# On Basing Search SIVP on NP-Hardness 

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#### Abstract

The possibility of basing cryptography on the minimal assumption NP $\nsubseteq \mathbf{B P P}$ is at the very heart of complexity-theoretic cryptography. The closest we have gotten so far is lattice-based cryptography whose average-case security is based on the worst-case hardness of approximate shortest vector problems on integer lattices. The state-of-the-art is the construction of a one-way function (and collision-resistant hash function) based on the hardness of the $\tilde{O}(n)$-approximate shortest independent vector problem $\operatorname{SIVP}_{\tilde{O}(n)}$.

Although SIVP is NP-hard in its exact version, Guruswami et al (CCC 2004) showed that gapSIVP $\sqrt{n / \log n}$ is in NP $\cap \operatorname{coAM}$ and thus unlikely to be NP-hard. Indeed, any language that can be reduced to $\operatorname{gapSIVP}_{\tilde{O}(\sqrt{n})}$ (under general probabilistic polynomial-time adaptive reductions) is in $\mathbf{A M} \cap$ coAM by the results of Peikert and Vaikuntanathan (CRYPTO 2008) and Mahmoody and Xiao (CCC 2010). However, none of these results apply to reductions to search problems, still leaving open a ray of hope: can NP be reduced to solving search SIVP with approximation factor $\tilde{O}(n)$ ?

We show that any language that can be reduced to solving search SIVP with approximation factor $\tilde{O}(n)$ lies in AM intersect coAM, eliminating the possibility of basing current constructions on NP-hardness.


## 1 Introduction

It is a long standing open question whether cryptography can be based on the minimal assumption that NP $\nsubseteq \mathbf{B P P}$. More precisely, one would hope to construct cryptographic primitives such that given an polynomial-time algorithm breaking the security of the primitive, one can efficiently solve SAT.

The closest we have gotten so far is lattice cryptography. This approach was born out of the breakthrough result of Ajtai [Ajt96] which constructs a one-way function family based on the worst case hardness of certain lattice problems such as the $\gamma$-approximate shortest independent vectors problem (SIVP ${ }_{\gamma}$ ), which can be stated as follows: given an $n$-dimensional lattice, find a set of $n$ linearly independent vectors whose length ${ }^{1}$ is at most $\gamma(n)$ (polynomial in $n$ ) times the length of the shortest such vector set. Since the work of Ajtai, the state of the art is a construction of a family of collision resistant hash functions (CRHF) based on the hardness of the shortest independent vectors problem with an approximation factor $\tilde{O}(n)$ [MR04]. One would hope that this approach is viable for constructing cryptography based on NP-hardness since Khot showed that SIVP $\gamma_{\gamma}$ is NP-hard for any constant factor [Kho05]. Presumably, if one could construct cryptographic primitives based on the hardness of $\operatorname{SIVP}_{O(1)}$, we would be golden. Alternatively, if one could extend the result of

[^0]Khot to show the NP-hardness of SIVP $\gamma_{\gamma}$ for larger $\gamma(n)$, we would be closer to the goal of basing cryptography on NP-hardness.

However, there are some negative results when one considers the corresponding gap version of the same lattice problem. The gap problem, denoted by gapSIVP $\gamma_{\gamma}$, is to estimate the length of the short independent vector set within a factor of $\gamma(n)$. Peikert and Vaikuntanathan show that $\operatorname{gapSIVP}_{\omega(\sqrt{n \log n})}$ is in SZK [PV08]. Thus there is no Cook reduction from SAT to gapSIVP $\tilde{O}_{(\sqrt{n})}$ unless the polynomial hierarchy collapses (as $\mathbf{B P P}{ }^{\text {SZK }} \subseteq \mathbf{A M} \cap \operatorname{coAM}$ [MX10]).

Fortunately, the hardness of SIVP is not contradicted by the fact that the gap problem with the same approximation factor is easy. For instance, if one considers any ideal lattice in the field $\mathbb{Z}[x] /\left\langle x^{2^{k}}+1\right\rangle$, its successive minima satisfy $\lambda_{1}=\ldots=\lambda_{n}$, thus gapSIVP ${ }_{\sqrt{n}}$ can be trivially solved using Minkowski's inequality. However, finding a set of short independent vectors in such ideal lattices is still not known to be easy. As none of these negative results apply to reductions to search SIVP, there is still a ray of hope: can NP be reduced to solving search SIVP with approximation $\tilde{O}(n)$ ?

Thus, in order to really understand the viability of the approach begun by the work of Ajtai, it seems one must study the search versions of lattice problems. In this work, we relate the hardness of the search version SIVP $_{\gamma}$, to the gap version gapSIVP. Informally, we show that if gapSIVP $\gamma_{\gamma}$ is not hard, neither is SIVP $_{\sqrt{n} \cdot \gamma}$.

Main Theorem. If gapSIVP ${ }_{\gamma} \in \mathbf{S Z K}$ and there exists a probabilistic polynomial-time adaptive reduction from a language L to $\operatorname{SIVP}_{\tilde{O}(\sqrt{n} \cdot \gamma)}$, then $\mathrm{L} \in \mathbf{A M} \cap \operatorname{coAM}$.

As a quick corollary from combining our result with gapSIVP $\tilde{O}_{(\sqrt{n})} \in \mathbf{S Z K}$ [PV08], any language that can be reduced to SIVP $_{\tilde{O}(n)}$ lies in $\mathbf{A M}$ intersect coAM and thus it is not NP-hard unless the polynomial hierarchy collapses. This eliminates the possibility of basing current constructions on NP-hardness.

Corollary. If there exists a probabilistic polynomial-time adaptive reduction from a language $\mathbf{L}$ to $\operatorname{SIVP}_{\tilde{O}(n)}$, then $\mathrm{L} \in \mathbf{A M} \cap \mathbf{c o A M}$.

### 1.1 Proof Overview

The first step is to shift from a search problem to a sampling problem. Our goal is to obtain a black-box separation between SIVP $_{\gamma}$ and NP-hardness by showing that any language $L$ that can be reduced to SIVP $_{\gamma}$ is in AM intersect coAM. Let $\mathcal{R}$ be the reduction from $L$ to SIVP ${ }_{\gamma}$. We will construct an $\mathbf{A M}$ protocol for L using reduction $\mathcal{R}$. For a first attempt, the naïve verifier samples a random tape and sends it to the prover. The prover simulates the reduction $\mathcal{R}$ and resolves any query to SIVP $_{\gamma}$ using its unbounded computational power. The simulation, including the answers to the reduction's query to $\mathrm{SIVP}_{\gamma}$, is sent to the naïve verifier, so that the verifier can check its correctness. But SIVP $_{\gamma}$ is a search problem and there is no unique right answer. The prover has the freedom to decide which answer is chosen upon each query. This freedom allows a malicious prover to fool the naïve verifier. Similar difficulty were faced by Bogdanov and Brzuska, which is resolved by inherently shifting to sampling problems. In order to separate size-verifiable one-way functions from NP-Hardness [BB15], they force the prover to sample a random answer uniformly among all correct ones. Thus the correct answer distribution for each query is unique.

Inspired by the work of Bogdanov and Brzuska, we consider a sampling problem correlated to SIVP $_{\gamma}$, called the discrete Gaussian distribution. A discrete Gaussian over a lattice is a distribution such that the probability of any vertex $\mathbf{v}$ is proportional to $e^{-\pi\|\mathbf{v}-\mathbf{c}\|^{2} / s^{2}}$, where $\mathbf{c}$ is its "center" and parameter $s$ is its "width". Lemma 4.2 shows that discrete Gaussian sampling is as hard as SIVP $\gamma_{\gamma}$
in the sense that there is a black-box reduction from SIVP $_{\gamma}$ to discrete Gaussian sampling with "width" $\gamma(n) / \sqrt{n}$. Therefore, if language L can be reduced to SIVP $_{\gamma}$, then it can also be reduced to discrete Gaussian sampling on lattices with "width" $s \leq \lambda_{n} / \sqrt{n}$.
Lemma 4.2 (Informal). For $\gamma \geq \sqrt{n}$, SIVP $_{\gamma}$ can be reduced to discrete Gaussian sampling on lattices with "width" $\sigma=\frac{\gamma}{\sqrt{n}} \lambda_{n}$.

The proof of Lemma 4.2 is quite intuitive. Repeatedly sample from the discrete Gaussian over the same lattice centered at $\mathbf{0}$. With good probability, the newly sampled vertex is short and is linearly independent from previously sampled verteces.

The next natural question is, which property separates a sampling problem from $\mathbf{N P}$-hardness? Here we introduce the notion of "probability-verifiability". Informally, a distribution family is probability-verifiable if for any distribution $\mathcal{D}$ in this family and for any possible value $v, \operatorname{Pr}[v \leftarrow \mathcal{D}]$, the probability that $v$ is sampled from $\mathcal{D}$, can be lower bounded within an arbitrarily good precision in AM.
Lemma 4.3 (Informal). If a language $L$ can be reduced to a probability-verifiable sampling problem S , then $\mathrm{L} \in \mathbf{A M} \cap \mathbf{c o A M}$.

Lemma 4.3 is a generalization of [BB15]. Assume language L can be reduced to sampling problem $S$. The input of $S$ is interpreted as the description of a distribution, let $\mathcal{P}_{\mathrm{pd}}$ denotes the distribution specified by input pd.

