\bar{M}_\varepsilon^{\text{CBCLCS}}(D_1, f)$ (resp. $\bar{M}_\varepsilon^{\text{CBCLCS}}(D_2, f)$), which approximates Z_y (resp. Z_{y-1}). For any sequence $\mathbf{r} = r_1, \dots, r_n \in \{0, 1\}^n$, the real representation of \mathbf{r} is $\text{REAL}(\mathbf{r}) \stackrel{\text{def}}{=} 0.r_1\dots r_n \in [0, 1]$. We obtain distribution \bar{Z}_y (resp. \bar{Z}_{y-1}) by sampling a sequence of bits $\mathbf{r} \in \{0, 1\}^n$ (resp. $\mathbf{r}' \in \{0, 1\}^n$) from a distribution $\text{BCL}(\delta, b, n)$ and outputting $\frac{k_1}{\varepsilon}$ (resp. $\frac{k_2}{\varepsilon}$) where $k_1 \in \mathbb{Z}$ (resp. $k_2 \in \mathbb{Z}$) is the unique integer such that $\text{REAL}(\mathbf{r}) \in \bar{I}_y(k_1)$ (resp. $\text{REAL}(\mathbf{r}') \in \bar{I}_{y-1}(k_2)$). \square

From the above construction, for all $k \in \mathbb{Z}$, we have

$$\frac{\Pr[\bar{M}_\varepsilon^{\text{CBCLCS}}(D_1, f) = \frac{k}{\varepsilon}]}{\Pr[\bar{M}_\varepsilon^{\text{CBCLCS}}(D_2, f) = \frac{k}{\varepsilon}]} = \frac{\Pr[\bar{Z}_y = \frac{k}{\varepsilon}]}{\Pr[\bar{Z}_{y-1} = \frac{k}{\varepsilon}]} = \frac{|\bar{I}_y(k)|}{|\bar{I}_{y-1}(k)|}.$$

Remark 4. It's easy to prove that $I_{y-1}(k) \cap I_y(k) \neq \emptyset$. The set of points $\{s_y(k)\}_{k \in \mathbb{Z}}$ partitions the interval $[0, 1]$ into infinitely many intervals $\{I_y(k) \stackrel{\text{def}}{=} [s_y(k-1), s_y(k)]\}_{k \in \mathbb{Z}}$. Similarly, the set of points $\{s_{y-1}(k)\}_{k \in \mathbb{Z}}$ partitions the interval $[0, 1]$ into infinitely many intervals $\{I_{y-1}(k) \stackrel{\text{def}}{=} [s_{y-1}(k-1), s_{y-1}(k)]\}_{k \in \mathbb{Z}}$.

Remark 5. Note that we can view $I_{y-1}(k)$ as having “shifted” $I_y(k)$ slightly to the right. Hence the truncation methods for the endpoints of $I_y(k)$ and $I_{y-1}(k)$ are different in order to guarantee BCL-complete sampling.

4.2 Concrete Results for Differential Privacy and Accuracy

In this section, we show that our construction satisfies compact $(\zeta, O(1))$ -BCL-consistent sampling and hence it's differentially private. Then, we give a relation between the result of [DY14] and ours.

Lemma 2 is one core step to achieve consistent sampling. Though it has essentially been proved by Dodis et al. [DLMV12], there still exist some typos there and the upper bound is not tight prior our work. Hence, we modify the Lemma A.1 of [DLMV12] and get Lemma 2. More concretely, recall that Lemma A.1. of [DLMV12] and partial proof are as follows.

Lemma 1. For all $y, k \in \mathbb{Z}$, $\frac{|I'_y(k)|}{|I_{y-1}(k)|} \leq 6\varepsilon$.

Proof.

...
 Case 3: If $s_y(k-1) < s_{y-1}(k-1) < \frac{1}{2} \leq s_{y-1}(k-1)$, then $\frac{|I'_y(k)|}{|I_{y-1}(k)|} \leq \frac{1-e^{-\varepsilon}}{2(e-1)}$.

...

□

It should be noted that: (1) It's obvious that “ $s_{y-1}(k-1) < \frac{1}{2} \leq s_{y-1}(k-1)$ ” never holds. (2) “ $\frac{|I'_y(k)|}{|I_{y-1}(k)|} \leq \frac{1-e^{-\varepsilon}}{2(e-1)}$ ” is wrong! Since $-1 - \frac{1}{\varepsilon} \leq v < -1$, without loss of generality, assume that $\frac{1}{\varepsilon}$ is an even integer and $v = -1 - \frac{1}{2\varepsilon}$. Then

$$\frac{|I'_y(k)|}{|I_{y-1}(k)|} = \frac{e^\varepsilon - 1}{2 \cdot e^{-\varepsilon v} - e^{-2\varepsilon v - \varepsilon - 1} - e^\varepsilon} = \frac{1 - e^{-\varepsilon}}{2(e^{\frac{1}{2}} - 1)} > \frac{1 - e^{-\varepsilon}}{2(e-1)},$$

which stands in contradiction to the inequality $\frac{|I'_y(k)|}{|I_{y-1}(k)|} \leq \frac{1-e^{-\varepsilon}}{2(e-1)}$.

By modifying Lemma 1, we get that

Lemma 2. Denote $I'_y(k) \stackrel{def}{=} I_y(k) \setminus I_{y-1}(k) = [s_y(k-1), s_{y-1}(k-1))$. For all $y, k \in \mathbb{Z}$ and $\varepsilon \in (0, 1)$, we have $|I'_y(k)|/|I_{y-1}(k)| < e \cdot \varepsilon$.

Proof. Note that if $x < y$, $\text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}(x) < \frac{1}{2}$; otherwise, $\text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}(x) \geq \frac{1}{2}$.

$$\frac{|I'_y(k)|}{|I_{y-1}(k)|} = \frac{s_{y-1}(k-1) - s_y(k-1)}{s_{y-1}(k) - s_{y-1}(k-1)} = \frac{\text{CDF}_{y-1, \frac{1}{\varepsilon}}^{\text{Lap}}(\frac{k-\frac{1}{2}}{\varepsilon}) - \text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}(\frac{k-\frac{1}{2}}{\varepsilon})}{\text{CDF}_{y-1, \frac{1}{\varepsilon}}^{\text{Lap}}(\frac{k+\frac{1}{2}}{\varepsilon}) - \text{CDF}_{y-1, \frac{1}{\varepsilon}}^{\text{Lap}}(\frac{k-\frac{1}{2}}{\varepsilon})}.$$

We consider four cases:

Case 1: If $\frac{1}{2} \leq s_y(k-1) < s_{y-1}(k-1) < s_{y-1}(k)$, then $\frac{|I'_y(k)|}{|I_{y-1}(k)|} = \frac{e^{\varepsilon+1} - e}{e-1}$.

Case 2: If $s_y(k-1) < \frac{1}{2} \leq s_{y-1}(k-1) < s_{y-1}(k)$, then

$$\frac{|I'_y(k)|}{|I_{y-1}(k)|} = \frac{1 - \frac{1}{2} \cdot e^{-\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon} - (y-1)]} - \frac{1}{2} \cdot e^{\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon} - y)}}{1 - \frac{1}{2} \cdot e^{-\varepsilon[\frac{k+\frac{1}{2}}{\varepsilon} - (y-1)]} - \{1 - \frac{1}{2} \cdot e^{-\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon} - (y-1)]}\}}.$$

For simplicity, denote $v \stackrel{def}{=} \frac{k-\frac{1}{2}}{\varepsilon} - y$. By the assumption, we have that $-1 \leq v < 0$. Correspondingly,

$$\begin{aligned} \frac{|I'_y(k)|}{|I_{y-1}(k)|} &= \frac{1 - \frac{1}{2}e^{-\varepsilon(v+1)} - \frac{1}{2}e^{\varepsilon v}}{-\frac{1}{2}e^{-\varepsilon(v+1+\frac{1}{\varepsilon})} + \frac{1}{2}e^{-\varepsilon(v+1)}} \\ &= \frac{-(e^{\varepsilon v} - 1)^2 - e^{-\varepsilon} + 1}{-e^{-1-\varepsilon} + e^{-\varepsilon}} \leq \frac{-e^{-\varepsilon} + 1}{-e^{-1-\varepsilon} + e^{-\varepsilon}} = \frac{e^{\varepsilon+1} - e}{e-1}. \end{aligned}$$

Case 3: If $s_y(k-1) < s_{y-1}(k-1) < \frac{1}{2} \leq s_{y-1}(k)$, then

$$\frac{|I'_y(k)|}{|I_{y-1}(k)|} = \frac{\frac{1}{2} \cdot e^{\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon}-(y-1)]} - \frac{1}{2} \cdot e^{\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon}-y)}}{1 - \frac{1}{2} \cdot e^{-\varepsilon[\frac{k+\frac{1}{2}}{\varepsilon}-(y-1)]} - \frac{1}{2} \cdot e^{\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon}-(y-1)]}}.$$

For simplicity, denote $v \stackrel{def}{=} \frac{k-\frac{1}{2}}{\varepsilon} - y$. By the assumption, we have that $-1 - \frac{1}{\varepsilon} \leq v < -1$. Correspondingly,

$$\begin{aligned} \frac{|I'_y(k)|}{|I_{y-1}(k)|} &= \frac{\frac{1}{2} \cdot e^{\varepsilon(v+1)} - \frac{1}{2} \cdot e^{\varepsilon v}}{1 - \frac{1}{2} \cdot e^{-\varepsilon(v+\frac{1}{\varepsilon}+1)} - \frac{1}{2} \cdot e^{\varepsilon(v+1)}} \\ &= \frac{e^{\varepsilon} - 1}{2 \cdot e^{-\varepsilon v} - e^{-2\varepsilon v - \varepsilon - 1} - e^{\varepsilon}} \\ &= \frac{e^{\varepsilon} - 1}{-(e^{-\varepsilon v - \frac{1+\varepsilon}{2}} - e^{\frac{1+\varepsilon}{2}})^2 + e^{1+\varepsilon} - e^{\varepsilon}} \\ &< \frac{e^{\varepsilon} - 1}{-(e^{\frac{\varepsilon-1}{2}} - e^{\frac{1+\varepsilon}{2}})^2 + e^{1+\varepsilon} - e^{\varepsilon}} \\ &= \frac{1 - e^{-\varepsilon}}{1 - e^{-1}}. \end{aligned}$$