Let $\mathcal{R}$ be the reduction from L to sampling problem S . On each input $x$, an execution $\mathcal{R}^{\mathrm{S}}(x)$ is determined by the random tape of reduction $\mathcal{R}$, denoted by $r$, and the answers to the reduction's queries to S . The transcript is defined as $\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \ldots, \mathrm{pd}_{T}, v_{T}\right)$ where $\mathrm{pd}_{t}$ is the $t$-th query to S and $v_{t}$ is the corresponding response. Note that $r, v_{1}, \ldots, v_{T}$ determine the execution, since $\mathrm{pd}_{t}$ is determined by $r, v_{1}, \ldots, v_{t-1}$. Then

$$
\begin{equation*}
\operatorname{Pr}\left[\mathcal{R}^{\mathrm{S}}(x) \text { accepts }\right]=\sum_{\substack{\sigma: \text { accepting transcript } \\ \text { of } \mathcal{R}^{\mathrm{S}}(x)}} \operatorname{Pr}[\sigma]=\sum_{\substack{\sigma \text { :accepting } \\ \text { of } \mathcal{R}^{\mathrm{S}}(x)}} \operatorname{Pr}[r] \cdot \mathcal{P}_{\mathrm{pd}_{1}}\left(v_{1}\right) \cdot \ldots \cdot \mathcal{P}_{\mathrm{pd}_{T}}\left(v_{T}\right) . \tag{1}
\end{equation*}
$$

For simplicity, assume for now that there is an efficient algorithm that computes the probability $\mathcal{P}_{\mathrm{pd}}(v)$ given pd and value $v$. This property is stronger than probability-verifiability. Then the probability that $\mathcal{R}^{\mathrm{S}}(x)$ accepts, which equals a sum (equation (1)) where each term can be efficient computed, can be lower bounded using the set lower bound protocol of Goldwasser and Sipser [GS86], so $L \in$ AM. Symmetrically, $L \in$ coAM. The proof of Lemma 4.3 shows the same result from the weaker condition that $S$ is probability-verifiable.

There is one last step missing between Lemma 4.2 and Lemma 4.3: Is discrete Gaussian sampling probability-verifiable? What is the smallest factor $\gamma$ such that discrete Gaussian sampling with "width" $\leq \gamma \lambda_{n}$ is probability-verifiable? Lemma 4.4 answers this question, and it connects the hardness of discrete Gaussian sampling with the hardness of gapSIVP.
Lemma 4.4 (Informal). Assume gapSIVP ${ }_{\gamma}$ is in SZK. There exists a real valued function $s(\mathbf{B}) \in$ $\left[\lambda_{n}, \tilde{O}(\gamma) \cdot \lambda_{n}\right]$ such that given a lattice basis $\mathbf{B}$, discrete Gaussian sampling over lattice $\mathcal{L}(\mathbf{B})$ with "width" $s(\mathbf{B})$ is probability-verifiable.

Lemma 4.4 has an easier proof assuming the stronger condition that gapSIVP ${ }_{\gamma}$ is in $\mathbf{P}$. If there were some deterministic polynomial time algorithm solving gapSIVP ${ }_{\gamma}$, there would exist $s(\mathbf{B}) \in$ $\left[\lambda_{n}(\mathbf{B}), \gamma \lambda_{n}(\mathbf{B})\right]$ that can be efficiently computed by binary search. As $s(\mathbf{B}) \geq \lambda_{n}(\mathbf{B})$, the verifier can ask the prover to provide a set of $n$ linearly independent vectors $\mathbf{w}_{1}, \ldots, \mathbf{w}_{n}$ whose length is no longer than $s(\mathbf{B})$. Given the lattice basis $\mathbf{B}$ and a set of short linearly independent vectors, there exists an efficient algorithm that samples from the discrete Gaussian with the desired parameter
$\left[\mathrm{BLP}^{+} 13\right]$. When the verifier can sample from a distribution, he can lower bound the probability of each value using the set lower bound protocol [GS86].

This informal proof assumes gapSIVP ${ }_{\gamma} \in \mathbf{P}$ in order to compute a function $s(\mathbf{B})$ that $s(\mathbf{B}) \approx$ $\lambda_{n}(\mathbf{B})$. This assumption can be weakened to gapSIVP $\gamma_{\gamma} \in \mathbf{S Z K}$, as the verifier only needs to compute such a function $s(\mathbf{B})$ in an AM protocol.

Corollary (of Lemma 3.1). Assume gapSIVP ${ }_{\gamma}$ is in SZK. There exists a real valued function $s(\mathbf{B}) \in\left[\lambda_{n}, \tilde{O}(\gamma) \cdot \lambda_{n}\right]$ that can be efficiently computed in $\mathbf{A M}$.

### 1.2 Related Works

Prior work exploring the problem of basing cryptography on worst-case NP-hardness has obtained several negative results for black-box reduction. Brassard [Bra79] first showed that one-way permutations cannot be based on NP-hardness. Goldreich and Goldwasser [GG98] showed that public-key encryption schemes satisfying certain very specific properties cannot be based on NP-hardness. The required properties include the ability to certify an invalid key.

Work of Akavia, Goldreich, Goldwasser and Moshkovitz [AGGM06] and Bogdanov and Brzuska [BB15] showed that a special class of one-way functions called size-verifiable one-way functions cannot be based on NP-hardness. A size-verifiable one-way function is one in which the size of the set of pre-images can be efficiently approximated via an AM protocol.

Bogdanov and Lee [BL13] showed that homomorphic encryption schemes satisfying a special property cannot be based on NP-hardness. The required property is that the homomorphic evaluation produces a ciphertext whose distribution is statistically close to that of a fresh encrypted ciphertext.

Recently, Liu and Vaikuntanathan [LV16] showed that single-server private information retrieval (PIR) schemes cannot be based on NP-hardness.

Several works have also obtained a separation results for restricted types of reductions, most notably non-adaptive reductions which make all oracle queries simultaneously. The work of Feigenbaum and Fortnow [FF91], subsequently strengthened by Bogdanov and Trevisan [BT06], showed that there cannot be a non-adaptive reduction from SAT to the average-case hardness of any problem in NP, unless the polynomial hierarchy collapses.

On basing lattice problems on NP-hardness, the work of Goldreich and Goldwasser [GG00], subsequently strengthened by Micciancio and Vadhan [MV03], showed that gapSVP $\sqrt{n / \log n}$ and $\operatorname{gapCVP}_{\sqrt{n / \log n}}$ are both contained in NP $\cap \mathbf{S Z K}$. The shortest vector problem (SVP) and the closest vector problem (CVP), roughly speaking, is the problem of finding the shortest non-zero vector in a lattice or finding the lattice vector that is closest to a given point. The corresponding gap problem gapSVP ${ }_{\gamma}, \operatorname{gapCVP}_{\gamma}$ is to estimate within a factor of $\gamma(n)$ the length of the shortest non-zero vector or the distance to the closest lattice vector from a given point. The problem gapSVP is connected to gapSIVP via so-called "transference theorems" for lattices [Ban93]. Aharonov and Regev [AR04] explored a slightly looser approximation factor and showed that gapSVP ${ }_{\sqrt{n}}$ and gapCVP $\sqrt{n}$ are both contained in NP $\cap$ coNP.

In prior work on the gap version of the SIVP problem, Guruswami, Micciancio and Regev [GMR04] showed that gapSIVP $\sqrt{n / \log n} \in \mathbf{N P} \cap \mathbf{c o A M}$. Peikert and Vaikuntanathan [PV08] showed that gapSIVP ${ }_{\gamma} \in \mathbf{S Z K}$ for any $\gamma=\omega(\sqrt{n \log n})$. In contrast to these results for promise problems, our work explores the approximate SIVP problem. With an approximation factor $\gamma(n)=\tilde{O}(n)$, this search problem is the basis of lattice-based collision resistant hash function (CRHF) constructions [Ajt96, MR04]. We separate SIVP $_{\gamma}$ from NP-hardness for the same approximation factor.

## 2 Preliminaries

Lattice A lattice in $\mathbb{R}^{n}$ is an additive subgroup of $\mathbb{R}^{n}$

$$
\left\{\sum_{i=1}^{n} x_{i} \mathbf{b}_{i}: x_{i} \in \mathbb{Z} \text { for } 1 \leq i \leq n\right\}
$$

generated by $n$ linearly independent vectors $\mathbf{b}_{1}, \ldots, \mathbf{b}_{n} \in \mathbb{R}^{n}$. The set of vectors $\mathbf{b}_{1}, \ldots, \mathbf{b}_{n}$ is called a basis for the lattice. A basis can be represented by matrix $\mathbf{B} \in \mathbb{R}^{n \times n}$ whose columns are the basis vectors. The lattice generated by the columns of $\mathbf{B}$ is denoted by $\mathcal{L}(\mathbf{B})$.

$$
\mathcal{L}(\mathbf{B})=\left\{\mathbf{B} \mathbf{x}: \mathbf{x} \in \mathbb{N}^{n}\right\} .
$$

The $i$-th successive minimum of a lattice $\mathcal{L}$, denoted by $\lambda_{i}(\mathcal{L})$, is defined as the minimum length that $\mathcal{L}$ contains $i$ linearly independent vectors of length at most $\lambda_{i}(\mathcal{L})$. Formally,

$$
\lambda_{i}(\mathcal{L}):=\min \{r: \operatorname{dim}(\mathcal{L} \cap r \mathcal{B}) \geq i\}
$$

where $r \mathcal{B}$ is the radius $r$ ball centered at the origin defined as $r \mathcal{B}:=\left\{\mathbf{x} \in \mathbb{R}^{n}:\|\mathbf{x}\|_{2} \leq r\right\}$. We abuse notations and write $\lambda_{i}(\mathbf{B})$ instead of $\lambda_{i}(\mathcal{L}(\mathbf{B}))$.