Case 4: If $s_y(k-1) < s_{y-1}(k-1) < s_{y-1}(k) < \frac{1}{2}$, then

$$\frac{|I'_y(k)|}{|I_{y-1}(k)|} = \frac{\frac{1}{2} \cdot e^{\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon}-(y-1)]} - \frac{1}{2} \cdot e^{\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon}-y)}}{\frac{1}{2} \cdot e^{\varepsilon[\frac{k+\frac{1}{2}}{\varepsilon}-(y-1)]} - \frac{1}{2} \cdot e^{\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon}-(y-1)]}} = \frac{1 - e^{-\varepsilon}}{e - 1}.$$

For $\varepsilon \in (0, 1)$, we have

$$\frac{1 - e^{-\varepsilon}}{e - 1} < \frac{1 - e^{-\varepsilon}}{1 - e^{-1}} = \frac{e - e^{1-\varepsilon}}{e - 1} < \frac{e^{\varepsilon} \cdot (e - e^{1-\varepsilon})}{e - 1} = \frac{e^{\varepsilon+1} - e}{e - 1} < e \cdot \varepsilon.$$

The last inequality holds according to the following three facts: (1) $g_1(x) \stackrel{def}{=} \frac{e^{x+1} - e}{e - 1}$ is a convex function; (2) $g_2(x) \stackrel{def}{=} e \cdot x$ is a linear function; (3) $g_1(0) = g_2(0)$ and $g_1(1) = g_2(1)$. \square

Compared with the above lemma, ours is much better. In fact, our upper bound is tight.

Theorem 2. Mechanism $\overline{M}_{\varepsilon}^{CBCLCS}$ is a compact $((2^b + 1) \cdot e \cdot \varepsilon, \log(\frac{e \cdot (2^b + 1)}{1 - e^{-1}}))$ -BCL-consistent sampling mechanism for (δ, b) -BCL sources. Therefore, $\overline{M}_{\varepsilon}^{CBCLCS}$ is $(\mathcal{U}, 2e \cdot \varepsilon)$ -differentially private and $(\mathcal{BCL}(\delta, b), \xi)$ -differentially private for $\xi = (\frac{1+\delta}{1-\delta})^{\log(\frac{e \cdot (2^b + 1)}{1 - e^{-1}})} \cdot (\frac{1+\delta}{2})^{-b} \cdot (2^b + 1) \cdot e \cdot \varepsilon$.

Proof Sketch. Denote $I'_y(k) \stackrel{def}{=} I_y(k) \setminus I_{y-1}(k) = [s_y(k-1), s_{y-1}(k-1))$. Assume that Y is a distribution $\text{BCL}(\delta, b, n)$ and $S_0 \stackrel{def}{=} \{\mathbf{r} \in \{0, 1\}^n \mid \Pr[Y = \mathbf{r}] \neq 0\}$. Let $\text{STR}(I, n) \stackrel{def}{=} \{\mathbf{r} \in \{0, 1\}^n \mid \text{REAL}(\mathbf{r}) \in I\}$ as the the set of all n -bit strings whose real representation lies in I . Let $T_1 = \text{STR}(\bar{I}_y(k), n) \cap S_0$ and $T_2 = \text{STR}(\bar{I}_{y-1}(k), n) \cap S_0$. Then $T_1 \setminus T_2 = \text{STR}(\bar{I}'_y(k), n) \cap S_0$. Let $\tau(y-1, k/\varepsilon) \stackrel{def}{=} \log \frac{1}{|I_{y-1}(k)|} + \log(2^b + 1)$. Then we prove that for all $y, k \in \mathbb{Z}$, $|\text{STR}(\bar{I}'_y(k), \tau(y-1, k/\varepsilon)) \cap S_0| \leq (2^b + 1) \cdot e \cdot \varepsilon$ (see Corollary 1 below) and $|\text{SUFFIX}(\mathbf{u}, \tau(y-1, k/\varepsilon)) \cap S_0| \leq e \cdot (2^b + 1)/(1 - e^{-1})$ (see Corollary 2 below), where \mathbf{u} be the longest common prefix of all strings in $\bar{I} \stackrel{def}{=} \bar{I}_y(k) \cup \bar{I}_{y-1}(k)$. Therefore, by Definition 10 and Theorem 1, we obtain Theorem 2.

Proof.

Let $I''_y(k) \stackrel{def}{=} I_{y-1}(k) \setminus I_y(k) = [s_y(k), s_{y-1}(k))$. Similarly, we can get that that there exists a constant C such that $\frac{|I''_y(k)|}{|I_y(k)|} < C \cdot \varepsilon$ for $y, k \in \mathbb{Z}$ and $\varepsilon \in (0, 1)$.

Lemma 3. *For all $y, k \in \mathbb{Z}$, we have*

- (1) $|\bar{I}'_y(k)| \leq |I'_y(k)|$,
- (2) $|I_{y-1}(k)| + 2^{-\tau(y-1, k/\varepsilon)} \geq |\bar{I}_{y-1}(k)| \geq |I_{y-1}(k)| - 2^{-\tau(y-1, k/\varepsilon)}$,
- (3) $|I_y(k)| + 2^{-\tau(y-1, k/\varepsilon)} \geq |\bar{I}_y(k)| \geq |I_y(k)| - 2^{-\tau(y-1, k/\varepsilon)}$.

Proof.

(1) Since $s_{y-1}(k-1) \geq \bar{s}_{y-1}(k-1)$, and $\bar{s}_y(k-1) \geq s_y(k-1) - 2^{-\tau(y-1, k/\varepsilon)} + 2^{-\tau(y-1, k/\varepsilon)}$, we get $|\bar{I}'_y(k)| \leq |I'_y(k)|$.

(2) One one hand, since $\bar{s}_{y-1}(k) \geq s_{y-1}(k) - 2^{-\tau(y-1, k/\varepsilon)}$ and $\bar{s}_{y-1}(k-1) \leq s_{y-1}(k-1)$, we have $|\bar{I}_{y-1}(k)| \geq |I_{y-1}(k)| - 2^{-\tau(y-1, k/\varepsilon)}$. One the other hand, since $s_{y-1}(k) \geq \bar{s}_{y-1}(k)$ and $s_{y-1}(k-1) \leq \bar{s}_{y-1}(k-1) + 2^{-\tau(y-1, k/\varepsilon)}$, we have $|I_{y-1}(k)| + 2^{-\tau(y-1, k/\varepsilon)} \geq |\bar{I}_{y-1}(k)|$. Hence, Lemma 3 (2) holds.

(3) One one hand, since $\bar{s}_y(k) \geq s_y(k) - 2^{-\tau(y-1, k/\varepsilon)} + 2^{-\tau(y-1, k/\varepsilon)}$ and $\bar{s}_y(k-1) \leq s_y(k-1) + 2^{-\tau(y-1, k/\varepsilon)}$, we have $|\bar{I}_y(k)| \geq |I_y(k)| - 2^{-\tau(y-1, k/\varepsilon)}$. One the other hand, since $\bar{s}_y(k) \leq s_y(k) + 2^{-\tau(y-1, k/\varepsilon)}$ and $\bar{s}_y(k-1) \geq s_y(k-1)$, we have $|\bar{I}_y(k)| \leq |I_y(k)| + 2^{-\tau(y-1, k/\varepsilon)}$. Hence, Lemma 3 (3) holds. \square

Assume that Y is a distribution $\text{BCL}(\delta, b, n)$ and $S_0 \stackrel{def}{=} \{\mathbf{r} \in \{0, 1\}^n \mid \Pr[Y = \mathbf{r}] \neq 0\}$. Denote $\text{STR}(I, n) \stackrel{def}{=} \{\mathbf{r} \in \{0, 1\}^n \mid \text{REAL}(\mathbf{r}) \in I\}$ as the the set of all n -bit strings whose real representation lies in I . Let $T_1 = \text{STR}(\bar{I}_y(k), n) \cap S_0$ and $T_2 = \text{STR}(\bar{I}_{y-1}(k), n) \cap S_0$. Then $T_1 \setminus T_2 = \text{STR}(\bar{I}'_y(k), n) \cap S_0$. By induction, it can be easily seen that $2^{n-b} \leq |S_0| \leq 2^n$.

Lemma 4. *Let $\tau(y-1, k/\varepsilon) \stackrel{def}{=} \log \frac{1}{|I_{y-1}(k)|} + \log(2^b + 1)$. Then for all $y, k \in \mathbb{Z}$,*

- (1) $|\text{STR}(\bar{I}'_y(k), \tau(y-1, k/\varepsilon)) \cap S_0| \leq (2^{-b} + 1) \cdot e \cdot \varepsilon$,
- (2) $|\text{STR}(\bar{I}_{y-1}(k), \tau(y-1, k/\varepsilon)) \cap S_0| \geq 1$.

Proof.

(1) Let $n \stackrel{\text{def}}{=} \tau(y-1, k/\varepsilon)$ for shorthand. Consider $|\bar{I}'_y(k)|$ as the probability of sampling a sequence \mathbf{r} from U_{S_0} such that $\mathbf{r} \in \text{STR}(\bar{I}'_y(k), n) \cap S_0$, where $2^{n-b} \leq |S_0| \leq 2^n$. Hence, $|\bar{I}'_y(k)| = \sum_{\mathbf{r} \in \text{STR}(\bar{I}'_y(k), n) \cap S_0} \frac{1}{|S_0|} \geq \sum_{\mathbf{r} \in \text{STR}(\bar{I}'_y(k), n) \cap S_0} \frac{1}{2^n} = |\text{STR}(\bar{I}'_y(k), n) \cap S_0| \cdot \frac{1}{2^n}$. Therefore, by Lemma 3, we get $|\text{STR}(\bar{I}'_y(k), n) \cap S_0| \leq 2^n \cdot |\bar{I}'_y(k)| \leq 2^n \cdot |I'_y(k)| = (2^b + 1) \cdot \frac{|I'_y(k)|}{|I_{y-1}(k)|} \leq (2^b + 1) \cdot e \cdot \varepsilon$.

(2) Since $|\bar{I}_{y-1}(k)| = \sum_{\mathbf{r} \in \text{STR}(\bar{I}_{y-1}(k), n) \cap S_0} \frac{1}{|S_0|} \leq \sum_{\mathbf{r} \in \text{STR}(\bar{I}_{y-1}(k), n) \cap S_0} (\frac{1}{2})^{n-b} = |\text{STR}(\bar{I}_{y-1}(k), n) \cap S_0| \cdot (\frac{1}{2})^{n-b}$, by Lemma 3, we get

$$|\text{STR}(\bar{I}_{y-1}(k), n) \cap S_0| \geq 2^{n-b} \cdot |\bar{I}_{y-1}(k)| \geq 2^{n-b} \cdot (|I_{y-1}(k)| - 2^{-n}) = 1.$$

□

Remark 6. We can guarantee that n is legal, in the sense that the modification of the endpoints in $I_{y-1}(k)$ and $I_y(k)$ with respect to n does not cause intervals to “disappear” or for consecutive intervals to “overlap”.