Shortest Independent Vectors Problem (SIVP) SIVP is a computational problem. Given a basis $\mathbf{B}$ of an $n$-dimensional lattice, find a set of $n$ linearly independent vectors $\mathbf{v}_{1}, \ldots, \mathbf{v}_{n} \in \mathcal{L}(\mathbf{B})$ such that $\max _{i}\left\|\mathbf{v}_{i}\right\|$ is minimized, i.e., $\left\|\mathbf{v}_{i}\right\| \leq \lambda_{n}(\mathbf{B})$ for all $1 \leq i \leq n$.

SIVP $_{\gamma}$ is the approximation version of SIVP with factor $\lambda$. Given a basis $\mathbf{B}$ of an $n$-dimensional lattice, find a set of $n$ linearly independent vectors $\mathbf{v}_{1}, \ldots, \mathbf{v}_{n} \in \mathcal{L}(\mathbf{B})$ such that $\left\|\mathbf{v}_{i}\right\| \leq \gamma(n) \cdot \lambda_{n}(\mathbf{B})$ for all $1 \leq i \leq n$. The approximation factor $\gamma$ is typical a polynomial in $n$.
gapSIVP $\gamma_{\gamma}$ is the decision version of SIVP $_{\gamma}$. An input to gapSIVP ${ }_{\gamma}$ is a basis $\mathbf{B}$ of a $n$-dimensional lattice and a scalar $s$. It is a YES instance if $\lambda_{n}(\mathbf{B}) \leq s$, and is a NO instance if $\lambda_{n}(\mathbf{B}) \geq \lambda(n) \cdot s$.

Discrete Gaussian For any vector $\mathbf{c}$ and any $s>0$, let

$$
\rho_{\mathbf{c}, s}(\mathbf{v})=e^{-\pi\|\mathbf{v}-\mathbf{c}\|_{2}^{2} / s^{2}}
$$

be a Gaussian function with mean $\mathbf{c}$ and width $s$. Functions are extends to sets in usual way, $\rho_{\mathbf{c}, s}(\mathcal{L})=\sum_{\mathbf{v} \in \mathcal{L}} \rho_{\mathbf{c}, s}(\mathbf{v})$. The discrete Gaussian distribution over lattice $\mathcal{L}$ with mean $\mathbf{c}$ and width $s$, denoted by $\mathcal{N}_{\mathcal{L}, \mathbf{c}, s}$, is defined by

$$
\forall \mathbf{v} \in \mathcal{L}, \mathcal{N}_{\mathcal{L}, \mathbf{c}, s}(\mathbf{v})=\frac{\rho_{\mathbf{c}, s}(\mathbf{v})}{\rho_{\mathbf{c}, s}(\mathcal{L})}
$$

In this work, most discrete Gaussian distributions considered are centered at the origin. Let $\rho_{s}, \mathcal{N}_{\mathcal{L}, s}$ denote $\rho_{0, s}, \mathcal{N}_{\mathcal{L}, \mathbf{0}, s}$ respectively.

Sampling Problems Besides computational problems and decision problems, we define sampling problems. The input of a sampling problem specifies a distribution, let $\mathcal{P}_{\text {pd }}$ denotes the distribution specified by input pd. The goal is to sample from the distribution $\mathcal{P}_{\mathrm{pd}}$. A probabilistic polynomialtime algorithm $\mathcal{S}$ perfectly solves the sampling problem if for any input pd

$$
\forall v, \operatorname{Pr}[\mathcal{S}(\mathrm{pd}) \rightarrow v]=\mathcal{P}_{\mathrm{pd}}(v) .
$$

The probability is over the random input tape of $\mathcal{S}$. In a more practical definition, $\mathcal{S}$ solves the sampling problem if the output distribution of $\mathcal{S}(\mathrm{pd})$ is close to $\mathcal{P}_{\text {pd }}$, i.e.

$$
\Delta_{\mathrm{sd}}\left(\mathcal{S}\left(\mathrm{pd}, 1^{\ell}\right), \mathcal{P}_{\mathrm{pd}}\right) \leq \frac{1}{\ell}
$$

where $\Delta_{\text {sd }}$ denotes the statistical distance.
For example, in this work, discrete Gaussian is considered as a sampling problem. For any function $s(\cdot)$ mapping lattice bases to positive real numbers, define sampling problem $\mathrm{DGS}_{s}$. The input of $\mathrm{DGS}_{s}$ is a lattice basis $\mathbf{B}$. The target output distribution $\mathcal{P}_{\mathbf{B}}$ is the discrete Gaussian distribution $\mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}$, where each vector $v \in \mathcal{L}(\mathbf{B})$ is sampled with probability

$$
\mathcal{P}_{\mathbf{B}}(\mathbf{v})=\mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}(\mathbf{v})=\frac{\rho_{s(\mathbf{B})}(\mathbf{v})}{\rho_{s(\mathbf{B})}(\mathcal{L}(\mathbf{B}))} .
$$

Probability-Verifiable A sampling problem is probability-verifiable if there exists an AM protocol to lower bound $\mathcal{P}_{\mathrm{pd}}(v)$ for any pd and $v$. More precisely, there exists a family of error function $\left\{\eta_{\mathrm{pd}, m}\right\}$ such that for any pd, $m$, the error function $\eta_{\mathrm{pd}, m}:\{0,1\}^{*} \rightarrow[0,+\infty)$ satisfies $\sum_{v} \eta_{\mathrm{pd}, m}(v) \leq \frac{1}{m}$, and the promise problem

- YES instance: $\left(\mathrm{pd}, v, \hat{p}, 1^{m}\right)$ such that $\hat{p}=\mathcal{P}_{\mathrm{pd}}(v)$
- NO instance: $\left(\mathrm{pd}, v, \hat{p}, 1^{m}\right)$ such that $\hat{p} \geq \mathcal{P}_{\mathrm{pd}}(v)+\eta_{\mathrm{pd}, m}(v)$
is in AM.
Sampling Oracles In order to formalize the (probabilistic) Turing reduction to a sampling problem, we also define sampling oracles, which is a generalization of traditional oracles studied by complexity theorists. Let $\mathcal{S}$ be a sampling oracle for a fixed sampling problem. $\mathcal{S}$ can be queried on any valid pd; upon query pd, sampling oracle $\mathcal{S}(\mathrm{pd})$ would always output a fresh sample from distribution $\mathcal{P}_{\mathrm{pd}}$. E.g. if the sampling oracle $\mathcal{S}$ is queried for the same pd multiple times, it would outputs i.i.d. samples from distribution $\mathcal{P}_{\text {pd }}$.

A probabilistic Turing reduction from a language $L$ to a sampling problem $S$ is a probabilistic poly-time oracle Turing machine $\mathcal{R}$, such that $\mathcal{R}$ can solve $L$ given a sampling oracle that samples from $S$ in the sense that

$$
\begin{aligned}
& x \in \mathrm{~L} \Longrightarrow \mathcal{R}^{\mathcal{S}}(x) \rightarrow 1 \text { w.p. } \geq 2 / 3 \\
& x \notin \mathrm{~L} \Longrightarrow \mathcal{R}^{\mathcal{S}}(x) \rightarrow 1 \text { w.p. } \leq 1 / 3
\end{aligned}
$$

If such a reduction exists, we say $L$ can be reduced to sampling problem $S$, denoted by $L \in \mathbf{B P P}^{S}$.
Similarly, a computational problem or a search problem can be reduced to a sampling problem S if they can be efficiently solved given the sampling oracle of $S$.
$\mathbb{R}$-TFAM and $\mathbb{R}_{\eta}$-TFAM The complexity class $\mathbb{R}$-TFAM is introduced by Mahmoody and Xiao [MX10]. Informally, it's consist of real-valued functions that can be efficiently computed in AM. A function $f:\{0,1\}^{*} \rightarrow \mathbb{R}$ is in $\mathbb{R}$-TFAM if the following promise problem is in AM:

- YES instance: $\left(x, f(x), 1^{m}\right)$.
- NO instance: $\left(x, y, 1^{m}\right)$ such that $|y-f(x)|>\frac{1}{m}$.

The definition of $\mathbb{R}$-TFAM emphasize on the absolute error. The complexity class $\mathbb{R}_{\eta}$-TFAM is defined to capture those functions that can be efficiently computed in AM with small relative error. A function $g:\{0,1\}^{*} \rightarrow \mathbb{R}^{+}$is in $\mathbb{R}_{\eta}$-TFAM if the following promise problem is in AM:

- YES instance: $\left(x, g(x), 1^{m}\right)$.
- NO instance: $\left(x, y, 1^{m}\right)$ such that $|y-g(x)|>\frac{1}{m} \cdot g(x)$.

It follows directly from the definitions that $g \in \mathbb{R}_{\eta}$-TFAM if and only if $\log g \in \mathbb{R}$-TFAM for any function $g:\{0,1\}^{*} \rightarrow \mathbb{R}^{+}$.

Statistical Zero Knowledge Statistical zero knowledge (SZK) is the class of decision problems that can be verified by a statistical zero-knowledge proof protocol. Entropy Difference (ED) is a complete problem for SZK [GV99], which is defined as the following: Given two polynomial-size circuits, $C$ and $D$, let $\mathcal{C}$ and $\mathcal{D}$ be the distributions of their respective outputs when $C, D$ are fed with uniform random inputs. The problem is to distinguish between

- YES instance: $(C, D)$ such that $H(\mathcal{C})-H(\mathcal{D}) \geq 1$;
- NO instance: $(C, D)$ such that $H(\mathcal{C})-H(\mathcal{D}) \leq-1$.

Where $H$ is the Shannon entropy. Moreover, the mapping $H: C \mapsto H(\mathcal{C})$ is in $\mathbb{R}$-TFAM.