From Lemma 4, we get that

Corollary 1. Denote $\tau(y-1, k/\varepsilon) \stackrel{\text{def}}{=} \log \frac{1}{|I_{y-1}(k)|} + \log(2^b + 1)$. For all $y, k \in \mathbb{Z}$, we have

$$\frac{|\text{STR}(\bar{I}'_y(k), \tau(y-1, k/\varepsilon)) \cap S_0|}{|\text{STR}(\bar{I}_{y-1}(k), \tau(y-1, k/\varepsilon)) \cap S_0|} \leq (2^b + 1) \cdot e \cdot \varepsilon.$$

Corollary 2. Denote $\tau(y-1, k/\varepsilon) \stackrel{\text{def}}{=} \log \frac{1}{|I_{y-1}(k)|} + \log(2^b + 1)$. Let \mathbf{u} be the longest common prefix of all strings in $\bar{I} \stackrel{\text{def}}{=} \bar{I}_y(k) \cup \bar{I}_{y-1}(k)$. Then

$$|\text{SUFFIX}(\mathbf{u}, \tau(y-1, k/\varepsilon)) \cap S_0| \leq \frac{e \cdot (2^b + 1)}{1 - e^{-1}}.$$

Proof. For simplicity, let $n \stackrel{\text{def}}{=} \tau(y-1, k/\varepsilon)$. Let \mathbf{u}' be the longest common prefix of all strings in $I \stackrel{\text{def}}{=} I_y(k) \cup I_{y-1}(k)$. Then we have $|\text{SUFFIX}(\mathbf{u}, n)| \leq |\text{SUFFIX}(\mathbf{u}', n)|$. We bound $|\text{SUFFIX}(\mathbf{u}, n)|$ by bounding the number of n -bit strings to the left or right of \bar{I} (depending on which endpoint of the interval $[0, 1]$ is closer to I).

Now we calculate the size of the interval $[s_y(k-1), 1]$ (resp. $[0, s_{y-1}(k)]$), which is an approximation of the size of $[\bar{s}_y(k-1), 1]$ (resp. $[0, \bar{s}_{y-1}(k)]$). Then we can upper bound how many n -bit strings there are in the interval $[\bar{s}_y(k-1), 1]$ (resp. $[0, \bar{s}_{y-1}(k)]$). Let $S \stackrel{\text{def}}{=} [s_y(k-1), 1]$.

Recall that $s_y(k) \stackrel{def}{=} \text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}(\frac{k+\frac{1}{2}}{\varepsilon})$ for all $k \in \mathbb{Z}$ and

$$\text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}(x) = \begin{cases} \frac{1}{2} \cdot e^{\varepsilon(x-y)}, & \text{if } x < y; \\ 1 - \frac{1}{2} \cdot e^{-\varepsilon(x-y)}, & \text{if } x \geq y. \end{cases}$$

Note that if $x < y$, $\text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}(x) < \frac{1}{2}$; otherwise, $\text{CDF}_{y, \frac{1}{\varepsilon}}^{\text{Lap}}(x) \geq \frac{1}{2}$.

$I'_y(k) = [s_y(k-1), s_{y-1}(k-1)]$ and $I'_{y+1}(k) = [s_{y+1}(k-1), s_y(k-1)]$.

For simplicity, denote $v \stackrel{def}{=} \frac{k-\frac{1}{2}}{\varepsilon} - y$. We consider four cases.

Case 1: Assume that $\frac{1}{2} \leq s_{y+1}(k-1) < s_y(k-1) < s_{y-1}(k-1)$. Then $v \geq 1$.

$$\frac{|I'_y(k)|}{|I'_{y+1}(k)|} = \frac{1 - \frac{1}{2} \cdot e^{-\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon} - (y-1)]} - 1 + \frac{1}{2} \cdot e^{-\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon} - y)}}{1 - \frac{1}{2} \cdot e^{-\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon} - y)} - 1 + \frac{1}{2} \cdot e^{-\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon} - (y+1)]}} = \frac{1}{e^\varepsilon}.$$

Case 2: Assume that $s_{y+1}(k-1) < \frac{1}{2} \leq s_y(k-1) < s_{y-1}(k-1)$. Then $0 \leq v < 1$.

$$\frac{|I'_y(k)|}{|I'_{y+1}(k)|} = \frac{e^{-\varepsilon v} - e^{-\varepsilon(v+1)}}{2 - e^{-\varepsilon v} - e^{\varepsilon(v-1)}} = \frac{1 - e^{-\varepsilon}}{-e^{-\varepsilon}(e^{\varepsilon v} - e^\varepsilon)^2 + e^\varepsilon - 1}.$$

Hence,

$$\frac{1}{e^\varepsilon} < \frac{|I'_y(k)|}{|I'_{y+1}(k)|} \leq 1.$$

Case 3: Assume that $s_{y+1}(k-1) < s_y(k-1) < \frac{1}{2} \leq s_{y-1}(k-1)$. Then $-1 \leq v < 0$.