## 3 Gap Problems

The lattice problem gapSIVP is essentially estimating $\lambda_{n}(\mathbf{B})$ given a lattice basis $\mathbf{B}$. This definition can be generalized to any real valued functions. Define the gap problem of function $f:\{0,1\}^{*} \rightarrow \mathbb{R}^{+}$ with gap $\gamma:\{0,1\}^{*} \rightarrow[1,+\infty)$, denoted by gap $f_{\gamma}$, as the promise problem

- YES instance: $(x, y)$ such that $y \leq f(x)$;
- NO instance: $(x, y)$ such that $y>\gamma(x) \cdot f(x)$.

In this work, estimating $\lambda_{n}(\mathbf{B})$ is of critical importance. Its gap problem, gapSIVP ${ }_{\gamma}$, alone is not sufficient for the proof. Instead, a stronger form of approximation is defined. Say $g:\{0,1\}^{*} \rightarrow \mathbb{R}^{+}$ is an approximation of function $f$ within factor $\gamma$ if $f(x) \leq g(x) \leq \lambda(x) \cdot f(x)$ for all $x$. Clearly, computing $g$ is a harder problem than gap $f_{\gamma}$, in the sense that there is a trivial reduction from gap $f_{\gamma}$ to computing $g$.

The following Lemma shows a reduction in the other direction: if gap $f_{\gamma}$ is in SZK, then there exists an approximation of $f$ within almost the same factor, which can be computed in AM.

Lemma 3.1. For any real valued function $f:\{0,1\}^{*} \rightarrow \mathbb{R}^{+}$and any gap $\gamma:\{0,1\}^{*} \rightarrow[1,+\infty)$ that $\log \gamma(x) \leq \operatorname{poly}(|x|)$, if gap $f_{\gamma} \in \mathbf{S Z K}$, then for any constant $\mu>1$, there exists $g:\{0,1\}^{*} \rightarrow \mathbb{R}^{+}$ such that $\forall x, g(x) \in[f(x), \mu \gamma(x) f(x)]$ and $g$ is in $\mathbb{R}_{\eta^{-}}$TFAM.

Lemma 3.1 can be combined with previous results about gapSIVP. Guruswami et al. [GMR04] showed that gapSIVP $\sqrt{n / \log n} \in \operatorname{coAM}$. Peikert and Vaikuntanathan [PV08] showed that gapSIVP ${ }_{\gamma} \in$ NISZK $\subseteq$ SZK for any $\gamma=\omega(\sqrt{n \log n})$. Thus there exists an approximation of $\lambda_{n}$ within a factor $\tilde{O}(\sqrt{n})$ that can be computed in AM.

Corollary. For any $\gamma(n)=\omega(\sqrt{n \log n})$, there exists a function $g$ maps lattice bases to real numbers such that $g \in \mathbb{R}_{\eta^{\prime}}$ TFAM and $\lambda_{n}(\mathbf{B}) \leq g(\mathbf{B})<\gamma(n) \cdot \lambda_{n}(\mathbf{B})$.

Proof of Lemma 3.1. Entropy Difference (ED) is a complete problem for $\mathbf{S Z K}$, so gap $f_{\gamma} \in \mathbf{S Z K}$ implies the existance of a reduction $(x, y) \mapsto\left(C_{x, y}, D_{x, y}\right)$ that maps input $x$ together with a real number $y$ to random circuits $C_{x, y}, D_{x, y}$. Let $\mathcal{C}_{x, y}$ and $\mathcal{D}_{x, y}$ be the output distributions of $C_{x, y}, D_{x, y}$. The reduction from gap $f_{\gamma}$ to ED satisfies the following properties:

- There is an efficient deterministic algorithm computing $C_{x, y}, D_{x, y}$ given input $(x, y)$.
- $H\left(\mathcal{C}_{x, y}\right)-H\left(\mathcal{D}_{x, y}\right)>2$ for any $x, y$ that $y \leq f(x)$.
- $H\left(\mathcal{C}_{x, y}\right)-H\left(\mathcal{D}_{x, y}\right)<-1$ for any $x, y$ that $y>\gamma(x) \cdot f(x)$.

Define the clamp function

$$
\operatorname{clamp}(y):= \begin{cases}1, & \text { if } y \geq 1 \\ y, & \text { if } y \in(0,1) \\ 0, & \text { if } y \leq 0\end{cases}
$$

For any fixed constant $\mu>1$, define
$g(x)=\exp \left(\ln \mu \cdot \sum_{i=0}^{+\infty} \operatorname{clamp}\left(H\left(\mathcal{C}_{x, \mu^{i}}\right)-H\left(\mathcal{D}_{x, \mu^{i}}\right)\right)+\ln \mu \cdot \sum_{i=1}^{+\infty}\left(\operatorname{clamp}\left(H\left(\mathcal{C}_{x, \mu^{-i}}\right)-H\left(\mathcal{D}_{x, \mu^{-i}}\right)\right)-1\right)\right)$.
As $\operatorname{clamp}\left(H\left(\mathcal{C}_{x, y}\right)-H\left(\mathcal{D}_{x, y}\right)\right)=1$ for $y \leq f(x)$,

$$
g(x) \geq \exp \left(\ln \mu \cdot\left\lceil\log _{\mu}(f(x))\right\rceil\right) \geq f(x) .
$$

As clamp $\left(H\left(\mathcal{C}_{x, y}\right)-H\left(\mathcal{D}_{x, y}\right)\right)=0$ for $y>\gamma(x) \cdot f(x)$,

$$
g(x) \leq \exp \left(\ln \mu \cdot\left\lceil\log _{\mu}(\gamma(x) \cdot f(x))\right\rceil\right) \leq \mu \gamma(x) \cdot f(x)
$$

In order to complete the proof, we show that $g$ is in $\mathbb{R}_{\eta}$-TFAM. For any input $x, \hat{g}$, the prover can prove $\hat{g} \approx g(x)$ if $\hat{g}=g(x)$.

Consider the following protocol, $\varepsilon=1 / \operatorname{poly}(m, \ln \gamma)$ will be fixed later.
AM "protocol" on input $(x, \hat{g})$
P: Send $\ldots, \hat{d}_{-1}, \hat{d}_{0}, \hat{d}_{1}, \hat{d}_{2}, \ldots$ such that $\log _{\mu} \hat{g}=\sum_{i=0}^{\infty} \operatorname{clamp}\left(\hat{d}_{i}\right)+\sum_{i=1}^{\infty}\left(\operatorname{clamp}\left(\hat{d}_{-i}\right)-1\right)$
$\mathbf{P}, \mathbf{V}:$ For each $i \in \mathbb{Z}$, convince the verifier that $\hat{d}_{i}-\varepsilon<H\left(\mathcal{C}_{x, \mu^{i}}\right)-H\left(\mathcal{D}_{x, \mu^{i}}\right)<\hat{d}_{i}+\varepsilon$

On any input $x$, define $d_{i}=H\left(\mathcal{C}_{x, \mu^{i}}\right)-H\left(\mathcal{D}_{x, \mu^{i}}\right)$. And the honest prover should send $\hat{d}_{i}=d_{i}$. The prover have to prove that $d_{i}-\varepsilon<\hat{d}_{i}<d_{i}+\varepsilon$. For $\mu^{i} \leq f(x), \hat{d}_{i} \geq d_{i}-\varepsilon \geq 1$, then $\operatorname{clamp}\left(\hat{d}_{i}\right)=1=\operatorname{clamp}\left(d_{i}\right)$. For $\mu^{i} \geq \mu \gamma(x) f(x), \hat{d}_{i} \leq d_{i}+\varepsilon \leq 0$, then $\operatorname{clamp}\left(\hat{d}_{i}\right)=0=\operatorname{clamp}\left(d_{i}\right)$. For $f(x)<\mu^{i}<\mu \gamma(x) f(x),\left|\operatorname{clamp}\left(\hat{d}_{i}\right)-\operatorname{clamp}\left(d_{i}\right)\right| \leq\left|\hat{d}_{i}-d_{i}\right|<\varepsilon$.

Thus

$$
\begin{aligned}
\left|\frac{\ln \hat{g}-\ln g(x)}{\ln \mu}\right| & \leq \sum_{i \in \mathbb{Z}}\left|\operatorname{clamp}\left(\hat{d}_{i}\right)-\operatorname{clamp}\left(d_{i}\right)\right| \\
& =\sum_{f(x)<\mu^{i}<\mu \gamma(x) f(x)}\left|\operatorname{clamp}\left(\hat{d}_{i}\right)-\operatorname{clamp}\left(d_{i}\right)\right| \\
& <\left\lceil\log _{\mu}(\mu \gamma(x))\right\rceil \varepsilon \\
& <\frac{\ln \gamma(x)+2}{\ln \mu} \varepsilon .
\end{aligned}
$$

If $\varepsilon$ is sufficiently small, $\hat{g}$ would be close to $g(x)$. To ensure $|\hat{g}-g(x)| \leq \frac{1}{m} g(x)$, it is sufficient to set $\varepsilon=O\left(\frac{1}{m(\ln \gamma(x)+2)}\right)$.

The above "protocol" is not a real protocol, as it requires the prover to send an infinite sequence to the verifier. To compress the proof, the prover need a succinct interactive proof that $d_{j}>1$ for all $j \leq i_{L}$ and $d_{j}<0$ for all $j \geq i_{H}$.

For an index $i$, if the prover can convince the verifier that $d_{i}=H\left(\mathcal{C}_{x, \mu^{i}}\right)-H\left(\mathcal{D}_{x, \mu^{i}}\right)<2$, the verifier also learns that $\mu^{i}>g(x)$, thus for any $j \geq i+\left\lceil\log _{\mu} \gamma(x)\right\rceil, \mu^{j}>\gamma(x) g(x)$ and $d_{j} \leq-1$. Similarly, if the prover can convince the verifier that $d_{i}=H\left(\mathcal{C}_{x, \mu^{i}}\right)-H\left(\mathcal{D}_{x, \mu^{i}}\right)>-1$, the verifier also knows that $d_{j} \geq 2$ for any $j \leq i-\left\lceil\log _{\mu} \gamma(x)\right\rceil$.

Thus the real AM protocol that proves $\hat{g} \in\left(g(x)-\frac{1}{m}, g(x)+\frac{1}{m}\right)$ is the following:
AM protocol on input $\left(x, \hat{g}, 1^{m}\right)$
P: Send $\hat{d}_{i_{L}}, \hat{d}_{i_{L}+1}, \ldots, \hat{d}_{i_{H}-1}, \hat{d}_{i_{H}}$ such that

- $\log _{\mu} \hat{g}=i_{L}+\sum_{i=i_{L}}^{i_{H}} \operatorname{clamp}\left(\hat{d}_{i}\right)$
- $i_{H}=i_{L}+2\left\lceil\log _{\mu} \gamma(x)\right\rceil$
- $\left.\hat{d}_{i_{L}+\left\lceil\log _{\mu}\right.} \gamma(x)\right\rceil>0$
- $\left.\hat{d}_{i_{L}+\left\lceil\log _{\mu}\right.} \gamma(x)\right\rceil+1<1$
$\mathbf{P}, \mathbf{V}:$ For each $i \in \mathbb{Z}$, convince the verifier that $\hat{d}_{i}-\varepsilon<H\left(\mathcal{C}_{x, \mu^{i}}\right)-H\left(\mathcal{D}_{x, \mu^{i}}\right)<\hat{d}_{i}+\varepsilon$ for $\varepsilon=O\left(\frac{1}{m(\ln \gamma(x)+2)}\right)$.


## 4 Search SIVP and NP-hardness

Theorem 4.1. For any factor $\gamma: \mathbb{N} \rightarrow \mathbb{R}$, if gapSIVP ${ }_{\gamma} \in \mathbf{S Z K}$ and there exists a probabilistic polynomial-time adaptive reduction from a language L to $\operatorname{SIVP}_{\sqrt{\pi n \log (2 n+4)} \cdot \gamma}$, then $\mathrm{L} \in \mathbf{A M} \cap$ coAM.

The smallest factor $\gamma$ we knows that makes problem gapSIVP ${ }_{\gamma}$ be in SZK if from [PV08]: for any factor $\gamma(n)=\omega(\sqrt{n \log n})$, problem gapSIVP ${ }_{\gamma}$ is in SZK.

Corollary. For any factor $\gamma(n)=\omega(n \log n)$, if there exists a probabilistic polynomial-time adaptive reduction from a language L to $\mathrm{SIVP}_{\gamma}$, then $\mathrm{L} \in \mathbf{A M} \cap$ coAM.

The proof of theorem 4.1 is the combination of Lemma 4.2, Lemma 4.3 and Lemma 4.4. Problem $\operatorname{gapSIVP}_{\gamma}$ is in $\mathbf{S Z K}^{\text {and }}$ there is a reduction from language L to search problem SIVP $\sqrt{\pi n \log (2 n+4)} \cdot \gamma$. Lemma 4.2 shows that there is another reduction from L to sampling problem $\mathrm{DGS}_{s}$ for any $s$ satisfying

$$
\begin{equation*}
s(\mathbf{B}) \in\left[2 \lambda_{n}(\mathbf{B}), \sqrt{\pi \log (2 n+4)} \cdot \gamma \lambda_{n}(\mathbf{B})\right] . \tag{2}
\end{equation*}
$$

Lemma 4.4 shows that there exists a function $s$ satisfying (2) such that the sampling problem $\mathrm{DGS}_{s}$ is probability-verifiable. Therefore, there exists a reduction from $L$ to a probability-verifiable sampling problem. Finally Lemma 4.3 shows that such a language $L$ must live in $\mathbf{A M} \cap$ coAM.

Lemma 4.2. Let $s(\cdot)$ be a function mapping lattice bases to real numbers, such that $\forall \mathbf{B}, 2 \lambda_{n}(\mathbf{B}) \leq$ $s(\mathbf{B}) \leq \frac{\gamma}{\sqrt{n}} \lambda_{n}(\mathbf{B})$. Then there exists a probabilistic Turing reduction from $\mathrm{SIVP}_{\gamma}$ to $\mathrm{DGS}_{s}$.

Lemma 4.3. If there exists a probabilistic Turing reduction from a promise problem $L=\left(L_{Y}, L_{N}\right)$ to probability-verifiable sampling problems, then $\mathbf{L} \in \mathbf{A M} \cap \mathbf{c o A M}$.

Lemma 4.4. For any factor $\gamma: \mathbb{N} \rightarrow \mathbb{R}$, if $\operatorname{gapSIVP}_{\gamma(n) / \sqrt{\pi \log (2 n+4)}} \in \mathbf{S Z K}$, then there exists $a$ function $s(\cdot)$ mapping lattice bases to real numbers, such that $\forall \mathbf{B}, s(\mathbf{B}) \in\left[2 \lambda_{n}(\mathbf{B}), \gamma(n) \cdot \lambda_{n}(\mathbf{B})\right]$ and $\mathrm{DGS}_{s}$ is probability-verifiable.

### 4.1 From Search SIVP to Discrete Gaussian Sampling

This section proves Lemma 4.2. The reduction from SIVP $_{\gamma}$ to discrete Gaussian sampling is straightforward: If we can sample from discrete Gaussian distribution of width $s \in\left[2 \cdot \lambda_{n}, \frac{\gamma}{\sqrt{n}} \lambda_{n}\right]$, keep sampling from it until $n$ short, linearly independent vectors are sampled.

When sampling from discrete Gaussian $\mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}$, there are two bad events that might occurs:

- The sampled vector is too long, its Euclidean norm is larger than $\gamma \lambda_{n}(\mathbf{B})$.
- The sampled vector is not linearly independent from previous ones, it lies in the subspace spanned by previously chosen vectors.

In order to prove Lemma 4.2, it's sufficient to show that there is a constant probability that none of these bad events occurs.

Lemma 4.5 (Lemma 1.5 in [Ban93]). For any $c>1 / \sqrt{2 \pi}$, $n$-dimensional lattice $\mathcal{L}$

$$
\begin{equation*}
\rho_{s}(\mathcal{L} \backslash c s \sqrt{n} \mathcal{B})<C^{n} \cdot \rho_{s}(\mathcal{L}) \tag{3}
\end{equation*}
$$

where $C=c \sqrt{2 \pi e} \cdot e^{-\pi c^{2}}$.
Lemma 4.5 bounds the probability that an overlong vector is sampled from a discrete Gaussian distribution. Let the constant $c$ in formula (3) equals 1 ,

$$
\operatorname{Pr}_{\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}}[\|\mathbf{v}\|>\sqrt{n} \cdot s(\mathbf{B})]=\frac{\rho_{s}(\mathcal{L}(\mathbf{B}) \backslash s \sqrt{n} \mathcal{B})}{\rho_{s}(\mathcal{L}(\mathbf{B}))}<\sqrt{2 \pi e} \cdot e^{-\pi}<0.2
$$

As $\gamma(n) \cdot \lambda_{n}(\mathbf{B}) \geq \sqrt{n} \cdot s(\mathbf{B})$,

$$
\operatorname{Pr}_{\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}}\left[\|\mathbf{v}\|>\gamma \lambda_{n}(\mathbf{B})\right] \leq \operatorname{Pr}_{\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}}[\|\mathbf{v}\|>\sqrt{n} \cdot s(\mathbf{B})]<0.2 .
$$

For any proper linear subspace $\mathcal{V} \subsetneq \mathbb{R}^{n}$, we would bound the probability $\operatorname{Pr}_{\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}}[\mathbf{v} \in \mathcal{V}]$. By the definition of successive minimum, there exists $\mathbf{u} \in \mathcal{L} \backslash \mathcal{V}$ such that $\|\mathbf{u}\| \leq \lambda_{n}(\mathbf{B})$. Let $\mathcal{L}^{\prime}$ denotes $\mathcal{L} \cap \mathcal{V}$. As $\mathcal{L}$ is close under addition, $\mathcal{L}^{\prime}+\mathbf{u}, \mathcal{L}^{\prime}-\mathbf{u}$ are subsets of $\mathcal{L}$. Moreover, as $\mathcal{V}$ is close under addition and $\mathbf{u} \notin \mathcal{V}, \mathcal{L}^{\prime}+\mathbf{u}, \mathcal{L}^{\prime}, \mathcal{L}^{\prime}-\mathbf{u}$ are disjointed.

$$
\begin{aligned}
& \operatorname{Pr}_{\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}}[\mathbf{v} \in \mathcal{V}]=\frac{\rho_{s}\left(\mathcal{L}^{\prime}\right)}{\rho_{s}(\mathcal{L})} \\
& \leq \frac{\rho_{s}\left(\mathcal{L}^{\prime}\right)}{\rho_{s}\left(\mathcal{L}^{\prime}-\mathbf{u}\right)+\rho_{s}\left(\mathcal{L}^{\prime}\right)+\rho_{s}\left(\mathcal{L}^{\prime}+\mathbf{u}\right)}=\frac{\sum_{\mathbf{v} \in \mathcal{L}^{\prime}} \rho_{s}(\mathbf{v})}{\sum_{\mathbf{v} \in \mathcal{L}^{\prime}}\left(\rho_{s}(\mathbf{v}-\mathbf{u})+\rho_{s}(\mathbf{v})+\rho_{s}(\mathbf{v}+\mathbf{u})\right)}
\end{aligned}
$$