$$\begin{aligned} \frac{|I'_y(k)|}{|I'_{y+1}(k)|} &= \frac{1 - \frac{1}{2} \cdot e^{-\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon} - (y-1)]} - \frac{1}{2} \cdot e^{\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon} - y)}}{\frac{1}{2} \cdot e^{\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon} - y)} - \frac{1}{2} \cdot e^{\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon} - (y+1)]}} \\ &= \frac{1 - \frac{1}{2} \cdot e^{-\varepsilon(v+1)} - \frac{1}{2} \cdot e^{\varepsilon v}}{\frac{1}{2} \cdot e^{\varepsilon v} - \frac{1}{2} \cdot e^{\varepsilon(v-1)}} \\ &= \frac{-(e^{-\varepsilon v - \frac{\varepsilon}{2}} - e^{\frac{\varepsilon}{2}})^2 + e^\varepsilon - 1}{1 - e^{-\varepsilon}}. \end{aligned}$$

Therefore,

$$1 < \frac{|I'_y(k)|}{|I'_{y+1}(k)|} \leq e^\varepsilon.$$

Case 4: Assume that $s_{y+1}(k-1) < s_y(k-1) < s_{y-1}(k-1) < \frac{1}{2}$. Then $v < -1$.

$$\frac{|I'_y(k)|}{|I'_{y+1}(k)|} = \frac{\frac{1}{2} \cdot e^{\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon}-(y-1)]} - \frac{1}{2} \cdot e^{\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon}-y)}}{\frac{1}{2} \cdot e^{\varepsilon(\frac{k-\frac{1}{2}}{\varepsilon}-y)} - \frac{1}{2} \cdot e^{\varepsilon[\frac{k-\frac{1}{2}}{\varepsilon}-(y+1)]}} = \frac{\frac{1}{2} \cdot e^{\varepsilon(v+1)} - \frac{1}{2} \cdot e^{\varepsilon v}}{\frac{1}{2} \cdot e^{\varepsilon v} - \frac{1}{2} \cdot e^{\varepsilon(v-1)}} = e^{\varepsilon}.$$

We only analyze **Case 1**, the other cases are analogous.

Since $I'_y(k)$ and $I'_{y+1}(k)$ are consecutive intervals for all $y \in \mathbb{Z}$, we have

$$|S| = \sum_{j=-\infty}^y |I'_j(k)| \leq \sum_{j=-\infty}^y |I'_y(k)|(e^{-\varepsilon})^{y-j} = \frac{|I'_y(k)|}{1 - e^{-\varepsilon}} \leq \frac{|I'_y(k)|}{(1 - \frac{1}{e}) \cdot \varepsilon}.$$

The last inequality holds from the facts: (1) $g_1(x) \stackrel{def}{=} 1 - e^{-x}$ is a concave function; (2) $g_2(x) \stackrel{def}{=} (1 - \frac{1}{e}) \cdot x$ is a linear function; (3) $g_1(0) = g_2(0)$ and $g_1(1) = g_2(1)$.

Let $\bar{S} \stackrel{def}{=} [\bar{s}_y(k-1), 1]$. Then $|\bar{S}| \leq |S| \leq \frac{|I'_y(k)|}{(1 - \frac{1}{e}) \cdot \varepsilon}$.

On the other hand, $|\bar{S}|$ can be considered as the probability of sampling a sequence \mathbf{r} from the uniform distribution U_{S_0} such that $\mathbf{r} \in \text{STR}(\bar{S}, n) \cap S_0$ and $2^{n-b} \leq |S_0| \leq 2^n$. Therefore,

$$|\bar{S}| = \sum_{\mathbf{r} \in \text{STR}(\bar{S}, n) \cap S_0} \frac{1}{|S_0|} \geq \sum_{\mathbf{r} \in \text{STR}(\bar{S}, n) \cap S_0} \left(\frac{1}{2}\right)^n = |\text{STR}(\bar{S}, n) \cap S_0| \cdot \left(\frac{1}{2}\right)^n.$$

Correspondingly,

$$|\text{STR}(\bar{S}, n) \cap S_0| \leq 2^n \cdot |\bar{S}| \leq 2^n \cdot \frac{|I'_y(k)|}{(1 - \frac{1}{e}) \cdot \varepsilon} = \frac{|I'_y(k)|}{|I'_{y-1}(k)|} \cdot \frac{(2^b + 1)}{(1 - \frac{1}{e}) \cdot \varepsilon} \leq \frac{e \cdot (2^b + 1)}{1 - e^{-1}}.$$

Hence,

$$|\text{SUFFIX}(\mathbf{u}, n) \cap S_0| \leq |\text{STR}(\bar{S}, n) \cap S_0| \leq \frac{e \cdot (2^b + 1)}{1 - e^{-1}}. \quad \square$$

Combining Theorem 1, Corollary 1, and Corollary 2, we get Theorem 2. \square

Theorem 3. $\overline{M}_\varepsilon^{\text{CBCLCS}}$ has $(\mathcal{BCL}(\delta, b), O(\frac{1}{\varepsilon} \cdot \frac{1}{1-\delta}))$ -utility and $(\mathcal{U}, O(\frac{1}{\varepsilon}))$ -utility.