For any vectors $\mathbf{u}, \mathbf{v}$ such that $\|\mathbf{u}\| \leq \lambda_{n}(\mathbf{B}) \leq s / 2$,

$$
\begin{aligned}
& \rho_{s}(\mathbf{v}-\mathbf{u})+\rho_{s}(\mathbf{v}+\mathbf{u})=e^{-\pi\|\mathbf{v}-\mathbf{u}\|^{2} / s^{2}}+e^{-\pi\|\mathbf{v}-\mathbf{u}\|^{2} / s^{2}} \\
&=\left(e^{-2 \pi\langle\mathbf{u}, \mathbf{v}\rangle / s^{2}}+e^{2 \pi\langle\mathbf{u}, \mathbf{v}\rangle / s^{2}}\right) e^{-\pi\|\mathbf{u}\|^{2} / s^{2}} e^{-\pi\|\mathbf{v}\|^{2} / s^{2}} \leq 2 e^{-\pi / 2^{2}} \rho_{s}(\mathbf{v})
\end{aligned}
$$

Thus

$$
\operatorname{Pr}_{\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}}[\mathbf{v} \in \mathcal{V}] \leq \frac{\sum_{\mathbf{v} \in \mathcal{L}^{\prime}} \rho_{s}(\mathbf{v})}{\sum_{\mathbf{v} \in \mathcal{L}^{\prime}}\left(1+2 e^{-\pi / 2^{2}}\right) \rho_{s}(\mathbf{v})}=\frac{1}{1+2 e^{-\pi / 2^{2}}}<0.6
$$

By union bound, the probability of bad event is at most 0.8 . By Chernoff bound, if sample $m$ times from the discrete Gaussian distribution (for $m \geq 10 n$ ), the probability that the sampled vectors contains $n$ linearly independent vectors of length at most $\gamma \lambda_{n}(\mathbf{B})$ is at least $1-e^{-\frac{m}{10}}$.

### 4.2 Probability-Verifiable Sampling Problem and NP-hardness

This section proves Lemma 4.3, which is a generalization of [BB15], the proof techniques are similar.
Let $\mathcal{M}$ be the reduction from a promise problem $\mathrm{L}=\left(\mathrm{L}_{Y}, \mathrm{~L}_{N}\right)$ to $\mathcal{S}$. For a given input $x$, we want to distinguish between $\operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right] \geq 8 / 9$ and $\operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right] \leq 1 / 9$ in AM. Notice that the randomness includes the random tape of $\mathcal{M}$ and the randomness $\mathcal{S}$ used to answer each query.

A transcript of an execution of $\mathcal{M}^{\mathcal{S}}(x)$ is an tuple $\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right)$ consists of the random tape of $\mathcal{M}$, all queries to $\mathcal{S}$ and the correlated answers. The transcript fully determined the execution $\mathcal{M}^{\mathcal{S}}(x)$, and

$$
\begin{aligned}
\operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right]= & \sum_{\begin{array}{c}
\text { transcript }\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \\
\text { determines a execution where } \mathcal{M}^{\mathcal{S}}(x) \rightarrow 1
\end{array}} \operatorname{Pr}\left[\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right)\right] \\
= & \sum_{\substack{\text { transcript }\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \\
\text { determines a execution where } \mathcal{M}^{\mathcal{S}}(x) \rightarrow 1}} \operatorname{Pr}[r] \prod_{t=1}^{T} \mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)
\end{aligned}
$$

In the proof, we construct an $\mathbf{A M}$ protocol that estimate this sum.

Proof of Lemma 4.3 It's sufficient to show that $L=\left(L_{Y}, L_{N}\right) \in \mathbf{A M}$. Then the same argument would shows $\overline{\mathrm{L}}=\left(\mathrm{L}_{N}, \mathrm{~L}_{Y}\right) \in \mathbf{A M}$, which implies $\mathrm{L} \in \mathbf{c o A M}$.

L can be efficiently reduced to a probability-verifiable sampling problem. Let $\mathcal{S}$ denote a correlated sampling oracle. The reduction is a probability polynomial-time oracle algorithm $\mathcal{M}$ such that

$$
\begin{align*}
& x \in \mathrm{~L}_{Y} \Longrightarrow \operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right] \geq \frac{8}{9} \\
& x \in \mathrm{~L}_{N} \Longrightarrow \operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right] \leq \frac{1}{9} \tag{4}
\end{align*}
$$

The probability is over the random tape of $\mathcal{M}$ and the randomness used by $\mathcal{S}$. Without loss of generality, assume there exists $T=\operatorname{poly}(n)$ that $\mathcal{M}$ uses $T$ bits of randomness and makes $T$ queries on any input $x \in\{0,1\}^{n}$.

Define a transcript of an execution $\mathcal{M}^{\mathcal{S}}(x)$ as a tuple $\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right)$ where $r \in\{0,1\}^{T}$ is the random tape of $\mathcal{M}, \operatorname{pd}_{t}$ is the $t$-th query to sampling oracle $\mathcal{S}$ and $v_{t}$ is the $t$-th sample returned by $\mathcal{S}$. The length of $v_{t}$ is bounded by some polynomial of $n$, let $\ell(n)$ be a polynomial that upper bound $\left|v_{t}\right|$.

Note that the input, the random tape and oracle's answers fully determine the reduction. Given the input and random tape, the reduction's first query is predictable; given the input, random tape and the oracle's previous answers, the reduction's next query is predictable. Therefore, we define a transcript $\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right)$ to be valid, if it's potentially a transcript of an execution $\mathcal{M}^{\mathcal{S}}(x)$, i.e. if for all $1 \leq t \leq T$, $\operatorname{pd}_{t}$ would the $t$-th query in execution $\mathcal{M}^{\mathcal{S}}(x)$ when $r$ is the random tape and $v_{1}, \ldots, v_{t-1}$ is the oracle's previous answers. By this definition, $\sigma$ is a valid transcript doesn't implies $v_{t}$ has non-zero probability under distribution $\mathrm{pd}_{t}$. Let $C(x)$ denote the set of all valid transcripts of $\mathcal{M}^{\mathcal{S}}(x)$.

The transcript also determines the output of the reduction. Define a transcript $\sigma$ to be accepting, if $\sigma$ is valid and the corresponding execution $\mathcal{M}^{\mathcal{S}}(x)$ output 1 . Let $C_{1}(x)$ denote the set of all accepting transcripts of $\mathcal{M}^{\mathcal{S}}(x)$.

Let $P_{x}(\sigma)$ denotes the probability that $\sigma$ is the transcript of $\mathcal{M}^{\mathcal{S}}(x)$ when the random tape is uniformly chosen and $\mathcal{S}$ is an ideal sampling oracle. Then by chain rule,

$$
P_{x}(\sigma)=\frac{1}{2^{T}} \prod_{t=1}^{T} \mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)
$$

for any valid transcript $\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right)$. For any input $x$, we know $C_{1}(x) \subseteq$ $C(x)$,

$$
\sum_{\sigma \in C(x)} P_{x}(\sigma)=1, \quad \sum_{\sigma \in C_{1}(x)} P_{x}(\sigma)=\operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right]
$$

by the definition of valid/accepting transcripts. Thus, by condition (4), to distinguish between $x \in$ $\mathrm{L}_{Y}$ and $x \in \mathrm{~L}_{N}$, it's sufficient to distinguish between $\sum_{\sigma \in C_{1}(x)} P_{x}(\sigma) \geq 8 / 9$ and $\sum_{\sigma \in C_{1}(x)} P_{x}(\sigma) \leq$ $1 / 9$.

Define $D(x)$ as the set of all tuple $(\sigma, k)$ such that $\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \in C_{1}(x)$, and $k$ is an integer that

$$
1 \leq k \leq K \cdot P_{x}(\sigma)=K \cdot \frac{1}{2^{T}} \prod_{t=1}^{T} \mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)
$$

where $K=10 \cdot 2^{T} \cdot 2^{T(\ell+1)}$. Then the size of $D(x)$ is roughly $K \cdot \operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right]$ if $K$ is sufficiently large.

The sampling problem is probability-verifiable. By definition, there exists a family of error function $\left\{\eta_{\mathrm{pd}, m}\right\}$ such that for any pd, $m$, the error function $\eta_{\mathrm{pd}, m}:\{0,1\}^{*} \rightarrow[0,+\infty)$ satisfies $\sum_{v} \eta_{\mathrm{pd}, m}(v) \leq 1$, and the promise problem

- YES instances: $\left(\mathrm{pd}, v, \hat{p}, 1^{m}\right)$ such that $\hat{p}=\mathcal{P}_{\mathrm{pd}}(v)$
- NO instances: $\left(\mathrm{pd}, v, \hat{p}, 1^{m}\right)$ such that $\hat{p} \geq \mathcal{P}_{\mathrm{pd}}(v)+\frac{1}{m} \eta_{\mathrm{pd}, m}(v)$
is in AM. Let ProbLowerBound be the corresponding AM protocol.
Let set $D^{\prime}(x)$ consist of all tuple $(\sigma, k)$ such that $\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \in C_{1}(x)$, and $k$ is an integer that

$$
1 \leq k \leq K \cdot \frac{1}{2^{T}} \prod_{t=1}^{T}\left(\mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{T} \eta_{\mathrm{pd}_{t}, T}\left(v_{t}\right)\right) .
$$

Here $K=10 \cdot 2^{T} \cdot 2^{T(\ell+1)}$ as in the definition of $D(x)$. By definition, $D(x) \subseteq D^{\prime}(x)$.

Claim. The promise problem

- YES instances: $(x, \sigma, k)$ such that $(\sigma, k) \in D(x)$
- NO instances: $(x, \sigma, k)$ such that $(\sigma, k) \notin D^{\prime}(x)$
is in AM.
Proof. TranscriptChecking is an AM protocol that solves this promise problem.
AM protocol TranscriptChecking on input $\left(x, \sigma=\left(r, \operatorname{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right), k\right)$
V: Check whether $\sigma$ is a valid accepting transcript of $\mathcal{M}^{\mathcal{S}}(x)$; Reject if not
$\mathbf{P}:$ Send $\hat{p}_{1}, \ldots, \hat{p}_{T}$, an honest prover should send $\hat{p}_{t}=\mathcal{P}_{\text {pd }_{t}}\left(v_{t}\right)$
$\mathbf{P}, \mathbf{V}$ : Run protocol ProbLowerBound $\left(\mathrm{pd}_{t}, v_{t}, 1^{10 T}\right)$ for all $1 \leq t \leq T$, repeat polynomial many times in parallel and take majority so that the total error probability is exponentially small; Reject if either of these protocols reject.

V: Check whether $1 \leq k \leq K \cdot \frac{1}{2^{T}} \prod_{i=1}^{q} \hat{p}_{i}$; Reject if not

For $(\sigma, k) \in D(x)$, an honest prover could convince the verifier that to accept $(x, \sigma, k)$.
Any prover, even if it's malicious, should send $\hat{p}_{t}$ such that $\hat{p}_{t} \leq \mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)$. Otherwise the prover will be caught in ProbLowerBound protocol with overwhelming probability. Thus no prover can make the verifier accept $(x, \sigma, k)$ with high probability if $(\sigma, k) \notin D^{\prime}(x)$.

Claim. The size of $D(x)$ is at least $\frac{2}{3} K$ if $x \in \mathrm{~L}_{Y}$.
Proof. $x \in \mathrm{~L}_{Y}$ implies that $\operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right] \geq \frac{8}{9}$. Thus

$$
\begin{aligned}
|D(x)| & =\sum_{\sigma \in C_{1}(x)}\left\lfloor K \cdot P_{x}(\sigma)\right\rfloor \\
& \geq \sum_{\sigma \in C_{1}(x)}\left(K \cdot P_{x}(\sigma)-1\right) \\
& =K \cdot \sum_{\sigma \in C_{1}(x)} P_{x}(\sigma)-\left|C_{1}(x)\right| \\
& \geq K \cdot \operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right]-|C(x)| \\
& \geq \frac{8}{9} K-2^{T} \cdot 2^{T(\ell+1)} \\
& =\frac{8}{9} K-\frac{1}{10} K \\
& \geq \frac{2}{3} K
\end{aligned}
$$

Claim. $D^{\prime}(x)$ has size at most $\frac{1}{3} K$ if $x \in \mathrm{~L}_{N}$.

Proof. $x \in \mathrm{~L}_{N}$ implies that $\operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right] \leq \frac{1}{9}$.

$$
\begin{aligned}
\left|D^{\prime}(x)\right|= & \sum_{\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \in C_{1}(x)}\left\lfloor K \cdot \frac{1}{2^{T}} \prod_{t=1}^{T}\left(\mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)\right)\right\rfloor \\
\leq & K \cdot \sum_{\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \in C_{1}(x)} \frac{1}{2^{T}} \prod_{t=1}^{T}\left(\mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)\right) \\
= & K \cdot \sum_{\sigma=\left(r, \mathrm{pd}_{1}, \ldots, v_{T}\right) \in C_{1}(x)}\left(\frac{1}{2^{T}} \prod_{t=1}^{T}\left(\mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)\right)-\frac{1}{2^{T}} \prod_{t=1}^{T} \mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)\right) \\
& +K \cdot \sum_{\sigma=\left(r,, \mathrm{pd}_{1}, v_{1}, \mathrm{pd}_{2}, v_{2}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \in C_{1}(x)} \frac{1}{2^{T}} \prod_{t=1}^{T} \mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right) \\
\leq & K \cdot \sum_{\sigma=\left(r, \mathrm{pd}_{1}, \ldots, v_{T}\right) \in C(x)}\left(\frac{1}{2^{T}} \prod_{t=1}^{T}\left(\mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)\right)-\frac{1}{2^{T}} \prod_{t=1}^{T} \mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)\right) \\
& +K \cdot \operatorname{Pr}\left[\mathcal{M}^{\mathcal{S}}(x) \rightarrow 1\right] \\
\leq & \left(e^{1 / 10}-1\right) K+\frac{1}{9} K \\
\leq & \frac{1}{3} K .
\end{aligned}
$$

The second to last inequality symbol relies on the following inequality,

$$
\begin{aligned}
& \quad \sum_{\substack{\sigma=\left(r, \mathrm{pd}_{1}, v_{1}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \in C(x)}}\left(\frac{1}{2^{T}} \prod_{t=1}^{T}\left(\mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)\right)\right) \\
& =\sum_{\substack{\left(r, \mathrm{pd}_{1}, v_{1}, \ldots, \mathrm{pd}_{T-1}, v_{T-1}, \mathrm{pd}_{T}\right) \\
\exists v_{T}\left(r, \mathrm{pd}_{1}, v_{1}, \ldots, \mathrm{pd}_{T}, v_{T}\right) \in C(x)}}\left(\frac{1}{2^{T}} \prod_{t=1}^{T-1}\left(\mathcal{P}_{\mathrm{pd}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)\right) .\right. \\
& \left.\leq \sum_{v}\left(\mathcal{P}_{\mathrm{pd}_{T}}(v)+\frac{1}{10 T} \eta_{\mathrm{pd}_{T}, 10 T}(v)\right)\right) \\
& \leq \sum_{\substack{\left(r, \mathrm{pd}_{1}, v_{1}, \ldots, \mathrm{pd}_{T-1}, v_{T-1}\right) \\
\exists \mathrm{pd}_{T}, v_{T}\left(r, \mathrm{pd}_{1}, \ldots, v_{T}\right) \in C(x)}}\left(\frac{1}{2^{T}} \prod_{t=1}^{T-1}\left(\mathcal{P}_{\mathbf{p d}_{t}}\left(v_{t}\right)+\frac{1}{10 T} \eta_{\mathrm{pd}_{t}, 10 T}\left(v_{t}\right)\right)\left(1+\frac{1}{10 T}\right)\right) \\
& \vdots \\
& \leq \sum_{r \in\{0,1\}^{T}} \frac{1}{2^{T}}\left(1+\frac{1}{10 T}\right)^{T} \\
& \leq \\
& \leq\left(1+\frac{1}{10 T}\right)^{T} \\
& \leq e^{1 / 10} .
\end{aligned}
$$

Combining the claims above, $L$ can be reduced to the following promise problem

- YES instances: $x$ such that $\left|D^{\prime}(x)\right| \geq|D(x)| \geq \frac{2}{3} K$;
- NO instances: $x$ such that $|D(x)| \leq\left|D^{\prime}(x)\right| \leq \frac{1}{3} K$.

This promise problem can be solved in AM using the set lower bound protocol of Goldwasser and Sipser [GS86]. Thus L $\in$ AM.

## 4.3 $\mathrm{DGS}_{s}$ is Probability-Verifiable

By Lemma 3.1, for any approximation factor $\gamma$, if gapSIVP ${ }_{\gamma / \mu} \in \mathbf{S Z K}$ for any constant $\mu>1$, there exists a function $g$ maps lattice bases to real numbers such that $g$ is in $\mathbb{R}_{\eta^{-}}$TFAM and $\lambda_{n}(\mathbf{B}) \leq g(\mathbf{B})<\gamma(n) \lambda_{n}(\mathbf{B})$.

For any base $\mathbf{B}$ and lattice point $\mathbf{v} \in \mathcal{L}(\mathbf{B})$. As $g \in \mathbb{R}_{\eta}$-TFAM, the verifier can force the prover to provide a sufficiently accurate estimation of $g(\mathbf{B})$, denoted by $\hat{g}$. As $\hat{g} \approx g(\mathbf{B}) \geq \lambda_{n}(\mathbf{B})$, the verifier can ask the prover to provide a set of linearly independent vectors $\mathbf{W}=\left(\mathbf{w}_{1}, \ldots, \mathbf{w}_{n}\right)$ such that $\|\mathbf{W}\| \leq \hat{g}$. Here the length of a vector set, e.g. $\|\mathbf{W}\|$, is defined as the length of the longest vector in the set.

Given such a short independent vector set $\mathbf{W}$, there exists an efficient algorithm that samples from discrete Gaussian distribution $\mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}$ such that $\hat{s}=\varphi(n) \cdot \hat{g}$ and $\varphi(n)=\Theta(\sqrt{n \log n})\left[\mathrm{BLP}^{+} 13\right.$, GPV08]. Moreover, the verifier can estimate the probability that $\mathbf{v}$ is sampled from $\mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}$ using the set lower bound protocol.

Let $s(\mathbf{B})=\varphi(n) \cdot g(\mathbf{B})$, then $\hat{s}$ is a good estimation of $s(\mathbf{B})$. If the bias between $\hat{s}$ and $s(\mathbf{B})$ is sufficiently small, one could expect $\operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}\right] \approx \operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), s(\mathbf{B})}\right]$.