Proof. We only need to prove that for all neighboring databases $D_1, D_2 \in \mathcal{D}$, all $f \in \mathcal{F}$, and all $BCL(\delta, b) \in \mathcal{BCL}(\delta, b)$, $\mathbb{E}_{\mathbf{r} \leftarrow BCL(\delta, b)}[|\overline{M}_\varepsilon^{\text{CBCLCS}}(D_1, f; \mathbf{r}) - f(D_1)|]$ and $\mathbb{E}_{\mathbf{r} \leftarrow BCL(\delta, b)}[|\overline{M}_\varepsilon^{\text{CBCLCS}}(D_2, f; \mathbf{r}) - f(D_2)|]$ are both upper bounded by $O(\frac{1}{\varepsilon} \cdot \frac{1}{1-\delta})$. Without loss of generality, assume that $f(D_1) = y$ and $f(D_2) = y-1$. Then $\mathbb{E}_{\mathbf{r} \leftarrow BCL(\delta, b)}[|\overline{M}_\varepsilon^{\text{CBCLCS}}(D_1, f; \mathbf{r}) - y|] = \sum_{k=-\infty}^{\infty} \Pr_{\mathbf{r} \leftarrow BCL(\delta, b)}[\overline{M}_\varepsilon^{\text{CBCLCS}}(D_1, f; \mathbf{r}) = \frac{k}{\varepsilon}] \cdot |\frac{k}{\varepsilon} - y|.$

Let \mathbf{a} be the longest common prefix of all strings in $\text{STR}(\bar{I}_y(k), \tau(y-1, k/\varepsilon))$. Denote $I_0 \stackrel{\text{def}}{=} \text{SUFFIX}(\mathbf{a}0, \tau(y-1, k/\varepsilon)) \cap \text{STR}(\bar{I}_y(k), \tau(y-1, k/\varepsilon))$ and $I_1 \stackrel{\text{def}}{=} \text{SUFFIX}(\mathbf{a}1, \tau(y-1, k/\varepsilon)) \cap \text{STR}(\bar{I}_y(k), \tau(y-1, k/\varepsilon))$. Thus, $I_0 \cup I_1 = \text{STR}(\bar{I}_y(k), \tau(y-1, k/\varepsilon))$. Correspondingly, we have

$$\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b)} [\overline{M}_\varepsilon^{\text{CBCLCS}}(D_1, f; \mathbf{r}) = \frac{k}{\varepsilon}] \leq \left(\frac{1+\delta}{2}\right)^{|\mathbf{a}0|} + \left(\frac{1+\delta}{2}\right)^{|\mathbf{a}1|} \leq 2 \cdot \left(\frac{1+\delta}{2}\right)^{\log\left(\frac{1}{|I_y(k)|}\right)}.$$

Similarly, we can conclude that

$$\Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b)} [\overline{M}_\varepsilon^{\text{CBCLCS}}(D_2, f; \mathbf{r}) = \frac{k}{\varepsilon}] \leq 2 \cdot \left(\frac{1+\delta}{2}\right)^{\log\left(\frac{1}{|I_{y-1}(k)|}\right)}.$$

Claim. For all $y, k \in \mathbb{Z}$, we have $|I_y(k)| \leq \frac{1}{2} \cdot e^{-\frac{1}{2}} \cdot (e-1) \cdot e^{-|k-\varepsilon y|}$.

Proof. We consider three cases.

Case 1: Assume that $\frac{k-\frac{1}{2}}{\varepsilon} - y \geq 0$ and $\frac{k+\frac{1}{2}}{\varepsilon} - y \geq 0$. Then

$$|I_y(k)| = 1 - \frac{1}{2} \cdot e^{-\varepsilon\left(\frac{k+\frac{1}{2}}{\varepsilon} - y\right)} - \left[1 - \frac{1}{2} \cdot e^{-\varepsilon\left(\frac{k-\frac{1}{2}}{\varepsilon} - y\right)}\right] = \frac{1}{2} \cdot e^{-\frac{1}{2}} \cdot (e-1) \cdot e^{-|k-\varepsilon y|}.$$

Case 2: Assume that $\frac{k-\frac{1}{2}}{\varepsilon} - y < 0$ and $\frac{k+\frac{1}{2}}{\varepsilon} - y \geq 0$. From the fact that $1 - \frac{1}{2}x \leq \frac{1}{2} \cdot \frac{1}{x}$ for all $x > 0$, we obtain

$$|I_y(k)| = 1 - \frac{1}{2} \cdot e^{-\varepsilon\left(\frac{k+\frac{1}{2}}{\varepsilon} - y\right)} - \frac{1}{2} \cdot e^{\varepsilon\left(\frac{k-\frac{1}{2}}{\varepsilon} - y\right)} \leq \frac{1}{2} \cdot e^{-\frac{1}{2}} \cdot (e-1) \cdot e^{-|k-\varepsilon y|}.$$

Case 3: Assume that $\frac{k-\frac{1}{2}}{\varepsilon} - y < 0$ and $\frac{k+\frac{1}{2}}{\varepsilon} - y < 0$. Then $|I_y(k)| = \frac{1}{2} \cdot e^{-\frac{1}{2}} \cdot (e-1) \cdot e^{-|k-\varepsilon y|}$.