Proof of Lemma 4.4. By Lemma 3.1, gapSIVP ${ }_{\gamma(n) / \sqrt{\pi \log (2 n+4)}} \in$ SZK implies the existence of a function $g$ maps lattice bases to real numbers such that $g$ is in $\mathbb{R}_{\eta}$-TFAM and $g(\mathbf{B}) \in[3$. $\left.\lambda_{n}(\mathbf{B}), \gamma(n) / \sqrt{\log (2 n+4) / \pi} \cdot \lambda_{n}(\mathbf{B})\right]$. Define $s(\mathbf{B})=\sqrt{\ln (2 n+4) / \pi} \cdot g(\mathbf{B})$, thus

$$
2 \lambda_{n}(\mathbf{B}) \leq 3 \cdot \sqrt{\ln (2 n+4) / \pi} \cdot \lambda_{n}(\mathbf{B}) \leq s(\mathbf{B})<\gamma(n) \lambda_{n}(\mathbf{B})
$$

Given any basis $\mathbf{B}$, vector $\mathbf{v} \in \mathcal{L}(\mathbf{B})$ and precision parameter $m$. The verifier can learn a good estimation on $g(\mathbf{B})$, denoted by $\hat{g}$. As $g(\mathbf{B}) \geq \lambda_{n}(\mathbf{B})$, the verifier could ask the prover to provide a set of linearly independent vectors of $\mathcal{L}(\mathbf{B})$, denoted by $\mathbf{W}$, such that $\|\mathbf{W}\| \leq \hat{g}$.

Given a set of linearly independent vectors $\mathbf{W}$ that $\|\tilde{\mathbf{W}}\| \leq \hat{g}$, there is an efficient algorithm which samples from discrete Gaussian $\mathcal{N}_{\mathcal{L}(\mathbf{B}), \sqrt{\ln (2 n+4) / \pi} \cdot \hat{g}}\left[\mathrm{BLP}^{+} 13\right]$. Let $\mathcal{S}$ denote this sampling algorithm. Let $\hat{s}=\sqrt{\ln (2 n+4) / \pi} \cdot \hat{g}$, then $\hat{s}$ is a good approximation of $s(\mathbf{B})$. Let $r$ be the random tape in the sampling algorithm $\mathcal{S}$, then

$$
\operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}\right]=\frac{\left\{r: \mathcal{S}\left(B^{\prime}, \hat{s}\right) \text { outputs } \mathbf{v} \text { when } r \text { is the random input tape }\right\}}{2^{|r|}}
$$

We could use the set lower bound protocol to lower bound this probability $\operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}\right]$. Thus the promise problem

- YES instances: $\left(\mathbf{W}, \mathbf{v}, \hat{s}, \hat{p}, 1^{m}\right)$ such that $\mathbf{v} \in \mathcal{L},\|\tilde{\mathbf{W}}\| \leq \frac{\hat{s}}{\sqrt{\ln (2 n+4) / \pi}}, \hat{p}=\operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}\right]$
- NO instances: $\left(\mathbf{W}, \mathbf{v}, \hat{s}, \hat{p}, 1^{m}\right)$ such that $\hat{p} \geq\left(1+\frac{1}{m}\right) \operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}\right]$
is in AM, as it can be solved by protocol ProbLowerBound.
To prove $\mathrm{DGS}_{s}$ is probability-verifiable, it is sufficient to show that ProbLowerBound is an AM protocol that estimate the probability $\operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}\right]$ with high accuracy. The estimation error of ProbLowerBound has two sources: (a) the inaccuracy of the set lower bound protocol, which

AM protocol ProbLowerBound on input (B, v, $\hat{p}, 1^{m}$ )
P: Send $\hat{g}$, an honest prover should send $\hat{g}=g(\mathbf{B})$
$\mathbf{P}, \mathbf{V}:$ Convince the verifier that $|\hat{g}-g(\mathbf{B})| \leq c \delta \cdot g(\mathbf{B})$, where $\delta=\frac{1}{n m^{2}}, c$ is a sufficiently small constant
$\mathbf{P}:$ Send $\mathbf{W}=\left(\mathbf{x}_{1}^{\prime}, \ldots, \mathbf{x}_{n}^{\prime}\right)$
$\mathbf{V}$ : Check if $\mathbf{W}$ is a basis of $\mathcal{L}(\mathbf{B})$ and $\|\tilde{\mathbf{W}}\| \leq \hat{g}$
$\mathbf{P}, \mathbf{V}:$ Run the set lower bound protocol to convince the verifier that $\hat{p} \leq\left(1+\frac{1}{2 m}\right) \operatorname{Pr}\left[\mathbf{v} \leftarrow \mathcal{N}_{\mathcal{L}(\mathbf{B}), \hat{s}}\right]$, where $\hat{s}=\sqrt{\ln (2 n+4) / \pi} \cdot \hat{g}$
introduce an $O\left(\frac{1}{m}\right)$ multiplicative error; and (b) the inaccuracy when estimating $s(\mathbf{B})$. Let $\eta_{\mathbf{B}}(\mathbf{v})$ be the estimation error, the error term satisfies

$$
\begin{equation*}
\mathcal{N}_{\mathbf{B}, s(\mathbf{B})}(\mathbf{v})+\eta_{\mathbf{B}}(\mathbf{v}) \leq\left(1+\frac{1}{2 m}\right) \max _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \mathcal{N}_{\mathbf{B}, \hat{s}}(\mathbf{v}) \tag{5}
\end{equation*}
$$

To complete the proof, it is sufficient to show that $\sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \eta_{\mathbf{B}}(\mathbf{v})=O\left(\frac{1}{m}\right)$. By summing (5) over $\mathbf{v} \in \mathcal{L}(\mathbf{B})$,

$$
1+\sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \eta_{\mathbf{B}}(\mathbf{v}) \leq\left(1+\frac{1}{2 m}\right) \sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \max _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \mathcal{N}_{\mathbf{B}, \hat{s}}(\mathbf{v}) .
$$

Thus it is sufficient to show

$$
\begin{align*}
& \sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \max _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \mathcal{N}_{\mathbf{B}, \hat{s}}(\mathbf{v}) \leq 1+O\left(\frac{1}{m}\right) .  \tag{6}\\
& \sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \max _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \mathcal{N}_{\mathbf{B}, \hat{s}}(\mathbf{v})=\sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \max _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \frac{\rho_{\hat{s}}(\mathbf{v})}{\rho_{\hat{s}}(\mathcal{L}(\mathbf{B}))} \\
& \leq \sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \frac{\max _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \rho_{\hat{s}}(\mathbf{v})}{\min _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \rho_{\hat{s}}(\mathcal{L}(\mathbf{B}))}  \tag{7}\\
& \leq \frac{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}))}{\rho_{(1-\delta) s}(\mathcal{L}(\mathbf{B}))}
\end{align*}
$$

For any short vector $\mathbf{v}$ that $\|\mathbf{v}\| \ll s / \sqrt{\delta}$, the relative difference between $\rho_{(1+\delta) s}(\mathbf{v})$ and $\rho_{(1-\delta) s}(\mathbf{v})$ is small.

$$
\frac{\rho_{(1+\delta) s}(\mathbf{v})}{\rho_{(1-\delta) s}(\mathbf{v})}=\exp \left(\frac{\pi\|\mathbf{v}\|_{2}^{2}}{s^{2}}\left((1-\delta)^{-2}-(1+\delta)^{-2}\right)\right)=1+O\left(\delta \cdot \frac{\|\mathbf{v}\|_{2}^{2}}{s^{2}}\right)
$$

Let radius $r=s \cdot \sqrt{n} \cdot \log m$, then

$$
\begin{equation*}
\frac{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}) \cap r \mathcal{B})}{\rho_{(1-\delta) s}(\mathcal{L}(\mathbf{B}) \cap r \mathcal{B})}=1+O\left(\delta \cdot \frac{r^{2}}{s^{2}}\right)=1+O\left((\log m)^{2} / m^{2}\right)=1+o(1 / m) \tag{8}
\end{equation*}
$$

For long vectors in lattice $\mathcal{L}(\mathbf{B})$, the sum of their probability in distribution $\mathcal{N}_{\mathcal{L}(\mathbf{B}),(1+\delta) s}$ is small. In particular, by Lemma 4.5

$$
\begin{align*}
\frac{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}))}{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}) \cap r \mathcal{B})} & =1+\frac{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}) \backslash r \mathcal{B})}{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}) \cap r \mathcal{B})} \\
& \leq 1+O\left(\log m \cdot e^{-\pi(\log m)^{2}}\right)  \tag{9}\\
& =1+o(1 / m)
\end{align*}
$$

Inequality (6) is proved by combining (7)(8)(9),

$$
\begin{aligned}
\sum_{\mathbf{v} \in \mathcal{L}(\mathbf{B})} \max _{|\hat{s}-s(\mathbf{B})| \leq \delta \cdot s(\mathbf{B})} \mathcal{N}_{\mathbf{B}, \hat{s}}(\mathbf{v}) & \leq \frac{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}))}{\rho_{(1-\delta) s}(\mathcal{L}(\mathbf{B}))} \\
& \leq\left(1+o\left(\frac{1}{m}\right)\right) \cdot \frac{\rho_{(1+\delta) s}(\mathcal{L}(\mathbf{B}) \cap r \mathcal{B})}{\rho_{(1-\delta) s}(\mathcal{L}(\mathbf{B}) \cap r \mathcal{B})} \\
& \leq\left(1+o\left(\frac{1}{m}\right)\right)\left(1+o\left(\frac{1}{m}\right)\right)
\end{aligned}
$$

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    ${ }^{1}$ The length of a vector set is defined as the length of the longest vector in the set.