□

By Lemma 3, $|\bar{I}_y(k)| \leq |I_y(k)| + 2^{-\tau(k-1, y)} = |I_y(k)| + \frac{1}{2^{b+1}} |I_{y-1}(k)|$. Hence,

$$\begin{aligned} \log\left(\frac{1}{|\bar{I}_y(k)|}\right) &\geq \log\frac{1}{\frac{1}{2}e^{-\frac{1}{2}}(e-1)\left(1 + \frac{1}{2^{b+1}}\right)} + \log(e^{\min\{|k-\varepsilon y|, |k-\varepsilon y+\varepsilon|\}}) \\ &\geq \min\{|k-\varepsilon y|, |k-\varepsilon y+\varepsilon|\} \geq |k-\varepsilon y| - 1. \end{aligned}$$

Similarly, $\log\left(\frac{1}{|\bar{I}_{y-1}(k)|}\right) \geq |k-\varepsilon y| - 1$. Therefore,

$$\begin{aligned}
& \sum_{k=-\infty}^{\infty} \Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b)} [\overline{M}_\varepsilon^{\text{CBCLCS}}(D_1, f; \mathbf{r}) = \frac{k}{\varepsilon}] \cdot \left| \frac{k}{\varepsilon} - y \right| \\
& \leq \sum_{k=-\infty}^0 2 \cdot \left(\frac{1+\delta}{2} \right)^{|\varepsilon y - k| - 1} \cdot \left| y - \frac{k}{\varepsilon} \right| + \sum_{k=1}^{\infty} 2 \cdot \left(\frac{1+\delta}{2} \right)^{|k - \varepsilon y| - 1} \cdot \left| \frac{k}{\varepsilon} - y \right| \\
& \leq \frac{2}{\varepsilon} \cdot \left(\frac{1+\delta}{2} \right)^{-1} \cdot \left[\sum_{k=1}^{\infty} \left(\frac{1+\delta}{2} \right)^{k-1} \cdot k + \sum_{k=-\infty}^0 \left(\frac{1+\delta}{2} \right)^{-k} \cdot (-k+1) \right] \\
& = \left(\frac{1+\delta}{2} \right)^{-1} \cdot \frac{4}{\varepsilon} \cdot \frac{1}{1 - \left(\frac{1+\delta}{2} \right)^2} = O\left(\frac{1}{\varepsilon} \cdot \frac{1}{1-\delta} \right).
\end{aligned}$$

Similarly, we get that

$$\sum_{k=-\infty}^{\infty} \Pr_{\mathbf{r} \leftarrow \text{BCL}(\delta, b)} [\overline{M}_\varepsilon^{\text{CBCLCS}}(D_2, f; \mathbf{r}) = \frac{k}{\varepsilon}] \cdot \left| \frac{k}{\varepsilon} - (y-1) \right| \leq O\left(\frac{1}{\varepsilon} \cdot \frac{1}{1-\delta} \right).$$

When $\delta = 0$ and $b = 0$, the BCL source degenerates into the uniform source.

Therefore, the mechanism $\overline{M}_\varepsilon^{\text{CBCLCS}}$ has $(\text{BCL}(\delta, b), O(\frac{1}{\varepsilon} \cdot \frac{1}{1-\delta}))$ -utility and $(\mathcal{U}, O(\frac{1}{\varepsilon}))$ -utility. \square

From Theorems 2 and 3, we get that

Theorem 4. *There exists an explicit $(\text{BCL}(\delta, b), \xi)$ -differentially private and (\mathcal{U}, ρ) -accurate mechanism M for the Hamming weight queries where*

$$\rho = \frac{2^{b \cdot \log(1+\delta) - 9}}{\xi} \cdot \left(\frac{2}{1+\delta} \right)^{b+1} \cdot \frac{2^b + 1}{(1+\delta)^b} \cdot \left(\frac{1+\delta}{1-\delta} \right)^{\log \frac{(2^b+1)e}{1-e^{-1}}} \cdot \frac{2^{11}}{1 - \left(\frac{1+\delta}{2} \right)^2} \cdot e > \frac{2^{b \cdot \log(1+\delta) - 9}}{\xi}.$$

One the other hand, recall that Dodis and Yao [DY14] obtained that

Theorem 5. *If $b \geq \frac{\log(\xi\rho) + 9}{\log(1+\delta)} = \Omega\left(\frac{\log(\xi\rho) + 1}{\delta}\right)$, then no $(\text{BCL}(\delta, b), \xi)$ -differentially private and (\mathcal{U}, ρ) -accurate mechanism for the Hamming weight queries exists.*

Therefore, we conclude that

Corollary 3. *Assume that the mechanism M is $(\text{BCL}(\delta, b), \xi)$ -differentially private and (\mathcal{U}, ρ) -accurate for the Hamming weight queries, then $\rho > \frac{2^{b \cdot \log(1+\delta) - 9}}{\xi}$.*

Corollary 3 implies that it's possible to construct a $(\text{BCL}(\delta, b), \xi)$ -differentially private and (\mathcal{U}, ρ) -accurate mechanism for Hamming weight queries, where $\rho > \frac{2^{b \cdot \log(1+\delta) - 9}}{\xi}$. In this paper, we show an explicit construction of such mechanisms.

Remark 7. If we replace the truncation method in [DLMV12] with the one in Step 2 of this paper, we can prove that the modified mechanism in [DLMV12] satisfies the compact $(\zeta', O(1))$ -SV-consistent sampling. Therefore, the resulting mechanism is differentially private. We can also prove that it's accurate. The proofs are similar to Theorems 2 and 3.

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